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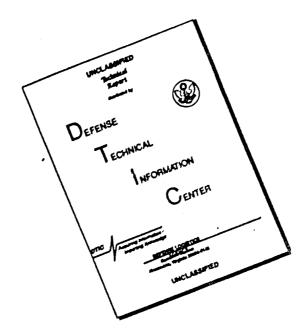
SCIENTIFIC AND TECHNICAL INFORMATION

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U.S. ARMY TANK-AUTOMOTIVE CENTER
WARREN, MICHIGAN

JOHN W. WISS

Lt Colonel, Ordnance Corps Chief, Components R&D Laboratories

Approved:

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A method of estimating statistical characteristics of stable ground roughness for use in off-road land locomotion is developed. Computer programs for obtaining these estimates from survey data have been prepared. Several estimated power spectral densities obtained from survey data taken at military installations are presented.

Pertinent features of these estimated power spectral densities are interpretable in terms of visual features of plotted ground profiles. The estimated power spectral densities may be approximated by similar and simple form.

OBJECT

Develop a statistical method of estimating characteristics of stable ground roughness for use in off-road land locomotion. Apply method to data obtained from surveys.

RESULTS

Computer programs have been constructed to estimate stable ground roughness characteristics. Several estimated power spectral densities are presented.

CONCLUSIONS

Power spectral density methods are useful for characterizing quantitatively stable ground roughness. Features of the estimated p.s.d.'s are interpretable in terms of visual features of plotted ground profiles. Estimated p.s.d.'s may be approximated by similar and simple form in cases considered.

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GLOSSARY

Stable ground roughness is the variation in elevation of hard ground whose influence on vehicle dynamics remains reasonably constant over distance, and is free of obstacles.

<u>Statistical estimation</u> is the technique of estimating average properties from sample data.

<u>Power spectral density</u> (p. s.d.) of a profile measures the amount of variation, by frequency bands, of the profile height.

Ground profile is the plot of survey height vs. distance.

SECTION I

INTRODUCTION

The Midwest Applied Science Corp. was asked by the Land Locomotion Laboratory, Detroit Arsenal, to identify and study the several areas which determine the speed of military vehicles under off-road hard ground conditions. Three areas have been identified as pertinent and are now under study; they are specification or description of stable ground roughness, comfort or tolerance limits of humans to vehicle oscillation, and dynamical characteristics of a military vehicle. Although results are usually presented separately, the separate studies in these areas must be synthesized into vehicle suspension system concepts which will permit increases in vehicle speeds.

This report is concerned only with the description of stable ground roughness.

The problem of describing the roughness of fluctuations in quantities which vary with time or distance is an old one. It has been and still is of interest in diverse areas such as physics, economics, mathematics, engineering, neurology, etc. Questions of the type: What is the composition of white light?, Will the stockmarket fall next week?, Is a function smooth?, Will noise in a telephone mask the message?, How rough is the road a car can travel safely at 60 m.p.h., etc., all relate to the roughness of quantities which change with time or distance.

Early ways of describing roughness are contained in the motion of bounded variation, magnitude of the Fourier coefficients of a periodic motion, power in frequency bands of light as determined by spectrographs, electrical power in frequency bands of fluctuating electrical current, etc. Some of these concepts are familiar to all of us.

That the designer of military vehicle suspension systems is vitally interested in ground roughness, or more particularly, stable ground roughness, follows from the substantial number of test courses at the various military installations around the country which are supposed to simulate "rough ground".

A military vehicle operating off roads encounters two types of hard ground roughness. One type consists of obstacles such as large rocks, tree stumps, stream beds, abrupt hills, etc. The other type consists of variations in elevation which are stable over reasonably large areas and change but gradually with distance. The latter type is what we have called stable ground roughness.

In general, obstacles require special driver attention dictated by the capabilities of the vehicle. What is an obstacle to one vehicle is not an obstacle to another.

Rapid traversibility of the stable roughness depends upon the vehicle, suspension system, and cargo. Again, what is rough to one vehicle may not be rough to another vehicle or to the same vehicle at another speed. Thus, stable ground roughness has meaning in two contexts: in one, only geometry is involved; in the other, geometry eoupled with vehicle-driver-cargo and speed is of importance. Hence, for the vehicle-driver-cargo system, it is impossible to identify the roughest terrain without knowledge of speed and system characteristics.

Environmental roughness exists and is now being studied for other than land vehicles. Descriptions of the roughness of the sea, atmospheric turbulence, and airport runways are needed to predict the vibrational state in ships and airplanes. Starts have been made in obtaining descriptions in terms of power spectral density and some of the preliminary results obtained from such descriptions appear useful to the designer. We note, of course, the lack of obstacles in air and sea travel, and the lack of long unimportant trends such as gradual hills. Here, as with land vehicles, what is rough to an airplane or ship depends upon speed and dynamical characteristics.

It is worth explicitly noting that there are no prior references to be cited in connection with statistical descriptions of stable offroad hard ground roughness.

OBJECT

The object of this report is to present statistical methods of describing or estimating quantitatively the characteristics of stable roughness of hard ground of interest in vehicle dynamics, and to present estimates of roughness obtained from survey data. Smoothing techniques must be developed to eliminate features of roughness unimportant in vehicle dynamics, such as long gradual hills. Further, the assumption that the underlying roughness is reasonably stable over lengths large in comparison with vehicle lengths must be verified.

RESULTS

Figures 1-4 present line profiles obtained from elevation surveys taken at the indicated places. Horizontal distance was measured with a tape; elevation was measured with level and rod.

In each of the first three figures, the complete profile is shown; each data point (represented by a dot) is shown, and the data points are two feet apart. Figure 4 shows every tenth data point and the data points are one foot apart. The vertical scale is substantially different from the horizontal; a twenty foot interval is marked to give an idea of the horizontal scale. In Figure 4, a complete section has been added at each end to indicate details.

The effect of ground roughness upon a vehicle is dependent on the vehicle's size and speed. It is obvious that in the context of today's military vehicles some of the variations in ground elevation in these figures are of such a scale as not to be of interest in vehicle dynamics. For example, in Figure 4, the profile has an upward sloping trends which indicates the ascent of a hill. Superimposed upon this hill are the variations in elevation or roughness which a vehicle will feel. In the other figures, such a striking trend is not present although there are large scale elevation variations which might be termed hills (in these figures) and are not of interest in this study. However, if one mentally eliminates these slowly varying elevation changes, one notes that the remaining variations or roughness patterns appear to be fairly constant. What is needed is a quantitative description of the roughness pattern.

Whatever the description adopted for roughness measurement, it is essential that it be usable by the suspension system designer. Moreover, it must classify as having the same or different roughness (two roads or fields classified by experience as having the same or different roughnesses). We shall, in this report, be concerned with the power spectral density description of ground roughness.

A complete description of power spectral density analysis of roughness on a line is given below for the interested reader. Briefly and loosely speaking, the power spectral density of an oscillating function or ground elevation along a line provides a measure of the power or intensity in frequency (cycles/ft.) bands; all oscillating functions having the same p.s.d. are classified as having the same roughness.

Figures 5 to 8 present the p.s.d.'s (power spectral densities) estimated from the profiles given in Figures 1 to 4, respectively.

The ordinate has units of feet and the abscissa is frequency with unit feet -1.

The general features of these p.s.d.'s are much the same. There is a sharp spike at the origin with a more or less rapid drop as frequency increases, and a long low tail. In two eases, Figure 5 and Figure 6, there is a bump in the rapid drop-off at approximately λ = .025 and λ = .035, respectively. The tails show irregularities at a very low level, in all cases below 0.30. Some of the irregularities in the tails are linked to inconsistencies in the profile data. However, the main features of interest in vehicle dynamics lie in the frequency range where the p.s.d. is descending to the tails.

Let us return to Figure 1. (Aberdeen) This profile shows no prominent hills. All elevations lie between $\stackrel{+}{-}$ 4 feet. This is more or less characteristic of the area within the Perryman Mud Course where the profile was taken. Qualitatively, the roughness seems to eonsist of changes in elevation of approximately 2 feet within distances of 20 to 30 feet. In the first third, there is a fairly periodic oscillation of ground elevation with a period of 30 to 40 feet; in the middle third, they are almost gone. One notices occasionally in the profile sharp irregularities in the elevation between adjacent neighboring points; these are due to ruts, holes, rocks, errors in data recording, etc.

The Aberdeen p.s.d., Figure 5, indicates just a hint of a bump at λ = .025, corresponding to the more or less regular oscillation of approximately 40 ft. wave length in the profile noted earlier. We see below that the p.s.d. of the first half has a much more pronounced bump at λ = 0.025 while the second half does not have this bump.

The profile from Ft. Knox, Figure 2, also does not exhibit any substantial long range variation, except after 1600 feet. As the profile was taken in bottom land bordering the Ohio River, this is reasonable. The roughness has amplitudes which are about half that observed in the Aberdeen profile. There is a pronounced oscillation between 400 ft. and 1600 ft. with a wave length of approximately 27 feet; this explains the presence of the bump in the corresponding p.s.d. estimate, Figure 6, at λ = 0.035. The minor irregularities in the profile are distributed through it as in the Aberdeen profile.

The profile taken at Yuma, Figure 3, exhibits major long range changes in elevation, varying between +12 feet and -10 feet. These are represented by hills of approximately 400 feet length; throughout the profile there are smaller hills of about 50 feet length; the pertinent roughness, from the point of view of vehicle dynamics, exhibits no significiant periodicities. The scale of this latter roughness appears smaller than present in either Aberdeen or Knox. The p.s.d. estimate, in Figure 7, indicates a much higher value near the origin, corresponding to the smaller hills, and a sharper drop due to less roughness. The ordinate of the p.s.d. at λ = 0.03 is 1.0 whereas Aberdeen and Knox are substantially higher, confirming this observation.

All 6201 data points of the Battlefield Day profile (taken at Fort Knox) could not be shown on the main profile, the spacing between points shown is 10 ft.; consequently the relevant roughness is largely masked. Two sections at either end have been plotted completely to

give samples of the type of roughness met. The main profile exhibits one long rising trend with superimposed minor hills. The scale of the roughness of interest, indicated only in part in the detailed portions, appears to be smaller even than in Yuma. Rarely do these roughness variations exceed 1/4 feet. The p.s.d., Figure 8, drops faster than in the Yuma p.s.d., as would be expected. The fluctuation in the tail of this p.s.d., as well as its elevation, will be discussed below; essentially it is traceable to inconsistencies in the data.

CONCLUSIONS

Our main conclusion is that the principal problems of characterizing ground roughness by p.s.d. methods are behind us. We must now turn to the more difficult task of relating the measurements to the suspension system design.

The raw data for the one and two dimensional cases are available in IBM card decks. Anyone interested in these data may obtain them by writing to MASC. These data will also be tabulated and presented in a later report.

The following practical conclusions appear to be consistent with the results obtained to date:

One-dimensional

- a. Visual observed roughness characteristics of the profile are related to characteristics of the estimated p.s.d.
- b. The estimated p. s.d.'s may be approximated with reasonable accuracy by a relatively simple class of functions.
- c. Abnormal height and periodic oscillations in the tails of the p.s.d. estimates are traceable to inconsistencies in the survey data.
- d. Sections of ground which appear visually to have constant roughness characteristics satisfy reasonably well the stationarity assumption on intervals of 1000-2000 feet length.
- e. The normality assumption for the smoothed data is reasonably well satisfied although there is a consistent deviation from it.

- f. There may be considerable difficulty in removing unwanted hills (long range trends) from the data.

 Running average methods have proved satisfactory with present data.
- g. The one-dimensional p.s.d. estimate computer program is almost ready for distribution.

Two-dimensional

h. P.s.d. estimates obtained to date, although not completely satisfactory, display general features observed in one-dimensional results.

In general, successes to date in obtaining p.s.d. estimates are quite gratifying. There are details, however, mostly related to smoothing and interpretation for two-dimensional p.s.d.'s, which must be worked out. We anticipate no substantial difficulties in cleaning up these details and arriving at programs available for distribution.

RECOMMENDATIONS

More surveys are necessary to provide p.s.d. estimates for a wider variety of ground types.

Smoothing procedures must be examined for their effects on characteristics relevant to vehicle dynamics.

Cross-correlation effects from the two-dimensional data must be ascertained.

Section II

One-dimensional Power Spectral Densities

1. Introduction, Interpretation of Spectra

The spectral analysis of random functions of one variable is the subject of a number of books and articles in the engineering literature. We will review in this chapter the principal ideas with emphasis on the descriptive aspects.

It is often the case that a random function is of interest as an input to a linear or almost linear system. Its effect in such a problem is conveniently separated into the sum of effects due to frequency components. The effect of terrain roughness on the vibration of a traveling vehicle is such a problem. As a first approach we will consider a single track vehicle. The profile heights on the path of the vehicle will be a function, h(x), of the distance, x along the track. The roughness of the profile will be a matter of the variation of h(x) about some smooth medial terrain consisting of the large hills etc. The problem of obtaining this smooth terrain is a practical one and is discussed in a later section. Let us denote the deviations of the profile height by d(x). These will average zero and their variation is what we mean by roughness. As a drastic simplification of separation into frequency components, we might have

$$d(x) = \sum_{k=0}^{M} (a_k \sin \omega_k x + b_k \cos \omega_k x)^{*}.$$
(1.1)

It is very unrealistic to consider, as in (1.1), that only a few frequencies are present, we should admit all frequencies each in an infinitesimal amount. These ideas are familiar to most engineers as an intro-

^{*}Generally we will distinguish between frequencies in radians per foot (ω) and cycles per foot (λ) .

duction to the Fourier integral. Here, however, the fact of randomness complicates matters. If d(x) with fixed frequencies is random, then it is the coefficients a_k , b_k that are random. How to deal with infinitesimal random quantities in the case of continuous frequencies, presents a special problem in probability theory which we cannot hope to deal with here. The unrealistic, fixed frequency, case, does "converge to" the realistic case and it is worthwhile to consider it in some detail.

If a function of the form (1.1) is used as an input to a linear system, the output also has this form

$$a(x) = \sum_{K} (a_{k}^{2} \sin \omega_{k} x + b_{k}^{2} \cos \omega_{k} x)$$
 (1.2)

The coefficients a'_k , b'_k of the k'th pair depend only on a_k , b_k of the k'th pair of (1.1). The form of dependence is stated most clearly if we change the form of (1.1), (1.2)

$$a_n \sin \omega_n x + b_n \cos \omega_n x = A_n \cos (\omega_n x + \phi_n)$$

$$A_n = \sqrt{a_n^2 + b_n^2} \qquad \phi_n = \arctan (a_n/b_n)$$
(1.3)

Then with the general term of (1.2) in the same form it may shown that

$$A_{n}^{*} = A_{n} P (\omega_{n})$$

$$\phi_{n}^{*} = \phi_{n} + \Phi (\omega_{n})$$
(1.4)

The functions $P(\omega)$ and $\Phi(\omega)$ combine to form what is called the transfer function of the linear system. The phase of the input does not affect the amplitude of the output, nor does the amplitude of the input affect the phase of the output. In many problems the ultimate

effect does not depend on the phase of the output so that only A_n and hence only A_n is of interest. The various A_n arranged according to the sizes of the ω_n form what is called the <u>frequency spectrum</u> of the input. Their sizes give the amounts of the frequencies present. For a continuum of frequencies this concept becomes that of a spectral density.

This is a quick account of the standard method of analyzing the effect of an input d(x) on a linear system. We have not mentioned that d(x) is random -- by random we mean that d(x) may turn out to be any one of an "ensemble" of functions and we have no hope of determining which one, but we can determine or estimate some ensemble averages. We will assume that only the coefficients a_n , b_n are random in the expression (1.1). This is no restriction, for we can assume that all the possible frequencies are represented among the ω 's and that their absence or presence is a matter of which coefficients are or are not zero.

It will frequently be convenient to use the exponential form of the trigonometric series (1.1). If the ω_k are given for k=0,1,2,...,M and ω_0 =) this is

$$d(x) = \sum_{K=-M}^{M} \alpha_k e^{i\omega_k x}$$

where

$$\alpha_{k} = \frac{1}{2} (b_{k} - i\alpha_{k}) \qquad k > 0$$

$$\alpha_{-k} = \alpha_{k}^{*} = \frac{1}{2} (b_{k} + i\alpha_{k}) \qquad k > 0$$

$$\alpha_{0} = b_{0}$$
(1.5)

Complex quantities appear on the right, but d(x) is real.

We assume two main properties for the randomness of d(x). The first is that the ensemble average at any point x is zero. From the linear property of the expected value (ensemble average) operation,

$$E \left\{ d(x) \right\} = \sum_{k} E \left\{ \alpha_{k} \right\} e^{i\omega_{k}x} = 0 \qquad (1.6)$$

The series above, being identically zero as a function of x, must have zero coefficients, i.e.

$$E \left\{\alpha_{k}\right\} = \frac{1}{2} \left(E\left\{b_{k}\right\} \pm iE\left\{a_{k}\right\} = 0$$

$$E\left\{a_{k}\right\} = E\left\{b_{k}\right\} = 0$$

Our second assumption is that of stationarity. A consequence of the assumption is that any expected value concerning the d's must remain unchanged for different placements of the origin of x. In particular the following does not depend on x:

$$E\left\{d\left(x+s\right)d\left(x\right)\right\} = E\left\{\sum_{k}\sum_{j}\alpha_{k}\alpha_{j}^{*}e^{i\omega_{k}(x+s)-i\omega_{j}x}\right\}$$
$$=\sum_{k}\sum_{j}E\left\{\alpha_{k}\alpha_{j}^{*}\right\}e^{i(\omega_{k}-\omega_{j})x}e^{i\omega_{k}s}$$

In this expression we have made use of the fact that d(x) = d*(x) for convenience. The series on the right, which must be identically constant for all x, falls into two parts,

$$\sum_{k} E \left\{ \alpha_{k} \alpha_{k}^{*} \right\} e^{i\omega_{k}s} + \sum_{k \neq j} E \left\{ \alpha_{k} \alpha_{j}^{*} \right\} e^{i(\omega_{k} - \omega_{j})x} e^{i\omega_{k}s}$$

If the ω_k are all different, which we may assume from the start, then the second series must be identically constant as a function of x hence its coefficients must be zero;

$$E\left\{\alpha_{k} \alpha_{j}^{*}\right\} = 0 \qquad k \neq j$$

Thus if d(x) is to be stationary, it is necessary that the complex coefficients be mutually uncorrelated. In this case

$$E\left\{d(x+s)\ d(x)\right\} = R(s) = \sum_{k} E\left\{\alpha_{k} \alpha_{k}^{*}\right\} e^{i\omega_{k}s} \qquad (1.7)$$

R(s) is called the covariance function of d(x). From (1.3)

$$|\alpha_k|^2 = \alpha_k \alpha_k^* = \frac{1}{4} (\alpha_k^2 + b_k^2) = \frac{1}{4} A_k^2$$
 (taking $A_k = A_k$)

$$|\alpha_0|^2 = \alpha_0^2 = A_0^2$$

This gives the important relation M

$$R(s) = E\left\{A_0^2\right\} + \sum_{K=-M}^{\infty} E\left\{\frac{1}{4}A_k^2\right\} e^{i\omega_k s}$$

$$\neq 0$$
(1.8)

which links the covariance fuction with the spectrum. Note that the components are all in phase, in fact we can write

$$R(s) = E\left\{A_0^2\right\} + \frac{1}{2}\sum_{k=1}^{M} E\left\{A_k^2\right\} \cos \omega_k s \qquad (1.8^{\dagger})$$

Therefore R(O) is a maximum (positive) value of R(s). The coefficients in the trigonometric series (1.8) are ensemble averages and are possible to know or to estimate. When they are arranged in order of the sizes of the ω_k , they form what is called the power spectrum of the random function d(x). From (1.8) The power spectrum is the frequency spectrum of the covariance function.

To summarize what we have discussed to now, we will deal with random inputs to linear systems. The spectrum will characterize the effect. However, the spectrum is random and cannot be known except for its ensemble averages. The power spectrum is a collection of such averages. The different spectral values are uncorrelated. The power spectrum is the spectrum of the covariance function.

To interpret the power spectrum, let us imagine that the frequency values $\boldsymbol{\omega}_{k}$ are very close together. We have from (1.8) that

R(0) = E
$$\{A_0^2\} + \sum_{k=1}^{M} E \{\frac{1}{2}A_k^2\}$$
 (1.9)

$$R(0) = E \left\{ d^{2}(x) \right\}$$
 (1.10)

Assuming that d(x) is approximately normal and has a mean of zero. R(O) is its variance and the probability is about .003 that d(x) will exceed in absolute value $3\sqrt{R(O)}$. Thus R(O) is a measure of the approximate extremes of sizes in profile height. The usual values of d(x) lie between $\pm 2\sqrt{R(O)}$. From (1.9) we see that the left side is limited by the range of oscillation of d(x), so that the terms of the sum of the right side are also limited in size. If there are many frequencies, then most of them have small values for $E\{A_{k}^{2}\}$.

Equation (1.10) has another interpretation through the ergodic hypothesis which says that ensemble averages may be found as limits of time averages. In particular

$$\lim_{T} \frac{1}{T} \int_{0}^{T} (d(x))^{2} dx = E \{(d(x))^{2}\} = R(0)$$

The left side is a measure of the average variability of d(x) over a long stretch. By equation (1.9) this is divided into averages of uncorrelated components each of which has a separate effect on the linear system. In our work we will usually be interested more in frequency bands, and it is natural to consider that the total variance can be partitioned into partial variances due to frequency bands. In fact, suppose we have several random functions $d_1(x)$, $d_2(x)$, etc. each of which may be represented as a trigonometric series of form (1.5) with random coefficients. Since the frequencies are different we will assume the coefficients to be uncorrelated. Thus $d_1(x)$, $d_2(x)$, etc. are also uncorrelated so that the variance of the sum is the sum of the variances

Var
$$\{d_1(x) + d_2(x) + ...\} = R_1(0) + R_2(0) + ...$$

Calling the sum d(x), we have

Var
$$\{d(x)\} = \sum_{i} E\{\frac{1}{2}A_{k}^{2}\} + \sum_{i} E\{\frac{1}{2}A_{k}^{2}\} + \dots$$
 (1.12)

which shows that the variance of d(x) can be partitioned in to the sums of variances of various frequency components each of which may be thought of as a separate random function the sum of which is d(x) itself.

Sometimes a power spectrum has a bump which we can "isolate" by sketching a "normal" spectrum on which it is superimposed. The power in the bump is the added area over the normal

^{*} Σ_1 , for example, means sum over those ω_k 's which are frequency components of $d_1(x)$, etc..

spectrum. We may visualize this power as being manifested approximately by a sine wave of the frequency at the center of the bump and amplitudes which may change but which average about the square root of twice the area. With pronounced bumps such oscillations are often visible in the profile data.

2. Relation of the theory to the estimation problem.

In estimating spectral values from profile data we will be motivated principally by two relations. The first is (1.8) which links the power spectrum to the covariance function, and the second is another aspect of the ergodic hypothesis (1.11),

$$R(s) = L_{\overline{I}}^{im} \frac{1}{T} \int_{0}^{T} \left[d(x+s) d(x) \right] dx \qquad (2.1)$$

These equations suggest that the values of R(s) might be estimated by taking mean lagged products of the data, and then the values of the power spectrum might be gotten by determining the Fourier coefficients in the expansion of R(s). Before going into details we must make some things definite that have, to now, been left vague.

We will assume a continuous power spectrum given by a spectral density function so that (1.8) becomes

$$R(s) = \int_{-\infty}^{\infty} f(\lambda) e^{2\pi i \lambda s} d\lambda = \int_{-\infty}^{\infty} f(\lambda) \cos 2\pi \lambda s d\lambda$$

$$(f(-\lambda) = f(\lambda)).$$
(2.2)

The quantity $f(\omega_k)d\lambda$ may be associated with the value of the sum of $E\{\frac{1}{4}|A_n^2\}$ for those ω_n within a band of width $d\lambda$ about ω_k . The dimension of the values of $f(\lambda)$ is $(distance)^3$ since $d\lambda$ is in eyeles per foot or $(dist.)^{-1}$ and A^2 is $(dist.)^2$.

We must also take into account the form of the data that we will have to work with. This will be n values of an outcome of d(x) taken at evenly spaced intervals:

$$d(x_0 + k\Delta)$$
, $k = 0, 1, 2, ..., n-1$.

We shall define a mean lagged product as

hall define a mean lagged product as
$$\begin{cases} \frac{1}{n-j} & \sum_{k=0}^{n-j-1} d(x_0+k\Delta+j\Delta)d(x_0+k\Delta) & \text{if } j \geq 0 \\ \frac{1}{n-|j|} & \sum_{k=j}^{n-1} d(x_0+k\Delta+j\Delta)d(x_0+k\Delta) & \text{if } j < 0 \end{cases}$$
 then

$$r_j = r_{-j}$$

It is apparent from the definition of R(s), (1.7) that

$$E\left\{r_{j}\right\} = R\left(j\Delta\right)$$

By the ergodic hypothesis, the probability limit, as n^{∞} of r_i is $R(j\Delta)$. It is also evident that we cannot hope to estimate any values of the covariance function other than $R(j\Delta)$ and that this limitation is imposed by the fact that we have only observed evenly spaced data. Our data is also limited to a finite number of values so that the number of covariance values is also limited. We should limit the number of lags so that the number in terms (n-|i|) in the mean (2.3) is sufficient to make it a good estimate. Let us denote the maximum lag by $m:r_i$ is calculated for $j=0,\pm 1,\pm 2...\pm m$.

The above indicates how we shall estimate the covariance function. The determination of its Fourier coefficients is accomplished by taking a linear combination of its values with cosine coefficients. We will investigate the properties of linear combinations of the r_i 's with any sort of coefficients w_i

$$\widetilde{W} = \sum_{j=-m}^{m} r_j w_j \qquad (2.6)$$

The expected value of this estimate is

$$E\left\{\widetilde{W}\right\} = \sum_{-m}^{m} R(j\Delta) w_{j}$$

$$= \sum_{-m}^{m} \int_{-\infty}^{\infty} f(\lambda) w_{j} \cos 2\pi \lambda j \Delta d\lambda$$

$$= \int_{-\infty}^{\infty} f(\lambda) \left[\sum w_{j} \cos 2\pi \lambda j \Delta\right] d\lambda \qquad (2.7)$$

$$= \int f(\lambda) W(\lambda) d\lambda$$

where

$$W(\lambda) = \sum_{-m}^{m} w_j \cos 2\pi \lambda_j \Delta \qquad (2.8)$$

Ideally if in (2.7) E $\{\widehat{W}\}$ were to be the value $f(\lambda_0)$ then $W(\lambda)$ should be a Dirac impulse or "delta" function centered at λ_0 . However $W(\lambda)$ is a finite Fourier series and cannot be a Dirac Function --in fact it is periodic with a period of $1/\Delta$ and cannot be zero except at isolated points.

The function $W(\lambda)$ of (2.8) is called the "spectral window". The estimates \widetilde{W} obtained by linear combinations of the covariance estimates are not estimates of the values of the power spectrum, but of averages of those values of $f(\lambda)$ admitted by the window as in (2.7). By making $W(\lambda)$ have a narrow peak of unit area near λ_0 and very low values elsewhere we will estimate very nearly the power over an interval of frequencies near λ_0 . This corresponds to one of the sums in the partition (1.12) of the total power.

The properties of $W(\lambda)$ arising from its form as a finite Fourier series (2.8) give several limitations of the linear estimate.

From the fact that $W(\lambda)$ is even and periodic, $W(\lambda_0) = W(1/\Delta - \lambda_0)$. This means that while trying to estimate $f(\lambda_0)$ we will unavoidably pick up what power there is in $f(\lambda)$ around $1/\Delta - \lambda_0$ since there is also a peak in the window around this point. All the power beyond the frequency $1/2\Delta$ must unavoidably be confused with power in the range 0, $1/2\Delta$. There is also another peak at $-\lambda_0$ which picks up the power around $f(-\lambda_0)$, but since the spectral density function is even, this is not a bad feature; it means, however, the $E\{\widehat{W}\}$ approximates $f(\lambda_0) + f(-\lambda_0)$ unless the total area of the two peaks is unity. The window has an infinity of other peaks, all of them contributing spurious power to the estimate, in the same manner as the peak at $1/\Delta - \lambda_0$. These spurious contributions are called "aliasing", and their exact nature depends on the sampling interval Δ . In designing the sampling method, Δ must be chosen so that the total power in the frequencies beyond $1/2\Delta$ is not appreciable.

From the fact that $W(\lambda)$ can be zero only at isolated points (actually $z^{m}W(\frac{1}{2\pi i\Delta}\log z)$ is a polynomial in z of degree 2m + 1 and cannot have more than 2m + 1 zeros) it follows that in trying to estimate $f(\lambda_0)$ we will unavoidably pick up some power in $f(\lambda)$ near the non-zero parts of $W(\lambda)$. If the spectrum has a high spot at one of the off center non-zero places of $W(\lambda)$ the power picked up may be considerable. If $W(\lambda)$ is negative, this spurious power is subtracted from the total and may make the expected spectral estimate go negative. These non-zero off center parts of the window are called "side lobes". To make them small one must at the same time make the peak broader so that $E\left\{ \widehat{W}\right\}$ becomes an average of a wider band of frequencies. The band width is inversely proportional to m and, though the exact definition of band width is not fixed (the window cannot be a rectangle) it is usually taken to be about $1/2m\Delta$. In this way the size of the sample is related, through the maximum lag number, m, to the resolution or band width of the estimate.

These limitations on the exactness of the estimate are not of a statistical nature. The statistical errors are random deviations of the estimate \widetilde{W} from its expected value $E\{\widetilde{W}\}$ (2.6), (2.7). What we have considered so far are properties of $E\{\widetilde{W}\}$; it is an average of spectral power over a band of frequencies with unavoidable contamination from aliasing and from the side lobes of the window. Considerable attention has been paid to the shape of the window (see, for example, Blackman and Tukey, "The Measurement of the Power Spectra", Dover 1958, Parzen, "Mathematical Considerations in the Estimation of Spectra", Technometrics Vol. 3, No. 2, May 1961, 167-190). Most of the considerations in the selection of a window are more a matter of art than of hard-eyed analysis, but some of the aspects of calculation are based on simple mathematical identities which we will now present.

The first window that comes to mind is the finite Fourier series whose coefficients are the first $\, m \,$ Fourier coefficients of the Dirac function centered at $\, \lambda_{\, 0} \,$. In terms of (2.8)

$$\overline{W}_{j} = \Delta \int_{-\frac{1}{2\Delta}}^{\frac{1}{2\Delta}} \delta(\lambda - \lambda_{0}) \cos 2\pi\lambda_{j} \Delta d\lambda = \Delta \cos 2\pi\lambda_{0} j \Delta (2.9)$$

which gives the window function

$$\overline{W}(\lambda) = \Delta \sum_{-m}^{m} \cos 2\pi \lambda_{0} j \Delta \cos 2\pi \lambda j \Delta \qquad (2.10)$$

This sum may be expressed in closed form by writing the cosines in exponential form and summing the resulting geometric series, or by using the formula $\sum \cos j\theta = \sin (m+1/2) \theta / \sin \theta / 2$:

$$\overline{W}(\lambda) = \frac{\Delta}{2} \left(\frac{\sin \pi (2m+1)(\lambda + \lambda_0) \Delta}{\sin \pi (\lambda + \lambda_0) \Delta} + \frac{\sin \pi (2m+1)(\lambda - \lambda_0) \Delta}{\sin \pi (\lambda - \lambda_0) \Delta} \right)$$
(2.11)

This function has high, sharp peaks at λ_0 and $-\lambda_0$ as well as the aliasing peaks at $-\lambda_0^+$ + k/ Δ . It also has large side lobes including negative ones (Graphs may be seen in Blackman and Tukey, referred to above, or in Jenkins, "General Considerations in the Analysis of Spectra, <u>Technometrics</u> Vol. 3, No. 2, May 1961, 133-166). Both of these are bad features of this window.--We have already discussed the side lobes, the narrowness has the effect of increasing the statistical errors as we shall see later. Integrating (210) term by term over $(-1/2\Delta, 1/2\Delta)$, all but the 0 term vanishes and the result is unity. This means that the two peaks of \overline{w} in this interval have a total area of unity and for $\lambda = 0$ the single peak has unit area. Thus \overline{w} always estimates an approximation to $f(\lambda_0)$.

All of the spectral windows presented in the literature are modifications of this simple window--actually of these simple windows since there are different ones for each λ_o . The modification does not depend on λ_o and consists in multiplying the coefficients \overline{w}_j of the simple window (2.9) by constants c_j , $(c_{-j}=c_j)$. These may be applied to the covariances r_j before the calculations (2.6) are made with the simple window coefficients so that the simple windows for each λ_o are all applied to the same modified covariances. The modified window function may be calculated in terms of the simple window. We will use the notation \overline{w} and \overline{w} to refer only to the simple window defined by (2.9) and (2.10). We will express (2.10) in exponential form:

Let us denote the cosine transform of the constants c_k by

$$C(\lambda) = \sum_{m=0}^{m} c_k \cos 2\pi \lambda k \Delta = \sum_{m=0}^{m} c_k e^{2\pi i \lambda k}$$

Then the convolution of C and \overline{W} ,

$$W(\lambda) = \Delta \int_{-\frac{1}{2\Delta}}^{\frac{1}{2\Delta}} \overline{W}(\omega) C(\lambda - \omega) d\omega$$

$$= \Delta \sum_{j} \sum_{k} \overline{w}_{j} c_{k} e^{2\pi i \lambda k \Delta} \int_{-\frac{1}{2\Delta}}^{\frac{1}{2\Delta}} e^{2\pi i (j - k) \omega \Delta} d\omega$$

$$= \sum_{k} c_{k} \overline{w}_{k} e^{2\pi i \lambda k \Delta}$$

is the modified window. This equation allows $W(\lambda)$ to be calculated directly in terms of the simple window. It also furnishes a basis for selection of $C(\lambda)$ or c_k to attain a desired $W(\lambda)$ We have shown, however, that we do not have complete freedom in selecting a window, and (2.12) can only be a guide in these matters.

The modified windows in usual use are wider and have smaller side lobes than does the simple window. Graphs of some of them may be seen in Blackmand and Tukey, or in Jenkins (Dover, and Technometrics, v. 3, 2 referred to above). In general they compare to an ideal rectange centered at λ_0 and of width $1/2m\Delta$. If this ideal were the case and we strung out such rectangles without overlapping we could put about 2m+1 in the interval $(-1/2\Delta, 1/2\Delta)$. Each rectangular window would "view" a separate part of the power spectrum and give estimates of $f(\lambda)$ at 2m+1 evenly spaced points in the interval where it is to be estimated. If we take

the spacing of the points and the width of the rectangles to be 1/(2m+1) then the rectangles are somewhat narrower but they do not go outside the interval where $f(\lambda)$ is to be estimated. Let us use the notation (see (2.6) and (2.9))

$$f_{k} = \Delta \sum_{-m}^{m} r_{j} \cos 2\pi (k\delta) j\Delta = \Delta \sum_{-m}^{m} r_{j} \cos 2\pi \frac{kj}{2m+1}$$
 (2.13)

for the spectral estimates at evenly spaced points using the simple window. Then by using either the exponential form of this series or the formula cited in connection with (2.10), we have the inverse relation

$$r_j = \sum_{k=0}^{m} f_k \cos 2\pi \frac{kj}{2m+1}$$
 (2.14)

The values $\mathbf{f}_{\mathbf{k}}$ are sometimes called the "raw spectrum". The modified window gives the spectral estimates

$$\oint_{\mathbf{k}} = \sum_{m=0}^{m} \mathbf{r}_{j} c_{j} \cos 2\pi \frac{j k}{2m+1}$$
(2.15)

The convolution theorem gives (taking f_k as periodic and even)

$$\oint_{\mathbf{k}} = \sum_{-m}^{m} f_{\mathbf{k}-\mathbf{j}} C (\mathbf{j} \delta)$$
(2.16)

where $C(j \delta)$ is given by (2.11):

$$C(j\delta) = \sum_{m=0}^{m} c_k \cos 2\pi \frac{jk}{2m+1}$$
 (2.17)

With evenly spaced spectral estimates, therefore, the modification of the window may be done by a running average smoothing of the raw spectrum, using as coefficients the finite Fourier transform of the c.'s.

The inverse relation (2.13), (2.14) and the convolution relation (2.15), (2.16), (2.17) are to be used later in a discussion of filtering.

In this section we have discussed the limitations imposed by the finiteness and discreteness of the data on what we must try to estimate. We shall go on now to discuss the random deviations from this.

3. Statistical problem

We have been considering the character of the expected value of linear spectral estimates. The actual outcome of the estimates is random, the ultimate randomness being in the smoothed profile data itself. We will assume that this is approximately Gaussian. Evidence in support of this assumption will be found in the following section. We discuss here the main ideas of the statistical problem without giving detailed proofs. These may be found in Blackman and Tukey (Dover, cited above) or in Grenander and Rosenblatt (Statistical Analysis of Stationary Time Series, Wiley 1957). In Chapter 3 of this report, the section corresponding to this one has similar proofs worked out for the two dimensional case.

The estimate \widetilde{W} (2.6) considered as a function of the n data points, is a homogeneous quadratic polynomial (a quadratic form). In practice such polynomials in Gaussian random variables are found to have distributions that are well approximated by a gamma distribution, i.e.

$$P\left\{\widetilde{W} < z\right\} = \int_{0}^{z} \frac{A^{k}}{\Gamma(k)} t^{k-1} e^{-At} dt$$
 (3.1)

For purposes of using tabulated values it is convenient to recognize this as a generalization of a chi-squared of χ^2 -distribution. In fact (3.1) would be stated equivalently as, $\widetilde{W}/2A$ has a χ^2 distribution with 2k degrees of freedom, except that 2k need not be an integer.

The gamma distribution is fitted by the method of moments, that is the values of A and k are chosen so that the first two moments of the gamma distribution match those of W.

$$4 A k = E \left\{\widetilde{W}\right\}$$

$$16 A^2 k = Var \left\{\widetilde{W}\right\}$$

from which
$$k \frac{\left[E\left\{\widetilde{W}\right\}\right]^{2}}{\operatorname{Var}\left\{\widetilde{W}\right\}} \qquad A = \frac{E\left\{\widetilde{W}\right\}}{4k} \qquad (3.2)$$

Using a χ^2 table we may find a 95% confidence interval for E $\{\widetilde{W}\}$ by getting numbers L and R such that

$$P \left\{ L < X^2 < R \right\} = .95$$

Then since $\widetilde{W}/2A = 2kW/E\{\widetilde{W}\}$ has a χ^2 distribution

$$P\left\{\frac{2k}{R}\widetilde{W} < E\left\{\widetilde{W}\right\} < \frac{2k}{L}\widetilde{W}\right\} = .95$$

The factors 2k/R and 2k/L give the ends of a confidence interval for the random deviations of the observed W from the expected W. A short table will illustrate how the width of the interval depends on the degrees of freedom, 2k.

2k = deg.fd	$\frac{2k}{R}$	$\frac{2k}{L}$
5	.39	6.02
10	. 495	3.08
15	. 546	2.40
20	.622	2.09
25	.615	1.91
30	. 640	1.79
40	. 675	1.64
50	.700	1.54
60	.720	1.48

The degrees of freedom, 2k, measure the amount of statistical error. Various equivalent measures of statistical error are listed in the paper by Jenkins cited above (<u>Technometrics v. 3</u>, No. 2). They all depend ultimately on the first two moments of \widetilde{W} . The first moment of \widetilde{W} is given in equation (2.7) which we repeat now

$$E\left\{\widetilde{W}\right\} = \int_{-\infty}^{\infty} f(\lambda) W(\lambda) d\lambda \qquad (2.7 rep)$$

The variance of W may be approximated in a similar form by making use of the assumption of approximate normality.

$$Var \left\{ \widetilde{W} \right\} = \frac{1}{\Delta(n-m)} \int_{-\infty}^{\infty} f^{2}(\lambda) W^{2}(\lambda) d\lambda \qquad (3.3)$$

The proof of this formula, which is long and tedious, may be found in Blackman and Tukey (Dover 1958, cited above) and will be given in the next chapter for the two dimension case. With these two formulas, the degrees of freedom formula (3.2) becomes

$$2k = 2\Delta (n-m) \frac{\left[\int f(\lambda) W(\lambda) d\lambda\right]^2}{\int f^2(\lambda) W^2(\lambda) d\lambda}$$

If the power spectral density function, $f(\lambda)$, does not have a steep slope in the region where the window is large, and if it vanishes substantially outside the interval $(-1/2\Delta, 1/2\Delta)$ then the integrals may be approximated by taking an average value of $f(\lambda)$ out and these average values will approximately eancel giving

$$2k = (n - m) \frac{\left[\int_{-\frac{1}{2\Delta}}^{\frac{1}{2\Delta}} W(\lambda) d\lambda \right]^{2}}{\int_{-\frac{1}{2\Delta}}^{\frac{1}{2\Delta}} W^{2}(\lambda) d\lambda}$$

The numerator of this should be unity for a properly normalized window function, so the size of the denominator gives the dependence of the statistical error on the window. Generally it is larger for the sharper windows. For an ideal window, a rectangle of width $1/2m\,\Delta$ and height $m\,\Delta$ (to make a unit area for the two peaks; for the window at λ =0 where the two peaks coincide, the height should be $2m\,\Delta$.) we have a final approximation

$$2k = 2\Delta(n-m) - \frac{1}{2m^2\Delta^2} = 2\frac{n-m}{m}$$

This approximation is on the optimistic or large side but it has the advantage of simplicity.

The magnitude of the statistical error may be indicated by confidence intervals. These may be found from the degrees of freedom as in the table given in this section. From the table one can see that precision increases with an increase in degrees of freedom. Formula (3.6) suggests that degrees of freedom increases as the ratio g = m/n decreases. To decrease this we must either raise the number, n, of observations or lower the number, m, of lags. The former almost always increases the expense of the investigation, and the latter decreases the resolution of the spectral estimates.

4. Smoothing the Profile

In the first section we immediately passed from actual heights h(x) to differences d(x) of the actual heights from smoothed heights, eqn. (1.1). We postponed discussion of this until the present section so that we could relate this discussion to the ideas of the intervening sections. Our mathematical assumption concerning the profile of ground heights, is that it is stationary—of the same character throughout the region of measurement. In most areas this property is clearly untrue for any possible region, though it might be "made true" for suitable regions if the hills could be flattened. In their turn the hills might appear as roughness in very large regions in which general statewide level changes were removed and so on. Our problem is to carry out the investigation on a scale relevant to our purpose.

The first part of limiting the scale of the investigation is to limit the extent of the region to be measured so that the roughness-other than "hills" is of a uniform character. The second part is to

remove the hills in some way. In the limited region they are not usually stationary. (In the Battlefield Day data the ground rises steadily so that the heights at the end are not nearly the same as those at the beginning.) They are also not relevant as vibration producing roughness. Nevertheless the variations in height due to these "hills" are ordinarily much larger than those variations we want to study. If we were to leave them in the data the spectrum would consist almost exclusively of a central value at zero frequency. The other values would be lost in the round-off of the computation.

We have used two methods of smoothing the profile. One method consists in constructing a smooth profile by fitting paraboloids to segments of the data. The second--so far in use only with the linear data--is a running average scheme.

The first method consists in building up a smoothed profile by fitting second degree polynomials H(x) to the data points $h'(x_i)$. Actually, several polynomials may be fitted each to a separate segment. We may imagine these to be graphed along with the data so that the data or actual ground profile cuts back and forth across the smooth parabolical curve. It is clear that a parabola can remove only one hill from each segment so that in a profile with several hills there should be several segments fitted by parabolas. This was done, but it was found that the smoothed profiles obtained in this way had steps where the segments came together. In some of the data the steps, though few in number, were about ten times the values of the d(x) values and an approximate analysis indicated that they could affect the resulting spectra. (See Appendix Y of this chapter.

The theoretical effect of parabolic smoothing is difficult to work out. Computations were made which indicated that this type of smoothing has several advantages over the running average type of smoothing. We have experienced some difficulties in the analysis of our data which suggest that parabolic smoothing works well when the parabolas fit closely but stronger smoothing is needed on hilly ground.

We have, therefore, done some calculations using another spectral program which has been equipped with a smoothing subroutine of the running average type. A theoretical discussion of the effects of this type of smoothing is, fortunately, very easy and we will present it for the linear case.

Let us denote the heights of the ground above points evenly spaced on a straight line of the datum plane by

$$h_{a} h_{a+1}, \dots, h_{0}, h_{1}, \dots, h_{n} h_{n+1}, \dots h_{n+n}$$
 (4.1)

We have added points on each end for a purpose which will be clear later. Running average smoothing requires that we select some coefficients

$$b_{\alpha}, b_{\alpha+1}, \dots, b_{\alpha}, \dots, b_{\alpha} \tag{4.2}$$

The smoothed values will be calculated by the formula

$$y_k = \sum_{\alpha = -\alpha}^{\alpha} b_{\alpha} h_{k+\alpha}$$
 $k = 1, 2, ..., n$ (4.3)

(Because of our foresight the numbers here run from 1 to n. Similar foresight in the computations, accomplished by faking in a few values at the beginning and end of the data accomplishes the same end.) The deviations used for spectral calculation and called "the smoothed profile" are:

$$d_k = h_k - y_k = h_k - \sum_{\alpha = -\alpha}^{\alpha} b_{\alpha} h_{k+\alpha} = \sum_{\alpha = -\alpha}^{\alpha} b'_{\alpha} h_{k+\alpha}$$
 (4.4)

where

$$b'_{\alpha} = \frac{-b_{\alpha}}{1-b_{0}}$$
 if $\alpha = 0$

To show what this type of smoothing can accomplish, let us suppose that the ground heights (5.1) are composed of a smooth "hilly" part P_k plus a rought part δ_k

$$h_k = P_k + \delta_k$$

Then

$$d_{k} = \sum_{\alpha} b'_{\alpha} P_{k+\alpha} + \sum_{\alpha} b'_{\alpha} \delta_{k+\alpha}$$
 (4.5)

Now if the Pk are values of a polynomial on integer points

$$P_k = C_0 + C_1 k + C_2 k^2 + \cdots + C_m k^m$$

Then the coefficients b (4.2) may be selected so that the first term of (4.5) is zero. Suppose, for example, the P_k is of 2nd degree. Then

$$\frac{\sum_{\alpha=-\alpha}^{\alpha} b_{\alpha}' P_{k+\alpha} = \sum_{\alpha=-\alpha}^{\alpha} b_{\alpha}' \left[C_0 + C_1 (k+\alpha) + C_2 (k+\alpha)^2 \right]}{\sum_{\alpha=-\alpha}^{\alpha} b_{\alpha}' \left[(C_0 + C_1 k + C_2 k^2) + \alpha (C_1 + 2C_2 k) + \alpha^2 C_2 \right] (4.6)}$$

$$= (C_0 + C_1 k + C_2 k^2) \sum_{\alpha=-\alpha}^{\alpha} b_{\alpha}' + (C_1 + 2C_2 k) \sum_{\alpha=-\alpha}^{\alpha} a_{\alpha} b_{\alpha}' + C_2 \sum_{\alpha=-\alpha}^{\alpha} a_{\alpha}' b_{\alpha}'$$

If we select the coefficients so that

$$\sum_{\alpha} b_{\alpha}^{\prime} = 0$$

$$\sum_{\alpha} \alpha b_{\alpha}^{\prime} = 0$$

$$\sum_{\alpha} \alpha^{2} b_{\alpha}^{\prime} = 0$$
(4.7)

Then the right side of (4.6) will be zero. In this sense the running average smoothing can remove a parabolic trend. Note that the smoothing is very strong in that the polynomial P may be different at each point. Since (4.7) is a set of three homogenous linear equations with coefficients

1, 1, ... 1, 1, 1, ... 1, 1
-a, -a+1, ., -1, 0, 1, ...
$$a-1$$
, a
 a^2 , $(a-1)^2$, ..., 1, 0, 1, ... $(a-1)^2$, a^2 .

For a = 0 or 1 the matrix is non singular so there is only the trivial--all zero--solution. Thus a must be at least 2. The solutions are not unique and for larger a's there is a large variety of them.

The removal of trend is not the only effect of this smoothing. In (4.5) the second term on the right is the smoothed time series whose spectrum we want. When we estimate the spectrum of the smoothed series we will not be estimating what we want. Suppose we denote

$$\gamma_{k} = \sum_{-\alpha}^{\alpha} b_{\alpha}^{\prime} \delta_{k+\alpha}$$
 $k = 1, 2, ..., n$

The covariance function for γ_k is

$$\tilde{r}_{d} = E \left\{ \gamma_{k} \gamma_{k+d} \right\} = E \left\{ \sum_{\alpha=-\alpha}^{\alpha} b_{\alpha} \delta_{k+\alpha} \sum_{\beta=-\alpha}^{\alpha} b_{\beta} \delta_{k+d+\beta} \right\}$$

$$= \sum_{\alpha} \sum_{\beta} b'_{\alpha} b'_{\beta} E (\delta_{k+\alpha} \delta_{k+d+\beta})$$

$$= \sum_{\alpha} \sum_{\beta} b'_{\alpha} b'_{\beta} r_{d+\beta+\alpha}$$

where $r_a = E\{\delta_k \delta_k + a\}$ is the covariance function of the δ 's. If we change variables in the summation so that $\beta - \alpha = \mu$ and consider the b's to be defined for all α , taking the value zero for $\alpha < -a$, $\alpha > a$ then the covariance function of the smoothed series may be written

$$\tilde{r}_{d} = \sum_{\mu} \left(\sum_{\alpha} b_{\alpha}' b_{\mu+\alpha}' \right) r_{d+\mu}$$
 (4.8)

which is a running average smoothing of the γ 's using the coefficients

$$B_{\mu} = \sum_{\alpha} b_{\alpha}' b_{\alpha+\mu}' \qquad (4.9)$$

A running average smoothing of the stationary part of a profile results in a profile whose covariances are a running average smoothing of the original covariances. The coefficients are the "covariances" of the coefficients used to smooth the profile.

The running average (4.8) is a convolution. If we rewrite the convolution relations (2.15), (2.16), (2.17) using the same conventions as to periodicity, we have

$$\widetilde{r}_{d} = \sum_{j=-m}^{m} r_{d+j} B_{j} \qquad (2.16 \text{ rep})$$

$$\tilde{r}_{d} = \sum_{m=1}^{m} f_{j} \beta_{j} \cos 2\pi \frac{jk}{2m+1} \qquad (2.15 \text{ rep})$$

where

$$\beta_{j} = \sum_{m}^{m} B_{k} \cos 2\pi \frac{jk}{2m+1}$$

and f_{j} is the r's given by (2.14).

A running average smoothing of the stationary part of a profile results in a profile whose spectral estimates are multiples of those of the original profile. The factors are the "power spectral estimates" of the smoothing coefficients. This is so since (2.15 rep) shows the spectrum of \tilde{r}_d to be $f_j\beta_j$. If we calculate \tilde{f}_j , the spectral estimates of the smoothed data, then we may correct it using as factors $1/\beta_j$, given by (2.17 rep).

Some work remains to be done on this type of smoothing that needs experience more than mathematical analysis. The smoothing coefficients, b', which we have tried, have a spectrum which is zero at $\lambda = 0$, and very close to zero for low frequencies. The extent of the frequency range for which these spectral values are zero indicates the severity of smoothing by indicating how much of the stationary part of the profile is removed. Correcting may not give it back since it also exaggerates round-off error as well as the low spectral value. Therefore, some care must be exercised in choosing the b' so as to leave in the important part of the desired spectrum. We may choose from the variety of smoothing coefficients, available to us as solutions of equations such as (5.7), those which give a desirable spectrum--equivalently a desirable covariance function (5.8). In general there may be many solutions of (5.7) giving exactly the same covariances. Each of these will act the same according to our mathematical model; they will remove the appropriate polynomial trend and affect the spectrum of the stationary part in the prescribed way. What is the actual effect in practice depends on how our actual problem differs from this mathematical problem. This will require an investigation by experience rather than by logic.

The problem of finding coefficients that have a prescribed spectrum i.e., prescribed correction factors seems to be easier to solve by computation than by analysis. This is another area for experience. As yet, we have used only one set of coefficients so these matters are for future investigation.

Inasmuch as the coefficients b'_j may be chosen so as to make the spectrum f_k have smoother slopes, it will help in the estimation problem. There will be no sharp bumps that leak in

through the side bands and so on. Smoothing for this purpose has been called "prewhitening". Our first computer program has an elaborate prewhitening section in which the spectrum was first estimated and then coefficients were calculated to correct it. Our experience indicates that the main bad part of the spectrum occurs at zero frequency so that simpler methods of prewhitening are more effective.

5. The Computing Programs

We have used two computer programs to estimate line spectra. A complete description of the first program is no longer available. The general operation of that program is as follows:

- 1) The data is "smoothed" by taking deviations of the profile heights from heights of quadratic polynomials fitted to segments by least squares.
 - 2) Covariances (mean lagged products) are computed.
- 3) A preliminary spectrum is calculated by making a cosine transform of the covariances and smoothing it by the Hamming coefficients.
- 4) A trigonometric polynomial of a given number of terms is fitted to the spectrum by least squares. The coefficients of this are used in a series of equations whose solutions are running average coefficients to be used on the profile data. This running average affects the spectrum as outlined in Section 4 and the computations are made so that it tends to flatten out steep places of the spectrum. This is the "Prewhitening" referred to in Section 4.
 - 5) The profile is smoothed by running average.
- 6) New covariances are computed and also their spectrum (cosine transform).

- 7) Correction coefficients for the prewhitening running average are computed.
 - 8) A corrected spectrum is computed.
- 9) The corrected spectrum is smoothed by the Hamming coefficients.

This program recently became inoperable and we have replaced it. We have obtained a spectral program made for the IBM 7090. The complete details of the program are available. Generally it performs these operations:

- 1) Subtracts the average of the entire profile from each profile value.
- 2) Computes covariances of any desired segment of the profile data.
- 3) Computes the spectrum of the covariances using the Hamming W function.

The first step is a sort of smoothing but it is not strong enough for our purposes. We have modified the program by inserting a subroutine which, previous to the above steps

(0) smooths the profile by running average using specified coefficients.

This modification is only temporary and it is lacking in several respects. The most serious respect is that the correction factors must be calculated and applied by hand. Since this type of smoothing seems to be very satisfactory, we are including a new subroutine which will

- (0.0) smooth the profile by running average using specified coefficients.
- (0.1) indicate any profile values that are inconsistent with their neighbors.
 - (0.2) compute and store the correction factors
 - (0.3) correct the spectral estimates.

6. Discussion of One-dimensional Spectral Results

In this section, we shall comment in some detail upon the one-dimensional p.s.d.'s, how they were obtained, and their corresponding profiles.

One of the significant problems from the point of view of subsequent use is the taking out or elimination of trends, hills, etc. which are not considered relevant. When we began the survey data analysis, we estimated trends by means of quadratic polynomials fit to segments of the data by least squares. These polynomials were subtracted from the profile heights to produce smoothed data (deviations from fitted polynomials). The roughness of these smoothed data were then analyzed in the spectral analysis program. The spectral estimated for the Yuma data contained anomalies such as high tail values and occasional negative values. In the case of Aberdeen and Knox, these anomalies were not observed. Moreover, in the Yuma data, spectral estimates of the halves and whole were different in character. It was not possible to interpret these anomalies in terms of the data. This was somewhat surprising in view of the consistancy of the results with Aberdeen and Knox and from highway survey data.

The origin of these anomalies remained a mystery for some time, but attention was gradually directed to the smoothing proceedure. We next obtained another p.s.d. estimate program which only removed the general average; it produced with the smoothed Yuma data precisely the same p.s.d. estimates which established beyond doubt the troubles in the smoothing proceedure. We then made a preliminary routine for running average smoothing which added to the program produced p.s.d. estimates free of most of the anomalies noted above. Moreover, the remaining anomalies were immediately traceable to inconsistant raw data points not previously detected. The results presented here are those obtained from this revised program.

The general features were already discussed in the Introduction and that discussion will not be repeated. We must describe however the significance of data processing and analysis.

As explained in Section II, Part 4, smoothing throws away part of the data and partially restores it by numerical correction. The nature and significance in vehicle dynamics of the part thrown away is not completely known to us and the control of this part by choice of smoothing coefficients is not yet entirely complete. It is important to understand that smoothing must bare a direct relationship to p.s.d. estimate use. Whether hills are removed or not depends upon vehicle size, speed and the presence and nature of the hills. Thus, use of the p.s.d. estimate procedures and the estimates without regard to the method of estimation employed may in unusual conditions lead to error.

Figure 5 shows the p.s.d. estimates of the full 750 data points, the first 350 data points, and the second 350 data points from Aberdeen. The raw data, the smoothed data, the covariance of the smoothed data, the p.s.d. estimates of the smoothed data, and the corrected p.s.d. estimates are given in Tables 1 to 1D. The plots are of the corrected p.s.d. estimates. The uncorrected p.s.d. estimates, calculated from the smoothed data, are distorted by the smoothing and are corrected by multiplication as explained in Section II, Part 4.

In general, the p.s.d. of the full set of data should be approximately the average of the p.s.d.'s of the two halves. This feature is noticeable upon Figure 5. The matching in the significant range λ =.005, to λ =0.075 between the halves is substantial except in the interval containing the bump. The bump is due to a periodic component in the profile, (Figure 1) which was most visible in the first half of the profile. This is confirmed by the presence of a strong bump in the first half p.s.d. and the absence of any bump in the second half. The generally high level in the low tails is due to inconsistencies in the data as evidenced by the readings at

112 feet and 1,254 feet in the smoothed data (Table 1B). As these tails are not of interest, these inconsistencies are not important to us at this time. The other significant difference between halves is in the λ = .09 to λ = .10, where the second half seems to have about twice as much power as the first. Examination of the profile to confirm this by inspection indicates that visual methods are not capable of explaining this difference, but since the power involved is low this is not surprising.

Figure 6 gives the p.s.d. estimates of the full and two half sets of data points from Ft. Knox. The tabulated results are given in Tables 2 to 2D.

The three spectra match better, in general, than those of Aberdeen. In particular, the tails fall off far more rapidly and to a lower value than in Aberdeen. The broad bump in the first half p.s.d. is spread over the wave lengths from 25 feet to 50 feet whereas in the second half this power is concentrated more in the wave length around 30 feet. Again these differences are not noticeable in scanning the profile (Figure 2) by eye.

Figure 7 gives the 3 p.s.d. estimates from the Yuma data. The tabulated results are given in Tables 3 to 3D.

The matching in Yuma is quite remarkable. If anything, the first half p.s.d. estimate indicates less roughness in the higher frequencies than the second half; this is evident in the profile (Figure 3). The drop off to a very low level of power in the tails is also very satisfactory.

Figure 8 gives the 5 p.s.d. estimates from the Battlefield Day data. Numerical data and computed results are tabulated in Tables 4 to 4D. In this figure, we give the p.s.d. estimates of the four quarters as well as of the whole. The estimates match very well up to $\lambda = .08$. There is no noticeable bump or any other feature

of interest in the range from λ = 0.00 to λ = .08. Two of the quarters have comparatively high tails and one exhibits a distinct periodicity. These features are transmitted to the spectra estimate of the full set of data. They are adequately explained by the presence of inconsistent data points, as mentioned in Appendix Y, Section 2.

The differences between p.s.d. estimates of the different tracts of ground are much greater than the differences encountered with estimates from sections of the same tract. This indicates that the assumption of stationarity over reasonable lengths is a good one and justifies the use of p.s.d. techniques in the measurement of ground roughness. It must be expected that the assumption of stationarity is not ideally fulfilled in any tract of ground. However, in ground selected with reasonable care, our data reveals that the assumption is reasonably well fulfilled over lengths of 1,000 feet or more. This indicates the possibility of adjusting suspension for slowly varying roughness.

These p.s.d. estimates have similar shapes though they are from widely different areas of the country. The fact that the spectral shapes are similar and simple in form is a matter of some importance. First, it might be possible to character ground roughness (spectral shape) in terms of three or four parameters. This may simplify classification of ground roughness types; it would permit easy simulation of roughness types; and it would define ranges of suspension system parameters for adjustment to changing ground roughness.

Figures 9, 10, and 11 show cumulative diagrams of the smoothed data from Aberdeen, Knox, and Yuma plotted on probability paper. The scale is the same in each; the horizontal scale is in percentage of the data, the vertical in feet of deviation. The points are plotted to indicate the percentages of data smaller than

various values. The arrangement of values on the horizontal scale is such that a theoretical Gaussian distribution plots as a straight line passing through the points

$$(15.87, \mu - \sigma)$$
, $(0, \mu)$, $(84.13, \mu + \sigma)$

where μ is its mean and σ^2 its variance. Gaussian data will also be straight except for random deviations. The values of μ and σ may be estimated by fitting a straight line and reading the ordinates of the points 15.87, 0, 84.13. Interpretation of deviations from straightness requires some experience but is not difficult.

For a random function satisfying the ergodic hypothesis, the time average of an indicator function, I (A), which is unity if the event A occurs and zero if not, equals the ensemble average or probability of A. This means for profile heights that the fraction of heights smaller than a number, z, should approximate the probability that a height is smaller than z. If the profile is a Gaussian random function, the data should therefore be Gaussian data. The curves for our data indicate systematic departures from straightness. In our opinion the amount of departure is not enough to invalidate the Gaussian approximations we must use.

The curves suggest that the distribution of the data is symmetric, as is the Gaussian distribution, but that it has a higher percentage of its probability in the region of small deviations than does the Gaussian. This might possibly be a result of the effects of erosion to round off peaks and fill in holes. The amount of this departure from the Gaussian is more for Yuma than for Aberdeen or Knox. It is possible that we may find in the future that the deviation from Gaussian for some locations is large enough that we should inquire into its effect on the Gaussian approximations. This effect will be confined to the formulas concerning the statistical errors.

Appendix Y

Effect of Certain Errors on Spectral Calculations.

In this appendix we will present the results of two investigations we carried out in an attempt to explain some anomalous features of the spectral calculations for Yuma. The first of these was done after one of the profile heights was found to be some thirty feet lower than its neighbors. It was extended somewhat in the light of similar errors found in other profiles. The second was concerned with the effect of a bad joint between parts of a profile that has been smoothed by fitting parabolas to two segments. The result of the latter indicated that the unruly behavior of the Yuma spectrum could not be due principally to this type of error so the investigation was not developed further.

For the first investigation we will suppose that the proper ground profile data $d_1d_2...d_n$ is actually presented to the computer as $d'_1d'_2...d'_n$ where

$$d'_{k} = d_{a} + \Delta \qquad k \neq a$$

$$k \neq a \qquad (Y1)$$

The covariances actually computed are

$$r'_{1} = \frac{1}{n-1} \sum_{k} d'_{k} d'_{k+1}$$

$$= \frac{1}{n-1} \sum_{k} d_{k} d_{k+1} + \frac{\Delta}{n-1} \left(d_{a+1} + d_{a-1} \right)$$
(Y2)

for $t \neq 0$, and

$$r_{o} = \frac{1}{n} \left[\sum_{k} d_{k}^{2} + 2 \Delta d_{o} + \Delta^{2} \right]$$

$$= r_{o} + \frac{2 \Delta d_{o}}{n} + \frac{\Delta^{2}}{n}$$
(Y3)

For large Δ the difference between γ_0 and the neighboring γ_1' , γ_2' etc. (on the order of Δ^2/n) is noticeable. If a is near one end of the data, the point where $d_a +_t$ cuts out of Y 2 may be noticeable (to avoid complicated limits of summation we take $d_k = 0$ for k < 0 or k > n.)

The computed spectral values are

$$f'_{u} = f_{u} + \frac{2 \Delta d_{a}}{n} + \frac{\Delta^{2}}{n} + 2 \Delta \sum_{T=1}^{m} \frac{d_{a+1} + d_{a-1}}{n-1} \cos 2\pi \frac{ut}{2m+1}$$
 (Y4)

In this formula the terms

$$\frac{2d_{a}}{n} + \sum_{i=1}^{m} \frac{d_{a+1} + d_{a+1}}{n-1} \cos 2\pi \frac{ut}{2m+1}$$

are a Fourier series of a small segment of the data and may be quite erratic. The dominant term in (Y.4) is Δ^2/n which is a constant or "white noise" term. This was a very prominent feature of the spectrum obtained from the Yuma data when it contained a single large error in the profile.

A single small error will have its size divided by n, the length of the profile record, in the Fourier series and its square divided by n in the white noise term. Since n is usually large, a small error will not have much effect. The effect of several errors may, however, accumulate to add noticeable white noise to the calculated spectrum.

The third quarter of the Battlefield Day profile data presented a spectrum that was affected, seemingly, by both a white noise and a periodic component. The latter had a period of almost exactly six stations, or about .06 cycles per foot. Since 1/.06 = 16.6 feet one can expect a pip on the covariances near station 17. Such a pip is indeed to be found there as well as a pip at zero to account for the white noise. The full run spectrum also had a similar shape and has similar pips so these are not machine errors.

The data was searched for bad points to explainthe pip at the zero covariances and the white noise part of the spectrum. A list of those found is

No. 3175	2.0 ft above adjacent points
4474	5.0 ft above adjacent points
4540	3.0 ft above adjacent points
4557	10.0 ft above adjacent points

The last two are exactly 17 stations apart and are without doubt the cause of the pip at the 17th lag. The white noise parts of the errors add together, as we shall show, and adding up the values Δ^2/n for the above we have

$$(4 + 25 + 9 + 100) / 1550 = .09$$

The height of the zero'th lag for the third quarter data is .0916, white the first lag is -.0112 making a pip of .0928 which is quite close to the above estimate.

If two errors are farther apart than the maximum lag then their effects in (Y2) and (Y3) are additive. However, if two errors, say of Δ_1 and Δ_2 are separated by a distance of to m stations in the data, say

$$y_{\alpha}' = y_{\alpha}' + \Delta_1$$
, $y_{\alpha+\uparrow_{\alpha}}' = y_{\alpha+\uparrow_{\alpha}} + \Delta_2$

Then the effects of these errors on the covariances are

$$\begin{aligned} &r_{t}' = r_{t} + \frac{1}{n - | t |} \left\{ \Delta_{i} (y_{\alpha + t} + y_{\alpha - t}) + \Delta_{2} (y_{\alpha + t_{0} + t} + y_{\alpha + t_{0} - t}) \right\} \\ &r_{t_{0}}' = r_{t_{0}} + \frac{1}{n - | t_{0}|} \left\{ \Delta_{i} (y_{\alpha + t_{0}} + y_{\alpha - t_{0}}) + \Delta_{2} (y_{\alpha + t_{0} + t} + y_{\alpha + t_{0} - t}) \right\} + \frac{\Delta_{i} \Delta_{2}}{n - | t_{0}|} \\ &r_{0}' = r_{0} + \frac{1}{n} \left(2 \Delta_{i} y_{\alpha} + 2 \Delta_{2} y_{\alpha + t_{0}} \right) + \frac{\Delta_{i}^{2} + \Delta_{2}^{2}}{n} \end{aligned}$$

The effects are additive as for widely separated errors except at a lag of to where there is a pip of

$$\frac{\Delta_1}{n-|t_0|}$$

For the Battlefield Day data third quarter errors, this is

$$3x10/(1550 - 17) = .0195$$

In the covariances the 17th is .0178 while the neighbors are -.0020 making .0198. These anomalies in the spectrum are therefore explained adequately by the errors or inconsistencies in the data.

The spectra for Ft. Knox and Yuma seem to be affected somewhat by periodic components and the spectra for Aberdeen seems to be rather high on the high frequency end. The compilation of this data for the purpose of plotting on probability paper shows several outlying in each profile, but none as spectacular as those found in the Battlefield Day data. The cumulative effect of these

could explain the behavior of the spectra. In these cases the actual amount of the bad behavior is small and seems to affect mostly those frequencies that are unimportant in our work.

The second investigation is concerned with a smoothed profile $d_1d_2...d_n \ d_{n+1}...d_{2n}$ which are deviations of the actual profile heights from parabolas fitted to each half. We will suppose that the fitted parabolas do not meet in the middle--thus the d's also do not meet--and we will try to estimate the effect of this error. Our attention will be on the relation between the spectrum f'(u) of the first half, f''(u) of the second, and f'(u) of the entire profile. In theory the average of the two parts should give the whole. To show this, let us recall

$$r_1 = \frac{1}{2n-1} \sum_{k=1}^{2n-\tau} d_k d_{k+\tau}$$
 $t \ge 0$

We will take t≥0 for the sake of having definite limits on the sum. This sum may be partitioned

$$\begin{split} r_{t} &= \frac{1}{2 \, n - 1} - \left[\begin{array}{c} \sum\limits_{k=1}^{n-1} d_{k} \, d_{k+1} + \sum\limits_{k=n-1+1}^{n} d_{k} \, d_{k+1} + \sum\limits_{k=n+1}^{2n-1} d_{k} \, d_{k+1} \right] \\ &= \frac{1}{2 \, n - 1} \, \left[\left(\, n - t \, \right) \, r_{t}^{\, \prime} + \sum\limits_{k=n-1+1}^{n} d_{k} \, d_{k+1} + \left(\, n - t \, \right) \, r_{t}^{\, \prime\prime} \, \right] \\ &= \frac{1}{2} \, \left[\begin{array}{c} \frac{n - \frac{1}{2} \, \tau}{n - \frac{1}{2} \, \tau} \, \right] \, \frac{r_{t}^{\, \prime} + r_{t}^{\, \prime\prime}}{2} \, + \, \frac{1}{2 \, n - t} \, \sum\limits_{n=1+1}^{n} d_{k} \, d_{k+1} \end{split} \end{split}$$

If t is a small fraction of n-t and 2n-t is large with respect to the sizes of the data near the joint, then

$$r_{\dagger} = \frac{r_{\dagger}^2 + r_{\dagger}^{\prime\prime}}{2} \tag{Y6}$$

Since the cosine transform is a linear operation this equation leads to

$$f_{u} \doteq \frac{1}{2} (f'_{u} + f''_{u})$$
 (Y7)

The approximation to the average in the first term on the right of (Y.5) may be made somewhat better by choosing a value of t near, say, 1/3m since the r's are largest near zero. With this (Y.7) may be modified a little

$$f_{u} = \left(1 - \frac{1}{6n}\right) \frac{1}{2} \left(f'_{u} + f''_{u}\right)$$

The central values, however, are not as predictable. The sum is the same kind as used to estimate $\,p_t^{}$ and if it weren't for the error we are discussing we might expect the second term to have the approximate size

$$\frac{1}{2 n-1}$$
 $r_{\uparrow} \approx \frac{1}{6 u} r_{\uparrow}$

which would cancel out the 1/6n part of (Y.8) on transformation. (Of course this is approximate and the amounts 1/6n are very small --no doubt the other approximations make larger effects than this.) The error in matching the two halves of the profile affects this central part only.

We will assume that the values d_k are roughly constant on each side of the half mark and are separated by a distance $\,D\,$

$$d_{k} = \begin{cases} d & n-m \le k \le n \\ d+D & n+1 \le k \le n+m \end{cases}$$

where m is the maximum lag. Then for any value of t $0 \le t \le m$ the central term is

$$\frac{1}{2n-t} \sum_{k=n-t+1}^{n} d_k d_{k+t} = \frac{t}{2n+t} d(d+D) = \frac{d(d+D)}{2n} t$$

Since r_+ is an even function we have for $-m \le t \le m$

$$r_t = \frac{r_1^2 + r_1^{22}}{2} + \frac{d(d+D)}{2n}$$
 | 1 |

The cosine transform of |t| maybe gotten in closed form in formula (428) of Jolley's "Summation of Series" (Dover, 1961)

$$\sum_{t=-m}^{m} |t| \cos 2\pi \frac{ut}{2m+1} = \frac{(m+1)\sin \pi u}{\sin \frac{\pi u}{2m+1}} = \frac{1-\cos \frac{m+1}{2m+1} 2\pi u}{2\sin^2 \frac{\pi u}{2m+1}}$$

$$= (m+1) \frac{\sin \pi u}{\sin \frac{\pi u}{2m+1}} - \left(\frac{\sin \pi \frac{m+1}{2m+1} u}{\sin \frac{\pi u}{2m+1}}\right)^2$$
(Y8)

The denominators both terms rise to near unity as u goes from 0 to m for most values of u they may be approximated by the argument. In fact for

$$0 < \frac{\pi u}{2m+1} < \frac{1}{2}$$

$$\frac{2 m+1}{\pi u} < \frac{1}{\sin \frac{\pi u}{2m+1}} < 1.05 \frac{2 m+1}{\pi u}$$

Thus we may approximate the right side of (Y.8) for 0<u<1/3 m by

$$\frac{(m+1)(2m+1)}{\pi u} \sin \pi u - \frac{(2m+1)^2}{\pi^2 u^2} \sin^2 \left(\frac{1}{2} + \frac{1}{4m+2}\right) \pi u \text{ (Y9)}$$

For u = 0 the left side of (Y.8) is at its maximum

$$m(m+1) \approx (m+\frac{1}{2})^2 = \frac{(2m+1)^2}{2}$$

which is the limit as u>0 of Y.9. Using (Y.8), (Y.9) to approximate the cosine transform of (Y.7) we obtain

$$f_{u} = \frac{1}{2} (f'_{u} + f''_{u}) + \frac{d(d+D)}{n} \frac{(2m+1)^{2}}{4} \left[2 \frac{\sin \pi u}{\pi u} \left(\frac{\sin \frac{\pi u}{2}}{\frac{\pi u}{2}} \right)^{2} \right]$$

Thus the total spectrum will deviate from the average of its parts by a term whose maximum size is approximately

d (d + D)
$$\frac{m^2}{n}$$

which oscillates with a period of about 2 and which damps out as u increases.

The approximations in this second investigation are so rough as to render the result of aid only in determining the order of magnitude of the effect of this type of error on the behavior of the spectrum. Since we will not smooth by parabolic fit to segments in the future, more exact work does not seem justified.

Section III

Two-dimensional Power Spectral Density

I. Introduction

Spectral analysis for random functions of one variable is common in engineering work. Problems involving random functions of two variables have appeared seldom in the literature. Indeed we are aware of but one article in which results of a two-dimensional spectrum were presented, "The Directional Spectrum of a Wind Generated Sea as Determined from Data Obtained by the Stereo Wave Observation Project", Willard J. Pierson ed. (Mcteorological Papers, N.Y.U. College of Engineering, Vol. 2, No. 6, June 1960). The material in the present section is a further development of work done by one of the authors in connection with the above article and appears in full detail for the first time here. The motivations and interpretations connected with this work will be given even though they are in many cases repetitions of similar ideas appearing in the preceeding sections. Thus the section will be self-contained.

To be able to speak mathematically about the height irregularities of a piece of ground, let us suppose that we have established a horizontal datum and a coordinate system on it. The coordinates (x,y) (Measurements in feet) locate a point on the datum and h(x,y) will denote the height in feet of the ground surface above the datum. The pattern of irregularity is given completely by h(x,y), but, in a sense, it is given in too much detail. For another piece of ground having the same kind of irregularity the function h(x,y) will in general be different. In order to extract the common character of the irregularity we will expand on the common mathematical model of a stationary random function.

The actual heights above datum h(x, y) include variations that are not of interest--gentle slopes, large hills, etc. We will invent a variable datum called the "smoothed heights" which consists of the slopes and hills together with the general level of h(x, y) above the actual datum. The differences

$$d(x,y) = h(x,y)$$
 [smoothed height at (x,y)] (1.1)

are the variations of ground height that affect the vehicle. We will conceive of d(x,y) as being random--that is, this plot of ground gives us one of an "ensemble" of possible roughnesses selected at random according to probabilities that characterize the roughness type. When we say that the general level has been taken out we mean that the expected value or ensemble average, denoted $E\{d(x,y)\}$, of d(x,y) at any point is zero, thus, unchanging or stationary from point to point. More than this we assume that all the characteristics of the probabilities are unchanging from point to point. This is the meaning of "stationary random functions".

Stationary random functions of one variable have been used in engineering for more than 20 years. Problems concerning the output of a linear system with a stationary random input are nowadays routine. A brief summary is given in Section II, Part 1 of this report. In our work, such a problem would be to determine the variable force on the driver of a motorcycle which is proceeding at a uniform rate across our ground on a straight line through (x_0, y_0) at an angle to the x-axis. The input is the random function of distance, s.

$$g(s) = d(x_0 + s \cos \alpha, y_0 \alpha s \sin \alpha)$$
 (1.2)

For a two track vehicle our input must be regarded as two correlated random functions (one for each track); the formulas are readily worked out as above. In any case, the output--variable force acting on the driver (or the variable displacement of the driver)--is also a random function. In engineering practice, the effects of such random functions have been found to be related to the covariance function or its spectrum. This is true in electrical engineering problems and in those problems of mechanical engineering which deal with vibration. In these problems frequency of oscillation is of obvious importance. However, the spectrum of frequencies has also been useful in the study of such "non-oscillatory" phenomena as atmospheric turbulence.

In this discussion of random functions of two variables, we will attempt to use standard notation and wording as far as possible and to make a development which is parallel to the usual one for functions of one variable. This may be found in Grenander & Rosenblatt, "Statistical Analysis of Stationary Time Series" (Wiley, 1957) and Blackman & Tukey, "The Measurement of Power Spectra", (Dover, 1958).

We will begin with the definitions of covariance and the spectrum.

R
$$(a,b) = D \{ d(x,y) d(x+a,y+b) \}$$
 (1.3)

In this the symbol E denotes the expected value operator or the ensemble average. Because of the stationarity or probabilistic sameness of one location to another, the average does not depend on the location (x,y) but only on the "lag" (a,b). The covariance function is not random--it is a fixed characteristic of the roughness.

a)
$$R(0,0) = E\{(d(x,y))^2\} \ge 0$$

b) $|R(a,b)| \le R(0,0)$ (1.4)

c)
$$R(-a,-b) = R(a,b)$$

These properties are related to the connection between covariance and correlation

corr.
$$(x,y) = \frac{\text{cov.}(x,y)}{\sqrt{\text{var}(x), \text{var}(y)}}$$

The first property, in other terms, states that the variance is the same for any point x, y. The second property is that of the correlation coefficient $-1 \le corr(x,y) \le 1$. And the third is corr(x,y) = corr(y,x). Since the correlation is a measure of the relationship between the outcomes of two random variables, we may expect R(a,b) to be large and positive if the height at a displacement (a,b) from a point tends to be about the same as the heights at the point itself R will be negative when large height at (x,y) is associated generally with small heights at x + a, y + b. The absolute size of R(a,b) will indicate the strength of the relationship. From continuity of the ground height h(x,y), R(a,b) should be near R(0,0) for small (a,b)--the correlation between d(x,y) and immediately adjacent points should be close to unity.

In many applications the covariance function has been useful mostly in the form of its Fourier transform

$$F(u,v) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i(xu+yv)} R(x,y) dx dy \qquad (1.5)$$

Since $e^{i\theta} = \cos\theta + i\sin\theta$, the quantity $2\pi (xu + yv)$ must be in radians. Since x and y are in feet, $2\pi u$ and $2\pi v$ are in radians/ft., or u, v in cycles/ft.

As with one variable transform this can be inverted

$$R(x,y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i (xu+yv)} F(u,v) du dv$$
 (1.6)

The fact that R(x,y) is an even function, (1.4c) implies that F(u,v) is real valued and in fact

$$F(u,v) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i(xu+yv)} 1/2 \left[R(x,y) + R(-x,-y) \right] dx dy$$

=
$$1/2$$
 $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i(xu+yv)} R(x,y) dx dy$

+
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i (-xu-yv)} R(x,y) dx dy$$

The second integral is obtained by transforming x' = -x, y' = -y then omitting the primes. The two may then be averaged to give

$$F(u,v) \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i(xu+yv)} \frac{1}{2} R(x,y) + R(-x-y) dx dy$$

$$= \frac{1}{2} \left[\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i(xu+yv)} R(x,y) dx dy \right]$$

$$+ \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-2\pi i(-xu-yv)} R(x,y) dx dy$$

$$(1.7)$$

which corresponds to the usual relation between even functions and cosine transforms in one variable Fourier analysis. From (1.7) it easily follows that F(u,v) is an even function so that (1.6) may be written

$$R(x,y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \cos 2\pi (xu + yv) F(u, v) du dv$$
 (1.8)

Some properties of S(uv) are

a)
$$F(-u,-v) = F(u,v)$$

b)
$$F(u,v) \ge 0$$
 (1.9)

c)
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} F(u,v) du dv = R(0,0)$$

(Actually (1.9a) is proved from the fact that R(a,b) is real valued while (1.9b) stems from a generalization of (1.4a,b) called the "positive semi-definite" property of R to wit, for any $C_1C_2...C_n$ and points $(x_1,y_1)(x_2,y_2)...(x_n,y_n)\sum_i\sum_j C_iC_jR(x_i-x_j,y_i-y_j)\geqslant 0$. The derivation of (1.9b) from this is from a simple extension of Bochner's theorem.)

Properties b) and c) suggest that the total variance, R(0,0) may be divided up into components F(u, v) at the frequency values (u, v). This is indeed the case and one can derive generalizations of the various representations of a stationary stochastic process of one variable. We will interpret the spectral density function F(u, v) as giving the "power" of the roughness in the frequency region u cycles/ft. paralle! to the x-axis, and v cycles/ft. parallel to the y-axis. An elementary wave at this frequency will be a cosine wave $\cos 2\pi$ (ux +vy) which has the appearance of corregated iron on the xy plane with the crests parallel to the line ux + vy = 0 and having crests evenly spaced every $\frac{1}{\sqrt{u^2 + v^2}}$ feet. A representation theorem makes us conceive of d(x, y) as a sum of such elementary waves at different frequencies, amplitudes and phases. For purposes of power those waves of the same frequencies are classified together. The power is analogous to the sum of the squared amplitudes. For a more detailed explanation, see Section II, Part 1 of this report.

A vehicle which runs over the rough ground is affected only by the roughness under its wheels. If its track is straight across a stationary field, the roughness is stationary. The simplest case deals only with one track but more realistically we will have two parallel tracks. Let us derive first the formulas for the single track. Equation (1.2) gives the random function involved. The covariance function for this is

$$C(\tau) = E\left\{g(s)g(s+\tau)\right\}$$

$$= E\left\{d(x_0 + s\cos\alpha, y_0 + s\sin\alpha) \quad d(x_0 + s\cos\alpha + \tau\cos\alpha, y_0 + s\sin\alpha + \tau\sin\alpha)\right\}$$

$$= R(\tau\cos\alpha, \tau\sin\alpha)$$

$$= R(\tau\cos\alpha, \tau\sin\alpha)$$

The spectral function for C (τ) is the function D (λ) that satisfies

$$C(\tau) = \int_{-\infty}^{\infty} e^{2\pi i \lambda \tau} D(\lambda) d\lambda$$

But we have from (1.6)

$$R(\tau \cos \alpha, \tau \sin \alpha) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{2\pi i (u\tau \cos \alpha + v \sin \alpha)} F(u,v) du dv$$

and if we take

$$u \cos \alpha + v \sin \alpha = \lambda$$

$$-u \sin \alpha + v \cos \alpha = \lambda$$

$$(1.11)$$

whence

$$\lambda \cos \alpha - u \sin \alpha = u$$

$$\lambda \sin \alpha + u \sin \alpha = v$$

and

$$|J| = \frac{\partial(\lambda, u)}{\partial(u, v)} = 1$$

the integral may be transformed to

$$\int_{-\infty}^{\infty} e^{2\pi i \lambda \tau} \left[\int_{-\infty}^{\infty} F(\lambda \cos \alpha - \mu \sin \alpha, \lambda \sin \alpha + \mu \cos \alpha) d\mu \right] d\lambda$$

from which we have

$$D(\lambda) = \int_{-\infty}^{\infty} F(\lambda \cos \alpha - \mu \sin \alpha, \lambda \sin \alpha + \mu \cos \alpha) du \qquad (1.12)$$

relating the linear spectra in various directions to the two dimensional spectrum.

For parallel tracks separated by a distance b the spectrum of each track is given by (1.12). The cross covariance of the two profiles is

$$H(\tau) = E\left\{d(x_0 + s\cos\alpha, y_0 + \sin\alpha) d(x_1 + s\cos\alpha + \tau\cos\alpha, y_1 + \sin\alpha + \tau\sin\alpha)\right\}$$

= R (
$$x_1 - x_0 + \tau \cos \alpha$$
, $y_1 - y_0 + \tau \sin \alpha$)

If (x_1y_1) is opposite (x_0y_0) on a line perpendicular to the track $x_1^-x_0^- = -b \sin \alpha \ y_1^-y_0^- = b \cos \alpha$. Since $H(\tau)$ is not an even function, its Fourier transform has both a real and an imaginary part. These are, respectively, its cosine and its sinc transform. They give the frequency--phase difference relationship between the two tracks. As above, this complex spectrum is the function $K(\lambda)$ such that

$$H(\tau) \int_{-\infty}^{\infty} e^{2\pi i \lambda \tau} K(\lambda) d\lambda$$

and (1.12) and (1.6) gives us

$$H(\tau) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{2\pi i \left[\left(-b \sin \alpha + \tau \cos \alpha \right) u + \left(b \cos \alpha + \tau \sin \alpha \right) v \right] F(u,v) du dv}$$

Making use again of the transformation (1.11)

$$H(\tau) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{2\pi i (bu + \tau \lambda)} F(\lambda \cos \alpha - \mu \sin \alpha, \lambda \sin \alpha + \mu \cos \alpha) du dv$$

whence

$$K(\lambda) = \int_{-\infty}^{\infty} e^{2\pi i b\mu} F(\cos \alpha - \mu \sin \alpha, \lambda \sin \alpha + \mu \cos \alpha) du$$

Formula (1.12) is a special case of this where b = 0. For each λ the integral is a line integral of the two dimensional spectrum along lines perpendicular to the direction (α) of the track and at a distance from the origin.

This work gives two possible derivations that may be made from the two dimensional spectrum. In both, the track of the vehicle is straight. When the track curves, the resulting roughness is not stationary unless the roughness of the field is the same in all directions. These spectra (1.14) for various α will give the conditions met by the vehicle in various directions and may be of aid in studying the effects of slowly varying non-stationarity.

In this introduction we have dealt with the theoretical quantities associated with the stochastic process. We have described them and indicated how they can be used. Our present problems are more

concerned with estimating these quantities, particularly the spectral function. Estimation of a two-dimensional spectrum is a simple extension of one dimensional spectral estimation. Since the formulas are not available in the literature we will present a fairly complete account of them here.

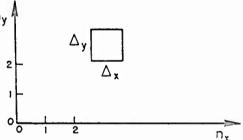
2. Relation of the Theory to the Practical Problem.

As in all statistical work we will use data from outcomes of the random scheme to make estimates of the desired characteristics of the probabilities. In our case the desired characteristic is the spectral density function of the covariance. Statistical estimation makes use of the fact that when repeated outcomes are averaged, the random part of the average tends to be small. This repetition may be accomplished with d(x,y) by taking the values at different places; the stationary property is, in effect, the property that the randomness repeats itself in different places. Actually a somewhat more stringent property, the ergodic property, is necessary so that averages over (x,y) tend to ensemble averages. We will assume as do most practitioners that this property holds.

It is most convenient for computation and for measurement to take data at intersection points of a rectangular lattice. Though the data we have taken has the separations equal in both directions, it is just as easy to deal with different separations Δ_x, Δ_y . Let us suppose that we make n_x by n_y observations, as indicated in the figure, at the points

$$(k\Delta_{x},j\Delta_{y})$$

$$k = 0,1,2...n_x-1$$
; $j = 0,1,2...n_y-1$



Supposing the smoothing of the profile to have been done, a topic we will discuss in Section 5, our data will consist of n_x by n_y values

$$d(k\Delta_x, j\Delta_y)$$

Only a part of the randomness is involved in this data. The covariances which underlie it are of the form

$$E \left\{ d(k \Delta_x, j \Delta_y) d((k+a) \Delta_x, (j+b) \Delta_y) \right\} = R(a \Delta_x, b \Delta_y)$$
 (2.1)

for

$$a = O_1 \pm 1, \dots \pm n_x$$
 $b = O_1 \pm 1, \dots \pm n_y$

No other values of the covariance function are used to "produce" the data and we cannot, therefore, hope to make direct estimates of any but these.

Supposing that we knew these values of R(x,y), what can we find out about the spectrum? This is a preliminary question and we must later refine the answer because we will not have the values of R(x,y) but only estimates of them. This, of course, will affect the way we answer the preliminary question.

The spectrum is related to these values of R(x,y) by equation (1.8)

$$R(a \Delta_{x}, b \Delta_{y}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \cos 2\pi (a \Delta_{x} u + b \Delta_{y} v) F(u, v) du dv \qquad (1.8)$$

In analogy to equation (1.7) we will later multiply the $R(a \Delta_x, b\Delta_y)$ by cosines and add. But we want to investigate the matter generally so instead of using cosines to obtain

$$\sum_{\mathbf{a}} \sum_{\mathbf{b}} \cos 2\pi (\mathbf{a} \Delta_{\mathbf{x}} \mathbf{u}_{\mathbf{0}} + \mathbf{b} \Delta_{\mathbf{y}} \mathbf{v}_{\mathbf{0}}) R(\mathbf{a} \Delta_{\mathbf{x}}, \mathbf{b} \Delta_{\mathbf{y}}) \Delta_{\mathbf{x}} \Delta_{\mathbf{y}}$$
(2.2)

let us use general coefficients to obtain

$$F^{*}(u_{o},v_{o}) = \sum_{a} \sum_{b} w(a,b;u_{o},v_{o}) R(a\Delta_{x}b\Delta_{y})$$

$$= \sum_{a} \sum_{b} w(a,b;u_{o},v_{o}) \iint \cos 2\pi (a\Delta_{x}u+b\Delta_{x}v)F(u,v)du dv$$

$$(2.3)$$

$$= \iiint \left\{ \sum_{a} \sum_{b} w(a,b;u_{o},v_{o}) \cos 2\pi (a\Delta_{x}u+b\Delta_{x}v) F(u,v) du dv \right\}$$

From this we see that F^* (u_0, v_0) is a weighted average of spectral values,

$$F^*(u_0, v_0) = \int \int W(u_0, v_0; u, v) F(u, v) du dv$$
 (2.4)

Our purpose is to obtain $F(u_0, v_0)$, therefore

$$W(u_{o},v_{o};u,v) = \sum_{a} \sum_{n} w(a,b;u_{o},v_{o}) \cos 2\pi (a\Delta_{x}u + b\Delta_{x}v)$$
 (2.5)

is ideally a Dirac or delta function of (u,v) centered at (u_0,v_0) . It is, of course, impossible to attain this ideal since W(`) is a finite Fourier series. An obvious W(`) to try is the finite part of the Fourier series for the Dirac function

$$\sum_{\alpha=-n_x}^{n_x} \sum_{b=-n_y}^{n_y} \Delta_x \Delta_y \cos 2\pi (\alpha \Delta_x u_0 + b \Delta_y v_0) \cos 2\pi (\alpha \Delta_x u + b \Delta_y v) \qquad (2.6)$$

which corresponds to (2.2).

We will not go into detail here about the form of W(u,v,:u,v). Lct us note, however, from (2.5) that it is even and periodic with a period of $1/\Delta_{v}$ in the u direction and $1/\Delta_{v}$ in the v direction so that if it has a sharp peak at (u_0v_0) it has identical peaks at $(+u_0+r/\Delta_x)$ $+ v_0 + s/\Delta_v$) for integers r,s. Referring to formula (2.4) we sec that the sharp peak of W(') at (u,v) picks up the values of F(u,v) near (u_0, v_0) while the other peaks get the values near their positions. $F*(u_0, v_0)$ is therefore made up of an average of the values of F(u, v)near (u_0, v_0) with the addition of similar averages for values near the other peaks of W('). These other averages are called "aliases". Their presence is, of course, unwanted. It is due to the fact that the data is taken at equally spaced points. The smaller the spacing $(\Delta_{_{\mathbf{Y}}}, \Delta_{_{\mathbf{Y}}})$ the farther apart are the peaks of W('). The usual spectrum "dies out" for large (u,v) and if $\Delta_{_{\mathbf{X}}},\Delta_{_{\mathbf{V}}}$ are small enough, the aliased values of F(u,v) will be at large enough frequencies that they do not affect the estimates. The periodicity of W(*) means that the function F*(u,v) of (2.3) is itself periodic. Actually the values outside the rectangle

$$\left(-\frac{1}{2\Delta_{\mathbf{x}}} \le \mathbf{U} \le \frac{1}{2\Delta_{\mathbf{x}}} \quad ; -\frac{1}{2\Delta_{\mathbf{y}}} \le \mathbf{V} \le \frac{1}{2\Delta_{\mathbf{y}}}\right) \tag{2.7}$$

are repetitions of those inside this rectangle. Moreover, the spectral function is even (1.9a). Consequently all the values of the estimated spectrum may be found from those within the rectangle $(-1/2\Delta_{_{\textstyle X}}\leqslant u\leqslant 1/2\Delta_{_{\textstyle X}};\ 0\leqslant v\leqslant 1/2\Delta_{_{\textstyle Y}})$ and these need be the only values reported. (For a one-variable spectral estimate the only values ordinarily presented are for positive frequencies out to the Nyquist frequency of $1/2\Delta$.)

The phenomenon of aliasing is a limitation on accuracy resulting from sampling at intervals. It arises from the form of Eqn. (2.5) and not from the specific coefficients $w(a,b; u_o, v_o)$. Another limitation of accuracy arises from our having only a finite number of data points. Equation (2.5) is a Fourier expansion with all coefficients zero for $a \ge n_x$ $b \ge n_y$. Since the coefficients may be calculated from $w(\cdot)$ by

$$w(a,b,u_ov_o) = \Delta_x \Delta_y \int_{-\frac{1}{2\Delta_x}}^{\frac{1}{2\Delta_x}} \int_{-\frac{1}{2\Delta_y}}^{\frac{1}{2\Delta_y}}$$
(2.8)

 $W(u_0, v_0; u v) \cos 2\pi (a \Delta_x u + b \Delta_y v) du dv$

it is clear that the ideal form for W('), (Dirac "delta" function), cannot be realized, otherwise all w() would be non-zero. In fact if W() is to be large near both positive and negative parts of $\cos 2\pi (a \triangle_x^u + b\triangle_x^v)$ for $a\ge n_x$, $b\ge n_x$ for the integrals to be zero. This indicates that the non-zero area of w(') should be at least a half period wide and a half period high or

$$\frac{1}{2 n_x \Delta_x} \quad \text{by} \quad \frac{1}{2 n_y \Delta_y} \tag{2.9}$$

As a check on this we note that the top line of (2.3) connects a finite number of R values, by a linear relation, to the values of F*. Under any but abnormal conditions an equal number of values of F* determine all the R's and hence also the other values of F*. If we space these values evenly, in the relevant area of frequencies, $(-1/2\Delta_x \leqslant u \leqslant 1/2\Delta_x)$, the spacing will be $1/2n_x\Delta_x$ by $1/2n_y\Delta_y$ as before,

3. Statistical Problem

In Section 2 we acted as though the values of $R(a\Delta_x, b\Delta_y)$ were known for $a=0,\pm 1,\ldots \pm n_x$ $b=0,\pm 1,\ldots \pm n_y$. Actually, we will estimate these values from the height data $d(k\Delta_x,j\Delta_y)$. The usual estimate for an expected value (2.1) is an average

$$\frac{1}{(n_x - |a|)(n_y - |b|)} \sum_{k} \sum_{j} d(k \Delta_x, j \Delta_y) d(k+a) \Delta_x, (j+b) \Delta_y = r_{a,b}$$
 (3.1)

where k and j range over all the pairs on hand. The actual limits of summation depend on the signs of a and b. Under the ergodic hypothesis this converges, as n_x and n_y grow larger, to $R(a\Delta_x,b\Delta_y)$ For values of a, b close to n_x , n_y , the sum has only a few terms and the estimate is not good. Therefore we must be content with estimates $r_{a,b}$ for $-m_x\leqslant a\leqslant m_x$, $-m_y\leqslant b\leqslant m_y$ where m_x and m_y are considerably smaller than n_x and n_y . The expected value of $r_{a,b}$ is $R(a\Delta_x,b\Delta_y)$. The different $r_{a,b}$ are averages of different numbers of terms, those of larger a and b have fewer terms and thus less precision. In general, also, the different r's are correlated.

In accordance with our findings in Section 2, we will try to estimate values of the spectral function which are evenly spaced throughout the relevant range. Since we will have covariance values only up to m_x , m_y , the spacing of spectral values is correspondingly larger. For mathematical convenience we will use a spacing of $1/(2m+1)\Delta$ rather than $1/2m\Delta$, where m is m_x in the x-direction and m_y in the y-direction. Thus we will try to estimate (see Eqn. (2.3)

$$F^* \left(\frac{\alpha}{(2 m_x + 1)} \frac{\beta}{(2 m_y + 1)} \right) \equiv F_{\alpha, \beta}^*$$
 (3.2)

$$\alpha = 0, \pm 1, \pm 2, ..., \pm m_x$$
, $\beta = 0, \pm 1, \pm 2, ..., \pm m_y$

Using the coefficients $w(a,b; u_0, v_0)$ of Eqn. (2.2) we can calculate the estimates

$$f_{\alpha,\beta} = \sum_{\alpha=-m_x}^{m_x} \sum_{b=-m_y}^{m_y} w \left(\alpha,b; \frac{\alpha}{(2m_x+1)\Delta_x}, \frac{\beta}{(2m_y+1)\Delta_y}\right) r_{\alpha,b}$$
 (3.3)

whose expected values are $F_{\alpha,\beta}^*$.

To approximate the distribution of the random variable $f_{\alpha,\beta}$ we will assume the Gamma distribution which has been found to fit well in this kind of work. The parameters A,K of the distribution will be estimated by the method of moments.

The following is the mathematical expression of our first assumption:

$$P\left\{f_{\alpha,\beta} < \Delta\right\} = \frac{A^{\kappa}}{\Gamma(\kappa)} \int_{0}^{\Delta} t^{\kappa-1} e^{-At} dt \qquad (A1)$$

As in Section II, Part 3 above, we have corresponding to

(3.2)

$$K = \frac{\left(F_{\alpha,\beta}^{*}\right)}{\operatorname{vor}\left\{f_{\alpha,\beta}^{*}\right\}}$$

$$A = \frac{F_{\alpha,\beta}^{*}}{4 \, \text{K}} \tag{3.4}$$

Confidence limits may be found as in Section II, Part 3 and interpreted by means of the table there.

The derivation of the formula for Var $\{f_{\alpha,\beta}\}$ which also gives cov $\{f_{\alpha,\beta},f_{\alpha',\beta'}\}=0$ is somewhat tedious and will be presented in an appendix to this section. Several simplifying assumptions are necessary and since these are relevant in practice we will discuss them below. The result is a formula for the degrees of freedom.

deg. fdm. =
$$2 \text{ K} \approx 2 \left(\frac{1-g}{g}\right)^2$$
 (3.5)

where $g = m_x / n_x = m_y / n_y$ is the ratio of the number of R values in each direction the number of profile values in each direction.

The assumptions made in the approximate calculation of var $\{f_{\alpha,\beta}\}$ and cov $\{f_{\alpha,\beta},f_{\alpha',\beta'}\}$ are as follows

- A.2 The profile height deviations d(x, y) of (1.1) are approximately Gaussian with means 0.
- A.3 The values of m_x and m_y are small in comparison to n_x and n_y so that the numbers of terms in the sums for different covariance estimates (3.1) are approximately the same.
- A.4 F(u,v) is essentially zero for (u,v) outside the range (2.8) so that aliasing does not occur. It is relatively constant within any of the small intervals (2.10).
- A.5 The non-zero parts of the functions W(') of Eqn. (2.4) are no wider than $(-\frac{1}{2n_{\chi}\Delta_{\chi}}, \frac{1}{2n_{\chi}\Delta_{\chi}})$ by $(-\frac{1}{2n_{\chi}\Delta_{y}}, \frac{1}{2n_{\chi}\Delta_{y}})$, so that their "overlap" is negligible.

We have little control over assumptions A.1 and A.2. Assumption A.5 is somewhat in conflict with the discussion of Eqn. (2.9) which says that the non-zero part of W(') must be at least that wide. To come near a compromise between A.5 and this discussion we must use a relatively long Fourier expansion (2.5) and choose the coefficients with care. This means a fairly large $m_{_{\rm X}}$, $m_{_{\rm Y}}$ also desirable for good resolution or closeness of spectral estimates. But A.3 requires an even larger $(n_{_{\rm X}}, n_{_{\rm Y}})$. The ratio g, of m to n should also be small to obtain a large degree of freedom (3.5). Assumption A.4 relates to the size of the spacings $\Delta_{_{\rm X}}$, $\Delta_{_{\rm Y}}$ in the profile data.

The desire for good resolution, statistical precision, and a wide range for the values of F(u, v) require a large number of profile values that are closely spaced. The expense of measuring the profile and handling the data must be balanced with these desires. Even if large amounts of data could be obtained and processed cheaply, the lack of stationarity and the inherent errors of measurement, which have not been taken into account here, would limit the accuracy and precision of the results in other ways so we will always be faced with the above dilemma. The principles outlined in this section will help us to deal rationally with it.

Appendix to Section 3

Derivation of degrees of freedom formula.

Let us consider two quantities of the form (3.3)

$$f_1 = \sum_{\alpha = -m_x}^{m_x} \sum_{b = -m_y}^{m_y} g_{\alpha b} r_{\alpha b}$$

$$f_2 = \sum_{\alpha = -m_x}^{m_x} \sum_{b = -m_y}^{m_y} h_{\alpha b} r_{\alpha b}$$

$$cov \{ f_1 f_2 \} = \sum_{a} \sum_{b} \sum_{a'} \sum_{b'} g_{ab} h_{a'b'} cov (r_{ab}, r_{a'b'})$$
 (X1)

In this sum the terms cov $(r_{ab}, r_{a'b'})$ are readily worked out using (3.1)

$$cov(r_{ab} r_{a'b'}) = \frac{1}{(n_x - |a|)(n_y - |b|)} \frac{1}{(n_x - |a'|)(n_y - |b'|)}$$

$$\sum_{i} \sum_{j} \sum_{i'} \sum_{j'} C(i, j; i', j')$$
(X2)

where the limits of the sums depend on a, b, a', b' and

$$C(i,j;i',j') = cov \left\{ d(i\Delta_x,j\Delta_y) d \left[(i+a) \Delta_x, (i+b) \Delta_y \right], \right.$$

$$d(i'\Delta_x,j'\Delta_y) + d \left[(i'+a') \Delta_x, (j'+b') \Delta_y \right] \right\}$$
(X3)

Using the Gaussian assumption A.2 we can evaluate this covariance from the formula

$$cov (MN,ST) = E \left\{ \left[MN - E(MN) \right] \left[ST - E(ST) \right] \right\}$$
$$= E \left\{ MNST \right\} - E \left\{ MN \right\} E \left\{ ST \right\}$$
$$= E \left\{ MS \right\} E \left\{ NT \right\} + E \left\{ MT \right\} E \left\{ NS \right\}$$

In this the Gaussian assumption is used only to evaluate the fourth moment.

Using this, (X.3) becomes

$$C (ij;i'j') = R \left[(i'-i)\Delta_{x},(j'-j)\Delta_{y} \right] R \left[(i'-i+a'-a)\Delta_{x},(j'-j+b'-b)\Delta_{y} \right]$$

$$+ R \left[(i'+a'-i)\Delta_{x},(j'+b'-j)\Delta_{y} \right] R \left[(i'-i-a)\Delta_{x},(j'-j-b\Delta_{y}) \right]$$

$$(X4)$$

We need, as our result, (X.1) in terms of F, so we will apply (1.6) to express the product of R's as an integral

$$R(ab) R(a'b') = \iint \int \int e^{2\pi i (au + bv + a'u' + b'v')} F(uv) F(u'v') du dv du' dv'$$
(X5)

Using (X.5) in (X.4) we obtain

$$C(ij;i'j') = \int \int \int EF(uv)F(u'v') du dv du' dv'$$
(X6)

where

$$\begin{split} E &= \exp 2\pi i \left[(i'-i) \Delta_{x} u + (j'-j) \Delta_{y} v + (i'-i+a'-a) \Delta_{x} u' + (j'-j+b'-b) \Delta_{y} v' \right] \\ &= \exp 2\pi i \left[(i'+a'-i) \Delta_{x} u + (j'+b'-j) \Delta_{y} v + (i'-i-a) \Delta_{x} u' + (j'-j-a \Delta_{y}) v' \right] \\ &= \exp 2\pi i \left[(i'-i) \Delta_{x} (u+u') + (j'-j) \Delta_{y} (v+v') - a \Delta_{x} u' - b \Delta_{y} v' \right] \times \\ &\times \left\{ \exp 2\pi i \left[a' \Delta_{x} u' + b' \Delta_{y} v' \right] + \exp 2\pi i \left[a' \Delta_{x} u + b' \Delta_{y} v \right] \right\} \\ &= 2\exp 2\pi i \left[(i'-i+\frac{a'}{2}) \Delta_{x} (u'+u) + (j'-j+\frac{b'}{2}) (v'+v) - a \Delta_{x} u' - b \Delta_{y} v' \right] \\ &= \cos 2\pi \left(a' \Delta_{x} \frac{u'-u}{2} + b' \Delta_{y} \frac{v'-v}{2} \right) \end{split}$$

Because F(u, v) is an even function, the integral (X.6) with u interchanged with v and u' with v' also equals C(i, j;i', j'). It has in place of E, a function identical with E except for the last two terms of the exponential. The average of the two integrals also equals C(i,j;i',j') and is like (X.6) except that in place of E it has

$$\begin{split} \exp \ 2\pi i \ \Big\{ (i'-i+\frac{a'}{2}) \ \Delta_{x}(u'+u) + (j'-j+\frac{b}{2}) \ \Delta_{y}(v'+v) \Big\} \cos 2\pi (a'\Delta_{x} \frac{u'-u}{2} + b'\Delta_{y} \frac{v'-v}{2}) \\ \Big[\exp 2\pi i (-a \ \Delta_{x} u'-b \ \Delta_{y} v') + \exp \ 2\pi i (-a \ \Delta_{x} u-b \ \Delta_{y} v) \Big] \\ = 2\exp 2\pi i \Big\{ (i'-i+\frac{a'-a}{2} \ \Delta_{x}(u'+u) + (j'-j+\frac{b'-b}{2}) \ \Delta_{y}(v'+v) \Big\} \\ \cos 2\pi (a \ \Delta_{x} \frac{u'-u}{2} + b \ \Delta_{y} \frac{v'-v}{2}) \cos 2\pi (a'\Delta_{x} \frac{u'-u}{2} + b'\Delta_{y} \frac{v'-v}{2}) \Big\} \end{split}$$

We will imagine this substituted into (X.6) and thence into (X.2) where we must sum the integrals. The sums are finite and can be taken into the integral where they are found to affect only the exponential factor of (X.7). This sum falls into the product of four sums of the form

$$\sum_{i'} e^{2\pi i \Delta_{\chi}(u'+u)i'} \qquad (X 8)$$

and a factor

$$2\pi i \left[\frac{a'-a}{2} \Delta_{x} (u'+u) + \frac{b'-b}{2} \Delta_{y} (v'+v) \right] \qquad (X 9)$$

The sum (X.8) is a geometrical series and its limits depend upon the value of a. It can be put into the closed form

$$\sum_{i'} e^{2\pi i \Delta_{x}(u+u')i} = e^{2\pi i \Delta_{x} \frac{u+u'}{2} (n_{x}-a'+1)}$$

$$= \frac{\sin 2\pi \Delta_{x} \frac{u+u^{1}}{2} (n_{x}-|a|)}{\sin 2\pi \Delta_{x} \frac{u+u^{1}}{2}}$$

The other sums may be worked out similarly. Their product consists of four factors that are sine ratios and an exponential factor which will be found to be the reciprocal of (X.9). Thus, the formula (X.2) for the covariance of two γ values becomes

$$\frac{2}{(n_{x}-|a|)(n_{x}-|a'|)(n_{y}-|b|)(n_{y}-|b'|)} \int \int \int \int$$

$$\frac{\sin \pi \Delta_{\mathbf{x}}(\mathbf{u} + \mathbf{u}') (\mathbf{n}_{\mathbf{x}} - |\mathbf{a}'|)}{\sin \pi \Delta_{\mathbf{x}}(\mathbf{u} + \mathbf{u}') (\mathbf{n}_{\mathbf{x}} - |\mathbf{a}|)} (\mathbf{x} \cdot \mathbf{10})$$

$$\sin \pi \Delta_{\mathbf{x}}(\mathbf{u} + \mathbf{u}')$$

$$\frac{\sin \pi \, \Delta_y (v + v^i) (u_y - |b^i|)}{\sin \pi \, \Delta_y (v + v^i)} \qquad \frac{\sin \pi \, \Delta_y (v + v^i) (u_y - |b|)}{\sin \pi \, \Delta_y (v + v^i)} \qquad (X 10 con.)$$

$$\cos 2\pi \, (a \, \Delta_x \frac{u^i - u}{2} + b \, \Delta_y \frac{v^i - v}{2}$$

$$\cos 2\pi (a' \Delta_x \frac{u'-u}{2} + b \Delta_y \frac{v'-v}{2} F(u,v) F(u' v') du dv du' dv'$$

Before substituting this long expression into (X.1) where it must be multiplied by coefficients and summed, let us examine the integrand. The sine ratios may be combined with the first factor to give four functions of the form

Each of these has a spike at ξ = 0 of height 1. and also spikes at $\xi = \frac{K\pi}{A}$ for the integer K. The factor N affects only the width of the spikes. If we adopt assumption A.3 the dependence of the sine ratio terms on a,b,a',b' largely disappears. The product of four factors is approximately a product of two squares, and the indices of summation in (X.1) affect only the cosine terms of (X.10). Let us take the sum and the factors $g_{ab} h_{a'b'}$ inside the integral whence we have the sum

$$\sum_{a} \sum_{b} g_{ab} \cos 2\pi \left(a \Delta_{x} \frac{u' - u}{2} + b \Delta_{y} \frac{v' - v}{2}\right) \sum_{a'} \sum_{b'} h_{a'b'} \cos 2\pi \left(c' \Delta_{x} \frac{u' - u}{2} + b' \Delta_{y} \frac{v' - v}{2}\right)$$

$$= G(\frac{u'-u}{2}, \frac{v'-v}{2}) H(\frac{u'-u}{2}, \frac{v'-v}{2})$$

where G and H are two of the functions W defined in (2.4), i.e., W for two different pairs (u_0v_0) .

We may now express (X.1) as an integral of the spectral function. Let us denote the "common value" of

$$\begin{split} & n_{x} - |a| = n_{x} - |a'| = N_{x} \\ & n_{y} - |b| = n_{y} - |b'| = N_{y} \\ & \text{cov} (f_{1} f_{2}) = 2 \int \int \int \int \left[\frac{\sin \pi \Delta_{x} (u + u') N_{x}}{N_{x} \sin \pi \Delta_{x} (u + u')} \right]^{2} \left[\frac{\sin \pi \Delta_{y} (v + v') N_{y}}{N_{y} \sin \pi \Delta_{y} (v + v')} \right]^{2} \\ & G \left(\frac{u' - u}{2}, \frac{v' - v}{2} \right) H \left(\frac{u' - u}{2}, \frac{v' - v}{2} \right) F (u v) F (u' v') du dv du' dv' \end{split}$$

Let us change variables so

$$\begin{cases} \frac{1}{2} (u^{i} + u) &= w \\ \frac{1}{2} (u^{i} - u) &= w^{i} \\ \frac{1}{2} (v^{i} + v) &= z \\ \frac{1}{2} (v^{i} - v) &= z^{i} \\ \frac{1}{2} (v^{i} - v) &= z^{i} \\ \frac{1}{2} (u^{i} - v) &= z^{i} \\ \frac{1}$$

$$\operatorname{cov}(f_1,f_2) = 2 \iiint \left(\frac{\sin 2\pi \Delta_x N_x w}{N_x \sin 2\pi \Delta_x w} \right)^2 \left(\frac{\sin 2\pi \Delta_y N_y z}{N_x \sin 2\pi N_y z} \right)^2 G(w',z') H(w',z')$$

$$\times F(w-w^{1}, z-z^{1}) F(w+w^{1}, z+z^{1}) 4 dw dw^{1} dz dz^{1}$$

Carrying out the integration first with respect to w and z we note that the squared sine ratios have spikes for w (or z) equal to $0, \pm \frac{1}{2\Delta_x}$ (or $0, \pm \frac{1}{2\Delta_y}$) etc. and that except for these spikes whose width is about $2 \cdot \frac{1}{2N_x\Delta_x}$ (or $2 \cdot \frac{1}{2N_y\Delta_y}$) the values are quite small. Using assumption A.4, F is relatively constant inside the peak at 0 and has very small values in the region of the other peaks. The first two integrations

$$\begin{split} \delta \int\!\int \; G\left(w^{\scriptscriptstyle I}z^{\scriptscriptstyle I}\right) \; H\left(w^{\scriptscriptstyle I}z^{\scriptscriptstyle I}\right) \left[\; F(w^{\scriptscriptstyle I}z^{\scriptscriptstyle I})\; \right]^2 \; \left\{\; \int\! \frac{\frac{1}{4\Delta_x}}{\frac{1}{4\Delta_x}} \left(\frac{\sin 2\pi \Delta_x N_{x}w}{N_x \sin 2\pi \Delta_x w}\;\right)^2 dw \right. \\ \left. \int\! \frac{\frac{1}{4\Delta_y}}{\frac{1}{4\Delta_y}} \left(\frac{\sin 2\pi \omega_x N_{x}w}{N_x \sin 2\pi \omega_x w}\;\right)^2 dw \right\} dw^{\scriptscriptstyle I} dz^{\scriptscriptstyle I} \end{split}$$

The two integrals of the sine ratio may be evaluated from Fejer's integral (Titschmarsh "Theory of Functions", p. 413). The result is

cov (f₁, f₂) =
$$\frac{2}{\Delta_x \Delta_y N_x N_y} \int \int G(w^1 z^1) \left[F(w^1 z^1) \right]^2 dw^1 dz^1$$
 (X 12)

From assumption A.5 when G and H are two different w functions their non zero parts do not overlap and so the covariance of two spectral estimates is zero.

The variance of f is the covariance of f with itself and formula (X.11)

$$var(f) = \frac{2}{\Delta_x \Delta_y N_x N_y} \int \int \left[G(w^i z^i) F(w^i z^i) \right]^2 dv^i dz^i \qquad (X 13)$$

Let us recall that our purpose is to compute the degrees of freedom using formula (3.5). Our result (X.12) will be in the denominator and the numerator will have

$$\left[E \left\{ f \right\} \right]^2 = \left[\int \int G(w^i, z^i) F(w^i, z^i) dw^i dz^i \right]^2 \tag{X 14}$$

This follows from taking the expected value on both sides of (3.3) then using (2.2). In both (X.12) and (X.13) we will take G(w'z') to be an abbreviation of $w(u_0, v_0; w', z')$ of (2.4)

To attain approximations of the two integrals we will make strong use of assumption A.5 and also assume an idealized rectangular form for G. This form proceeds from our discussion at the end of Section 2; it is that G is 0 in the relevant region of frequencies (w',z') except in two rectangles of dimensions $\frac{1}{m_\chi\Delta_\chi} \times \frac{1}{m_\chi\Delta_\chi}$ centered at (u_0v_0) and $(-u_0-v_0)$. The total volume of these rectangles should be unity so the height of G will be $\frac{1}{2}$ $(m_\chi m_\chi \Delta_\chi \Delta_\chi)$. With this (X.12) and (X.13) work out to be approximately

$$\operatorname{var} \left\{ f \right\} = \frac{2}{\Delta_{x} \Delta_{y} N_{x} N_{y}} \left[F(u_{o} v_{o}) \right]^{2} (m_{x} m_{y} \Delta_{x} \Delta_{y})$$

$$\left[E\left\{ f \right\} \right]^{2} = \left[F(u_{o} v_{o}) \right]^{2}$$
(X 15)

The degrees of freedom formula (3.5)

d. f. =
$$2K = \frac{2[F(u_0 v_0)]^2}{[F(u_0 v_0)]^2 \frac{2m_x m_y}{N_x N_y}} = \frac{N_x N_y}{M_x M_y}$$

Now if we recall the definition (X.11) of $\,N_{_{\scriptstyle X}}\,$ and $\,N_{_{\scriptstyle Y}}\,$ we have as an approximation

d.f. =
$$\frac{n_x - m_x}{m_x} = \frac{n_y - m_y}{m_y}$$

which, when we use the fixed fraction g for both m_x and m_y we get the formula (3.6). This approximation is probably as good as the assumptions justify but more exact results may be found by approximating more closely (X.13) and (X.14) using the actual W function instead of the rectangular idealization.

4. Filtering

The assumptions at the end of Section 3 are used in both the statistical work and, as we pointed out in the discussion of them, in relating the theoretical work to the practical. We will discuss here assumptions A.4 and A.5 beginning with the latter.

The shape of the function W(') should be two symmetrically placed peaks on a field of zeros. Though considerations in Section 2 indicat that the peaks should be as narrow and high as possible around the central point, the statistical considerations expounded in equations (X.14) and (X.15) indicate that the variance of f increases with the narrowness of the peak. Thus we are led to seek a box shaped W(') function. The choice of this function is the subject of a good deal of literature. Here we will follow some of the early work--that given in Blackman and Tukey, "Measurement of Power Spectra". This will take the direction of using as standard the W(') given by (2.6) and modifying it by filtering. Let us note first several formulas concerning finite Fourier expansions.

Suppose we have an expansion C(u,v) with coefficients $c_{a,b}$

$$C(u,v) = \sum_{\alpha=-m_x}^{m_x} \sum_{b=-m_y}^{m_y} c_{\alpha b} e^{2\pi i (\alpha \Delta_x u + b \Delta_y v)}$$

and we want to select evenly spaced values of (u, v) to "represent" W(u, v). Let us take

$$\begin{aligned} \mathbf{u} &= \mathbf{0} \ , \ \pm \ \delta_{\mathbf{x}} \ , \ \pm \ 2\delta_{\mathbf{x}} \ , \dots , \pm \ \mathbf{m}_{\mathbf{x}} \delta_{\mathbf{x}} \\ \\ \mathbf{v} &= \mathbf{0} \ , \ \pm \ \delta_{\mathbf{y}} \ , \ \pm \ 2\delta_{\mathbf{y}} \ , \dots , \pm \ \mathbf{m}_{\mathbf{y}} \delta_{\mathbf{y}} \end{aligned}$$

$$\mathbf{C}_{\mathbf{j},\mathbf{k}} \equiv \mathbf{C}(\mathbf{j} \delta_{\mathbf{x}}, \mathbf{k} \delta_{\mathbf{y}}) = \sum_{\alpha = -m_{\mathbf{x}}}^{m_{\mathbf{x}}} \sum_{b = -m_{\mathbf{y}}}^{m_{\mathbf{y}}} \mathbf{c}_{\alpha,b} \ \mathbf{e}^{2\pi \mathbf{i} (\Delta_{\mathbf{x}} \delta_{\mathbf{x}} \alpha \mathbf{j} + \Delta_{\mathbf{y}} \delta_{\mathbf{y}} b_{\mathbf{y}})}$$

Multiplying both sides by $2\pi i (\Delta_x \delta_x \overline{a} j + \Delta_y \delta_y \overline{b} k)$

and summing on

(j,k)

gives

$$\sum_{j=-m_x}^{m_x} \sum_{k=-m_y}^{m_y} C_{j,k} e^{2\pi i (\Delta_x \delta_x \overline{a} j + \Delta_y \delta_y \overline{b} k)}$$

$$= \sum_{\alpha} \sum_{b} c_{\alpha,b} \sum_{j} \sum_{k} e^{2\pi i (\Delta_{x} \delta_{x} j(\alpha - \overline{\alpha}) + \Delta_{y} \delta_{y} k(b - \overline{b}))}$$

The second sums are easy to work out. The exponential is readily separated into a product and gives a product of two geometric series, each of the form

$$\sum_{j=-m_x}^{m_x} e^{2\pi i \Delta_x \delta_x j(\alpha-\overline{\alpha})} = \frac{e^{2\pi i \Delta_x \delta_x m_x (\alpha-\overline{\alpha})} - e^{2\pi i \Delta_x \delta_x (m_x+1)(\alpha-\overline{\alpha})}}{1-e^{2\pi i \Delta_x \delta_x (\alpha-\overline{\alpha})}}$$

$$= \frac{\sin \pi \, \Delta_{x} \, \delta_{x} (2 \, m_{x} + 1)(\alpha - \overline{\alpha})}{\sin \pi \, \Delta_{x} \, \delta_{x} (\alpha - \overline{\alpha})}$$

This formula holds for $a \neq \overline{a}$. For $a = \overline{a}$ the sum is evidently $(2m_x + 1)$.

Now if we take $\Delta_x \delta_x = \frac{1}{2m_x + 1}$ the numerator is zero for all the (integral) values of $(a - \sqrt{a})$ while the denominator is not. From this

$$\sum_{j} \sum_{k} e^{2\pi i \left[\Delta_{x} \delta_{x} j(\alpha - \overline{\alpha}) + \Delta_{y} \delta_{y} k(b - \overline{b}) \right]}$$

$$= \begin{cases} (2m_{x} + 1)(2m_{y} + 1) & \text{if } \alpha = \overline{\alpha}, b = \overline{b} \\ 0 & \text{otherwise} \end{cases}$$

This means we can solve (4.1) for the $c_{a,b}$ in terms of the $C_{j,k}$ when

$$C_{j,k} = C \left[\frac{j}{(2m_x+1)\Delta_x} \cdot \frac{k}{(2m_y+1)\Delta_y} \right]$$

Let us display the two formulas

$$C_{j,k} = \sum_{\alpha = -m_{x}}^{m_{x}} \sum_{b = -m_{y}}^{m_{x}} c_{\alpha,b} e^{2\pi i \left(\frac{\alpha j}{2m_{x}+1}, \frac{bk}{2m_{y}+1}\right)}$$

$$= \sum_{\alpha} \sum_{b} c_{\alpha,b} \cos 2\pi \left(\frac{\alpha j}{2m_{x}+1}, \frac{bk}{2m_{y}+1}\right)$$

$$c_{\alpha,b} = \frac{1}{(2m_{x}+1)(2m_{y}+1)} \sum_{j} \sum_{k} C_{j,k} e^{-2\pi i \left(\frac{\alpha j}{2m_{x}+1} + \frac{bk}{2m_{y}+1}\right)}$$

$$= \frac{1}{(2m_{x}+1)(2m_{y}+1)} \sum_{j} \sum_{k} C_{j,k} \cos 2\pi \left(\frac{\alpha j}{2m_{x}+1} + \frac{bk}{2m_{y}+1}\right)$$

$$(4.2)$$

These relations are analogous to the usual reciprocal relations for Fourier series, e.g., (2.5) and (2.8). More important to the problem of filtering is the following convolution formula. Suppose

C and \overline{C} are given by c and \overline{c} through formula (4.2), then their convolution is defined as

$$\frac{1}{(2m_x+1)(2m_y+1)} \sum_{\alpha=-m_x}^{m_x} \sum_{\beta=-m_y}^{m_y} C_{j-\alpha,k-\beta} \overline{C}_{\alpha\beta} = T_{jk}$$
 (4.4)

where C is regarded here as a periodic function having values outside the range $(-m_X, m_X; -m_y, m_y)$. To be specific, $C_{m_X}+1, k^{\pm}$ $C_{-m_X,k}$; $C_{m_X}+2, k^{\pm}C_{-m_X}+1, k$; etc. If we transform T_{jk} by a formula such as (4.3)

$$\uparrow_{\alpha b} = \frac{1}{(2m_{x}+1)^{2}(2m_{y}+1)^{2}} \sum_{j} \sum_{k} \sum_{\alpha} \sum_{\beta} C_{j-\alpha k-\beta} \overline{C}_{\alpha \beta} e^{2\pi i \left(\frac{\alpha j}{2m_{x}+1} + \frac{bk}{2m_{y}+1}\right)}$$

$$= \frac{1}{(2m_{x}+1)^{2}(2m_{y}+1)^{2}} \sum_{\alpha} \sum_{\beta} \overline{C}_{\alpha \beta} e^{2\pi i \left(\frac{\alpha \alpha}{2m_{x}+1} + \frac{b\beta}{2m_{y}+1}\right)} \times$$

$$\times \sum_{j} \sum_{k} C_{j-\alpha k-\beta} e^{2\pi i \left(\frac{\alpha(j-\alpha)}{2m_{x}+1} + \frac{b(k-\beta)}{2m_{y}+1}\right)}$$

Since in the second sum both "C" and the exponential term are periodic with period ($2m_X+1$) by ($2m_y+1$) we can rearrange the summands to achieve the sum that could be gotten by erasing α and β . This gives us, from (4.3),

$$t_{ab} = c_{ab} \overline{c}_{ab} \tag{4.5}$$

So the Fourier coefficients of a convolution are the products of the coefficients of the two functions being convolved.

If we define a convolution of coefficients as

$$\sum_{\alpha = -m_x}^{m_x} \sum_{\beta = -m_y}^{m_y} c_{\alpha - \alpha, b - \beta} \overline{c}_{\alpha \beta} = y_{\alpha b}$$
 (4.6)

then we may derive similarly,

$$\sum_{a} \sum_{b} y_{ab} e^{2\pi i \left(\frac{aj}{2m_x + 1} + \frac{bk}{2m_y + 1}\right)} = C_{jk} \overline{C}_{jk} = Y_{jk}$$
 (4.7)

so that the expansion whose coefficients are convolutions is the product of two expansions.

We will apply (4.4) and (4.5) to the situation where $\bar{c}_{a,b} = r_{a,b}$ of (3.1) so that $c_{j,k}$ are spectral estimates of the type (2.2) whose filter function is (2.6). With this we may take $c_{a,b}$ so that

$$c_{ab} \cos 2\pi \left(\frac{aj}{2m_x+1} + \frac{bk}{2m_y+1}\right) = w(a, b; j\delta_x, k\delta_y)$$
 (4.9)

where w are the coefficients for the spectral estimates (3.3), and correspond to a desirable filter function $W(\cdot)$, (2.5). Then from (4.4) and (4.5) and (3.3) we have,

$$f_{\alpha,\beta} = \sum_{a} \sum_{b} w(a,b;\alpha \delta_{x},\beta \delta_{y}) r_{ab} = \sum_{a} \sum_{b} c_{ab} \cos 2\pi \left(\frac{a\alpha}{2m_{x}+1} + \frac{b\beta}{2m_{y}+1}\right) r_{ab}$$

$$= \frac{1}{2m_{x}+1} + \frac{1}{2m_{y}+1} \sum_{a} \sum_{b} c_{\alpha-a}, \beta-b \overline{c}_{ab}$$

$$(4.10)$$

From this we have a fundamental relation. C is the "raw spectrum" of the covariances $r_{a,b}$ and the coefficients C are used to "smooth" it by taking a running average. By this we may achieve a variety of W(') functions. Note that the actual W(') function is a transform (2.5) of the product of the C's with cosines and is therefore a convolution of the type in (4.8). All the filters used in practice are given either in terms of the running average weights

C or in terms of the coefficients c which modify the covariances before making a Fourier transformation. Corresponding to the two sides of the convolution theroem, we have two ways of computing the spectrum. We may first compute a "raw spectrum" or cosine transformation of $\mathbf{r}_{a,b}$ and then smooth it by a running average. This is convenient when only a few of the running average coefficients $\mathbf{C}_{j,k}$ are non-zero. Or we may modify the $\mathbf{r}_{a,b}$ by multiplying them by the factors $\mathbf{c}_{a,b}$ then take the cosine transformation. This is done when the c's are easy to compute.

An application of the convolution relations is to smoothing the profile which was discussed for the one-dimensional case in Section II, Part 4. We should note that the work of the section deals primarily with the data and not with the mathematical model of randomness. The various sequences \mathbf{r}_{ab} , $\mathbf{f}_{\alpha\beta}$, are handled as even, periodic sequences. In actual computation of the sums we do not go back to the beginning after exhausting the table so that the numerical results are affected by "end effects" which cause them to depart somewhat from the formulas presented here.

Section 5. The two-dimensional spectral computer program.

The program operates using data as outlined in the preceding sections. The general operation of the program involves the steps:

- (1) The area data is smoothed by taking differences from one fitted quadratic polynomial.
 - (2) The covariances of the smoothed data are computed.
- (3) A preliminary spectrum of these covariances is computed.
- (4) A trigonometric polynomial is fitted to the reciprocals of the spectral values, its coefficients are used to make a running average smoothing of the covariances.
- (5) A new spectrum of the smoothed covariances is computed.
 - (6) The spectrum is smoothed by the Hamming method.
- (7) The spectrum is corrected using the reciprocals of the polynomial (4) as factors.

The program was checked by running some sea surface data from the report by Pierson mentioned in Section I. The results checked very well except for the latter part of the program.

Recently this part has been made to operate correctly.

The model of this program was the one (Section II, Part 5) which we were using on the linear data at the time the specifications were sent to the programmer. Our later experience with the linear program indicates that the quadratic polynomial method of smoothing is not enough for our purposes, and that the prewhitening steps may be accomplished better by using an improved method of smoothing. In order to complete the two-dimensional analysis, the program is being modified.

6. Preliminary Spectral Estimate

Table 5 contains a preliminary spectral estimate from our present program using the Aberdeen data. These results are reproduced to illustrate some of the comments we have made regarding two-dimensional results. Complete studies of this two-dimensional data will be presented in a later report.

	٠,		F	TABLE 1					•	
		ABERDEEN -	SEN - 1500	0 Ft.	TWO F(TWO FOOT SPACING	CING			
0	C	2	4	9	80	_10	13	14	91	18
0	,7	0	5		6	0.0	0			-1.17
20	-1-25	-1.31	-	-1.06	o	-0.30	•	0.72	0.72	0.55
	4.0	.2		0	0	0.0	0			
	-2	.5	6.	2.2	1.9	6	.5			£.
	8		4	9.	8	0	0	•	•	0
	0	4.	. 2	0.	•	~	•2	•		6
	8	0	6.	•		7	0.		•	7
	0		4.	•	.5	9	7.		1.88	30
	9.	3	•	•	80	S	• 5	•	1.	0
200	-2	9.	3.		6	•0		0.87	0.93	8
200	.7	1.	9	•	•	•	• 6	-0.53	•	.2
	.5	9.	4.	•	.5	~	6.	0.73	•	9.
	9	0	0	•	-2	~	6.	0.68	0.36	G
	.2	.5	3	•	6	C	6.	-0.92	-0.92	6
	6.	6.	6.	•	6.		4.	-0-22	0.35	8
	0	4.	8	•	-	.5	•2	0.89	99.0	9
	6	5	3	•	.6			2.03	2.01	6
	0	6	8		4.	-2	3.	-0.05	-0.25	
	6	-2	₹.	•	•	3	8	1.92	1.83	w.
400	6.	5.	. 7	•	•	C	7	0.11	0.04	0
	0	8		•	6	4.	.5	1.55	1.68	
	8	6.	0	•	-2	3	.2	2.24	2.21	7
		- 2	~	•	3	4.	4	2.51	2.58	•
	-	. 7	8	•	0	3	.5	3.57	3.40	6
	6.	. 7	.2	•	4	7	7	1.10	1.23	5
	8	8		•	9.	5	4.	1.22	1.15	_
	6	-		•	.5		4.	3.24	3.05	6.
	6.	6.	. 1	•	.2	4.	.5	3.76	4.03	8
		4.	7	•	.2	7	0	3.27	3.49	7
600	0	8	.5	•	4	4.	4	2.31	2.19	7
000	2	.2	.2	•	7	-2	0	2.00	1.89	8
			.2		4.	4	6	2.79	3.22	.2
	6	3		•	.2		7	3.04	5.65	3
	0	6.	.7		4.	7		1.42	2.00	7
	7.	. 7	6		8	8		0.88	0.72	0.74
	6		•	•	9.	E,	.5	2.70	2.70	4.
	2		1	•	9.	ŝ	3	1.26	1.37	1.60
	9.	-	•		9.	1.52	7.	06.0	79.0	9.
	.5	.2	4.	•	4.	6.	9.	1.93	2.00	6.
	0	0	• 1	•	0	7	8	1.84	1.76	3-1

TABLES AND CHARTS

Each table of results consists of five parts. The first part is the raw profile heights given in feet and hundredths above the surveyors' datum. The heights are presented in reading order, left to right and down. The numbers of feet from the beginning for certain of the heights in the first column are given on the left margin. On the top margin are given the numbers of feet to add to the position of an item in the left column so as to obtain the positions of the heights in other columns of the same row. The second parts of the tables, labled A, are indexed similarly. These parts contain smoothed heights.

The third tables, labled B, give autocorrelations of the smoothed data in ft. 2 for the various lag numbers, k. The fourth tables, labled C, give the spectral density estimates of the smoothed data in ft. 2 for various frequencies, λ , in cycles/ft. The last tables, D, give the corrected spectral estimates, and are labled similarly.

After making the calculations, several small errors were found in the raw data of Aberdeen, Knox, and Yuma. The number of the errors was small, at most fifteen in the Yuma data, and the sizes were on the order of a few tenths of a foot. The effects on the calculations were deemed to be minor so they were not done over. The raw data portion of each table is presented in a correct version, but the other parts are as calculated.

800	1.72	1.81	1.81	1.80	1.82	1.88	1.92	1.87	1.86	1.84
	1.83	1.75	1.67	1.46	1.40	1.30	1.00	0.68	0.76	0.8
	0.70	0.95	1.17	1.48	1.11	0.80	0.52	0.28	-0.20	-0.24
	-0.24	-0.10	0.10	0.65	0.91	0.78	0.62	0.45	0.38	0.35
	0.34	0.40	0.56	0.83	1.20	1.45	1.46	1.42	1.07	6.0
	0.81	0.70	0.58	0.48	0.35	0.15	90.0	-0.03	-0.15	-0.28
	-0.62	-0.92	-1.12	-1.30	-1.76	-1.79	-1.63	-1.58	-1.53	-1.2
	-1.09	-0.86	-0.66	-0.06	-0-	20.0	0.27	0.36	0.46	0.4
	0.13	0.23	0.05	-0.19	-0.25	-0.47	-0.72	-1.05	-0.83	-0.7(
	-0.85	-0.93	-1.12	-1.27	-1.26	-1.18	-1.16	-1.17	-1.40	-1.05
1,000	-0.95	-0.95	-1.18	-1.40	-1.74	-1.84	-1.38	-1.62	-1.83	-1.6
	-0.98	-1.18	-0.81	-0.58	-0.31	-0.08	93.0	0.34	0.58	9.0
	0.74	96.0	1.05	1.11	1.15	1.21	1.15	1.16	1.30	1.3
	1.33	1.37	1.33	1.47	1.61	1.03	0.10	-0.23	-0-17	-0-5
	-0.29	-0.14	0.33	-0.27	C.37	1.04	1.05	6.95	0.86	0.7
	09.0	09.0	0.37	0.20	-0.03	-0.19	-0.35	-0.42	99.0-	9.0-
	-0.74	-0.75	-0.13	-0.66	-0.63	-0.42	90.0	0.61	1.09	6.0
	1.17	1.17	1.19	1.23	1.20	1.23	1.27	0.74	1.30	1.3
	1.39	1.44	1.41	1.43	1.41	1.40	1.39	1.38	1.50	1.5
	1.53	1.57	1.59	1.63	1.71	1.73	1.78	1.79	1.81	1.8
1200	1.87	1.88	1.97	1.98	2.04	2.09	2.15	2.17	2.19	2.2
	2.27	2.39	2.41	2.40	2.32	2.32	2.31	2.25	2.25	2.3
	2.30	2.35	2.33	2.18	5.04	1.90	1.85	1.67	1.54	1.5
	1.57	1.59	1.74	2.01	2.12	2.17	2.19	2.15	2.17	2.2
	2.25	2.25	2.13	2.09	2.09	2.06	2.05	1.99	2.02	2.1
	2.17	2.15	2.13	2.09	2.07	2.12	2.13	2.07	2.12	2.14
	2.09	1.98	1.86	1.63	1.37	2.07	2.07	0.90	0.85	1.7
	1.50	1.27	-0.52	-0.73	-0.81	-0.63	-0.53	-0.50	-0.33	-0-1
	0.19	1.05	0.63	0.65	0.63	0.53	14.0	0.46	0.42	0.3
	0.54	0.77	66.0	1.17	2.07	1.87	1.75	0.74	0.27	-0.0
1400	0.05	0.05	-0.13	-0.18	-0.27	-0.18	-0.15	-0-83	-0.63	-0.5(
	-0.32	0.04	0.15	0.29	76.0.	0.62	0.10	19.0	0.45	0.2
	-0.06	-0.26	-0.85	-0.61	-0.67	-1.00	-1.07	-0.91	-0.81	-0.6
	-0.39	-0-17	-0.13	-0.23	-0.61	-0.41	-0.93	-1.01	-0.91	-0.6
	-0.35	-0.01	0.27	0.49	09.0	0.65	0.75	0.83	0.89	0.8

TABLE 1 CONT.

TABLE 1A SMOOTHED ABERDEEN - 1500 Ft. TWO FOOT SPACING

					1									-!	95										-										1				-	
80	22222	•			0.53889	-0.33222	-0.27111	0.13222	-6.17444	-0.16111	-0.27222	-0.29000	-0.07333	-0.00000-	0.12556	-0.30222	-0,20111	-0.31667	0.09556	-6.30776	0.0000	-0-14222	-0.02222	0.03556	-0.03222	-0.35444	-C-00883	0.26333	0.12667	-0.17222		.3558		.5277	-0322	-	0.05667	• 6744	-1322	-6-17639
91	13000	0.50222	-36	87780.0-	0431000	-0.56889	-0.27009	0.14778	-C.26333_	-0.24000	-0-46444	-0.24889	0.06869	-0.00000	-0.00567	-0.31000	0.32889	-0.21222	0.48889	-0.19657	0.01556	-0.02222	-0.01333	0.31667	-0.25111	-0.33556	0.05778	0.46111	0.39333	-0-11000		0.53444	-0688	4	+0466	.4233	-0-16778	-2133	.3277	-0.05111
14	4777C C-	4 .		-0.29657	0.37333	-0.58778	-0.20000		-0.32889	-0.46778	-0.74444	-0.24000	0-16444	-0.00333		-0.15333	0.29222	-0.05889	0.56556	-0.03444	-0.04889.	0.03111	-0.03778	0.40657	-0.32839	-0.24889	0.34657		.1033	-0.02889	0.05000	.3065			.1244			.0922	.4577	-0.02556
12	0000	• •	0	4	0.27222	30	•		-0.27889	0.71111	•	+4410-0-	0.22111	-0.010.0-		0.00667	0.43556	-0.06111	0.56778	-0.32778	9.05667	•	0-	0	0	-0-	d	•		0.02333	0-02000	-0.32667			0	5	-0-19222			-0.00778
10	0000	• •	.0544	.5433	0.24333	-0.60333	-0.16778		-0.24222	0.05444	-0.37333	.1222	0.27889	-0.05000	-0.27000	0.22556	0.20556	-0.06778	0.25778	-0.28444	0.03778	0.17333	-0.02000	0.15333	-0.38556		79960-0	0.01000			-1344	.5533	0-14111	-0.53111	-0.21111	0.46778	-0.04444		-0-13444	0.17111
œ	, , c	366	-0-11889	-0.44556	0.17667	-0.61000	-0.09889	0.31778	-0.07889	0.33444	-0.12111	0.27667	0.34778	-0.12444	-0.31889	0.41333	0.13222	-0.17667	0.18222	-0.24000	0.13556	0.09556	-0.01111	-0.06889	-0.36111	0.05222	0.05889	-0.11000	-0.04667	-0.08111		-0.43222	0-03556	-0.17556	-0-31556	0.02000	-0.01778		4	0.02889
œ		-0.38667		-0.64444	0.13000	-0.78444		•	-0.08000	0.50667	0.08444		0.23111	-0.19444	-0.14778	•	0.09889	-0.23556	•	-0.12556	•	0.05667	•	•	-0.28333	•	-0-06444	-0.08444	-0-12444		0-12333	-0.17778	0.11444	0.01389	0-03667	-0.19222	-0.08111	0.19333	.4871	0.03444
4	0000	-0.32778	55	-0.46889	0.19444	0.31444	0.45778	-0.30889	-0.13333	0.39778	0.29222	0.34111	0.07111	-0.30667	00070.0-	0.51556	-0.03667	-0.13444	-0-37000	0.10333	0.31333	-0.14111	0.02333	-0.05444	0.10111	0.43000	0.10778	-0.04556	-0.29889	-0.13889	0.05222	-	0.07444	-0.06778	-	-0.51222	-0.23111	0.11111	-0-13444	0.08667
6		-6.35556	-0.06778	-0.26111	0.31567	-0.44389	0.61778	-0.39444	0.00778	0.17556	0.23889	0.71444	0.12111	-0-18444	-0.02000	0.31333	-0.46778	0.16778	-0.51111	0.02889	0.12556	0.03111	-0.02444	-0.07111	0.25222	0.18333	-0.09667	-0.20111	-0.09444	-0.06556	0.02556	-0.26556	0.06333	-0.05556	0-16222	-0.66333	0.03889	0.14778	-0.30667	-0.01778
0	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	31889	.06889	0	0.33889	4				-0.05778	-0.04444	0.84000	-0.27000	-0.07778	0.00333	0.11778	-0.59778	0.61778	-0-42444	-0.19000	-0-46333	-0.00556	-0.15889		0.21222	0.22222	-0.31111			0.08778	0.00778	0.06889	0.26556	-0.21222	0.54333	-0.49556	0-04444		9	0.05778
	c	200	,	•	i						200										400				1						009				1				1	

TABLE 1A CONT.

800	87730.0-	0,02333	0.01444	0.001/18	-0.01222	0.03444	0.07222	0.02889	0.03333	0.05333
	.2555	0.0166	.254	.6111	.353	1477	0000	.0988	.4255	414
	-0.42567	-0.31555	-0.153	. 3244	.515	3200	11111	.0922	.1522	1733
	-0.23000	-0.25222	-0.214	0000.	.230	4144	.3785	.3233	.0011	55.50
	-0.05667	-0.92111	0.0100	.0322	.022	.0566	.0000	.0766	.1344	18//
	0.06000	-0.03144	-0.0566	1110.	.398	.3522	.1833	1622	.1833	.0622
	-0.13111	-0.09000	-0.0955	.2944	.270	.1588	.2538	0.2500	.5822	3448
	0.06444	0.24667	0.1366	.0122	.205	.0777	.0522	0.2733	.9500	.2333
	0.23111	0.20222	0.0244	.1933	.111	.0122	.0100	0.0188	.2588	1366
1000	0.26889	0.34444	0.1377	1780.	.307	.3355	.1271	.1122	.3877	.2866
0001	0,16333	-0.13111	0.0022	.0058	.013	.0655	1100.	0.0755	.1344	0000
	0.00000	0.07222	0.0922	.0877	.057	.0500	.0511	0.0766	.0388	.053
	-0.01111	0.04222	0.1200	.4300	.738	.3322	.4133	.5300	.3933	.2446
	-0.23111	-0.13556	0.1422	.5822	.063	. 4966	.4244	.2344	.1333	100
	0.01111	0,14839	0.0533	.0355	.043	.0622	.0733	.0188	.1555	.0955
	-0,11667	-0.11889	-0.1522	.2233	.333	.3711	1044	.2344	.5088	.2022
	0.20222	0.07222	0.0133	1760.	.024	.0311	.0466	.5111	.1088	0833
	0.07333	0.10339	0.0066	.0266	900.	.0288	.0488	0.0766	.0255	.0011
	-0.00333	-0.39111	-0.0255	1110.	.027	.0133	.0300	.0077	0010.0	0100
1200	-0.01444	-0.03389	0.0111	.0133	.002	.0144	.0311	0.0044	.0233	.0433
	-0.00889	0.09222	0,0366	.0900	.004	.0088	.0086	.0622	0.0544	0211
1	0.04222	0.13778	0.1622	1144.	.332	.3277	.3644	3.6733	0.2855	.1108
	0.19444	0.17333	-0.1000	.1133	.152	.1500	.0765	0.0200	.0133	0960
	0.07889	0.07333	-0.6155	.0355	.013	.0344	.0355	0.0977	0.0722	0.177
	0.07667	0.04889	0.0133	.0322	.046	.0066	.0233	0.6200	.0555	1244
	6.15778	0.05444	-0.0655	.1666	.233	.4544	.5017	0.5366	0.4077	1255
	6.79554	0,35555	-0.7584	8183	.657	.3083	.088b	6.2333	0.2144	0.1566
	0.00333	0.74556	0.2177	.1500	.071	.0511	.0544	0.6300	0.1577	0.2411
	-0.26889	-0.13444	-0.1177	.0265	.940	. 3056	.7666	0.1388	,4644	0.5444
1400	-0.19667	0.01778	-0.0633	.0083	.017	.1333	.2044	0.4344	0.3311	0.2633
0011	-0.16667	0.10773	0.0477	.0433	.217	.2100	.3011	.3166	.2233	0.1044
	-0.01222	-0.02333	-0.420	0288	.023	.2377	.2435	.2088	.1622	100
	0,16556	C.31222	0.3544	.2766	.077	.1555	.3444	4317	.3933	0.2955
	-0.06118	0.03667	0.1311	.2077	.142	.0544	.0611	.087	306.	

TABLE 1B

AUTOCORROLATION ABERDEEN

SMOOTHED DATA

k		1-1500 ft.		1-750 ft.		7,50-1500 ft.
4		0830.c		0.0891	and I had the day the bearing	0.0369
	and the second or other party of the second or other party or othe	0.0379		0.0489		0.0268
*		0.0135		0.0244		U.0026
		0.0073		0.0005	-	0.0141
.3		0.0149		0.0118	-	0.0181
4		0.0171		0.0252	-	0.0092
5		0.0225_	_	0.0318	-	0135
		0.0219	_	0.0276	-	0.0162
,		0.0203	_	0.0263	-	0.0142
		0.0132	_	0.0180	-	0.0079
9		0.0029		0.0138		0.0086
10		0.0002		0.0084		0.0093
11		0.0026	_	0.0018	A~*	0.0073
13		0.0031		0.0041		0.6020
13		0.0040		0.0101		_1.027
-		0.0062		0.0118	-	0.0003
25		0.0057_		0.0117		_0.0016
17		0.0044		0.0112	-	0.036
1.2		(023_		0.0082	1 programme and the first control of	0.0044
14		0.0003		0.0045		0.0041
26		0.0020_		0.0046	*4. 1	0.0000
87		0.0018		0.0040		0.0004
21 22	12	0.0001		0.0018	-	_0.0011_
		0.0015	que peu les s	0.0047		0.0031
23	_	0.0018		0.0073		0.0055
		0.0014		0.0074		0.0068
26		0.0042	_	0.0092		0.0024
2,		0.0644	_	0.0068	-	0.0011
21		0.0042		0.0048		0.0036
2.1	_	0.0021		0.0030	_	0.0722
*		0.0002_		0.0001	_	0.0019
31	-	0.0007		0.0005	-	0:0054 0:0053
32	_	0.0015	-	0.0015		
33		0.0002	_	0.0010	un-de	0.0022
3		0.0044		0.0034		_ (`.)O'+!}
35		0.0090		0.0084		0.0111
34		0.0087_		0.0089		0.0173
- ,,	upo upo	0.0076	· ·	0.0097		0.11042
		0.0026		0.0101	-	
				0.0069	•	0.7254
7.5	-	0.0008	M as are not de-	2.0062		0.1345
4,70		0.0065		0.0093	-	0.,334
4.1	_			0.0130		V.)180
42		0.0156		0.0127	-	0.7162
65		0.0147	-	0.0082	-	0.0139
. 44	_	0.0116	_	3.0102		0.0092
45		0.0013	Carrie	0.0063		0.0106
4,1		0.0023	_	0.0029		0.0107
44 j		0.0049		0.0011		0.0082
48		0.0062			1.46	1. 737
45	ameter or a	0.0057	-	0.3080		117
51		0.0099		•		

-98-TABLE 1C

ABERDEEN

POWER SPECTRAL DENSITY

1 OWER SI EC	SMO	OTHED DATA	1122112211
λ	1-1500 ft.	1-750 ft.	750-1500 ft.
0.00000	0.00239	0.01560	0.00479
0.00500	0.02014	0.02946	0.01434
0.01000	0.10575	0.10283	0.10202
0.01500	0.27612	0.30310_	0.23916
0.02000	0.53145	0.69460	0.37944
0.02500	1.07736	1.69098	0.49692
0.03000	1.50238	2.20492	0.70204
0.03500	1.28241	1.71860	0.79930
0.04000	1.33669	1.44279	1.30112
0.04500	1.07281	1.02973	1.15177
0.05000	0.84050	0.89131	0.77151
0.05500	0.69880	_0.10279_	0.72524
0.06000	0.50667	0.47875	0.50718
0.06500	0.50175	0.57522	0.39657
0.07000	0.41796	0.50602	0.34049
0.07500	0.37456	0.32,400	0.42598
0.08000	0.35928	0.23250	0.44397
0.08500	0.44999	_ 0.21263	0.69028
0.09000	0.43022	0.24646	0.60723
_0.09500	0.42759	0.23562	0.60828
0.10000	0.45689	0.25063	0.67066
_0.10500	0.47007	0.19130	0.35643
0.11000	0.23433	0.20663	0.20713
_0.11500	0.23103	0.27404	0.18538
	0.20347	0.20156	0.21/07
0.12000	0.19745	0.17218	0.22506
0.12500		0.14506	0.1 794
0.13000	0.15222	0.14854	0.10353
_0.13500	0.16626	2.13139	0.15790
0.14000	0.17385	0.16407	0.17463
_0.14500	0.17077	0.13023	0.21114
0.15000	0.16942	0.08670	0.16351
.0.15500	0.12592	0.11308	0.15:73
0.16000	0.14245	0.10067	0.17185
. 0 . 16500	0.13742	0.05909	0.15352
0.17000	0.11710	0.03707	0.17734
_0.17500	0.12539	0.05179	0.14161
0.18000	0.10080	0.05556	0.2 701
_0.18500	0.13301	0.03813	0.25531
0.19000	0.15245	0.00046	0.23216
0.19590	0.15353		0.32725
0.20000	0.22136	_0.11709_ 0.12716	0.24235
0.20500	0.13520		0.21327
0.21000	0.17411	0.13782	
0.21500	0.10536	0.15758	0.17378
0.22000	0.10584	0.15678	0.17666
0.22500	0.15657	0.17531	0.13.80
0.23000	0.14259	0.13922	0-14-34
0.23500	0.14512	0.12477	0.15547
0.24000	0.15285	0.12633	0.13639
0.24500	0.15040	0.11010	0.13736
0.25000	0.14710	0.10104	0.106(8
0 1 2 7 0 0 0			

TABLE 1D

	TILDLIL	117	
POWER S	SPECTRAL DENSIT	Y.	ABERDÈEN
λ	0-1500 ft	0-750 ft	750-1500 ft
0.000			
0.005	116.19168	169.96063	82.73032
0.010	39.36861	38. 28155	37.98000
0.015	21.11489	23. 17805	18.28856
0.020	13.57748	17.74564	10.20489
0.025	12.09444	18.98294	5.57842
0.030	8.86158	13.00153	4.13964
0.035	4.56807	6.12182	2.84718
0.040	3.10900	3.35578	3.02627
0.045	1.78086	1.70935	1.91193
0.050	1.06373	1.12804	. 97642
0.055	. 71340	. 71747	. 74044
0.060	. 43838	. 43153	. 43882
0.065	. 38445	. 44074	.30385 .24070
0.070	. 29503	. 35719 . 21837	. 28710
0.075 0.080	. 25244 . 23881	. 15454	.32169
0.085	. 30367	. 14349	. 46582
0.090	. 30238	. 17322	. 42679
0.095	.31971	. 17625	. 45502
0.100	. 37008	. 21111	. 54323
0.105	. 24321	. 16980	. 31638
0.110	. 22928	. 20222	. 25154
0.115	. 24893	. 29527	. 19974
0.120	. 24534	. 23721	. 25546
0.125	. 24989	. 21791	. 28584
0.130	. 20277	. 19323	. 21305
0.135	. 22719	. 20297	. 25086
0.140	. 23721	. 24750	. 22909
0.145	. 22681	. 21791	. 23725
0.150	. 21441	. 16481	. 26721
0.155	. 14953	. 10295 . 13040	. 19416 . 18634
0.160	. 15732 . 14081	. 10315	. 17609
0.165 0.170	. 11172	.06592	.15601
0.175	. 11172	. 06636	.15838
0.110	. 09087	.05171	.12730
0.185	. 10908	.04638	. 16978
0.190	. 12246	. 03062	. 21317
0.195	. 13490	.04516	. 22572
0.200	. 17865	. 09646	. 26669
0.205	. 15426	. 10591	. 20186
0.210	. 15188	. 11967	. 18519
0.215	. 15227	. 14424	. 15906
0.220	. 16123	. 15242	. 17175
0.225	. 16222	. 18163	. 14484
0.230	. 15723	. 15351	. 16026
0.235	. 16916	. 14544	. 19288
0.240	. 18621	. 15390	.21976
0.245	. 18860	. 13798	. 23732
0.250	. 18870	. 16584	.08363

	ø								Þ	
				TA	TABLE 2					
			FORT	ORT KNOX -	1800 Ft.	TWO FO	TWO FOOT SPACING	ING		
	0	2	4	9	8	10	12	14	16	
0	0	0	0.	. 1	.2		4.			
•	9	4.	3	8	9.	5	7	.3	5	4.
	0.24	2.		0.	-	٠,4	æ	∞	.2	. 7
	4	0	•	3	1.	-	.3	5	.2	.2
	0	5	~	8	.2	Τ.	4.	• 2	.1	•2
	-0-	0	9	0.1	0	0.1	•2	0.3	0.1	.2
	2	.2	-	0.2	0.2	0.3	•2	0.3	0.3	0.
	4	4.	0.4	0.4	0.8	9		0.6	4.	0.3
	4	4.	0.7	0.7	0.7	9.0	4.0	0.5	9.0	9.
	9	5	. 7	0.6	9.0	0.5	.5	0.5	0.5	0.5
200	9	•	0.5	0.7	0.7	9.0	1.0	9.0	0.1	6.
•	9		1.2	0.9	6.0	0.9	6.0	6.0	0.9	1.0
	5	-1.05	-0.98	6.	-0.72		8	-0.72	-0.71	1.0
	9	0	0.8	0.8	6.0	0.8	0.7	9.0	9.0	0.5
	9	7.	9.0	0.5	0.5	0.6	0.7	9.0	9.0	8
	8	8	0.7	0.7	0.5	0.4	7.	0.7	6.0	1.0
	2	.2	1.3	1.1	1.3	1.3	1.3	1.2	1.3	1.3
	2	.2	1.3	1.3	1.3	1.3	1.4	1.4	1.4	1.3
	0	٤.	1.3	٠,٠	1.4	٠,	1.4	1.5	1.5	9.
	5	. 5	1.5	1.6	9.	1.4	1.5	1.5	1.3	.2
400	2	-	0.1	.3	4.0	r.	• 5	. 7	• 2	.3
	2	0.	0.	• 2	• 1	0	0.0	0.1	0.2	0.2
	-	. 1	~	• 2	• 2	r.	• 6	• 6	. 7	1.
	7	. 7	• 6	5	٠,		• 2	4.	4.	3
	2	0.2	٠4	. 1	7.	٠,	• 6	3		7.0
	9	0.8	0.8	0.8	6.0	5	1.0	9.0	0.8	0
	1	9.0	0.5	0.4	9.0	0.7	0.7	0.7	0.8	0.8
	Φ,	0.9	्		-	⁻.	0	6.	\$	0.8
	7 .	9.4	0.5	0.5	0.4	•	0.7	0.1	0.7	6.7
	7	0,7	0.7	0.3	0.3	0.3	0.5	٦.	0.5	1.1
009	4	1.6	l. 6	1.5	1.4	٠,	1.3	1.2	1.0	1.1
	2	1.3	1.4	1.5	1.5	S.	1.5	1.4	1.1	1.0
	0	0.8	0,8	0.8	1:1	1.4	1.7	2.0	1.8	1.8
	∞	1.0	6.	0.7	0.3	0.3	.8	6.0	1.1	.5
	9.	1.6	1.6	1.5	1.2	5	9.0	4.0	0.2	0.2
	9	0.7	1.0	1.2	1.3	1.3	1:1	0.1	0.8	9.
	4.	0.4		0.5	0.7	ಎ	8.0	173	6.	
	8	9.0	4.0	0.3	0.4	0.5	0.7	6.3	1.0	
	-1.01	-0.71	-0.53	-0.39	-0.21	-0.15	-0.11	-0.22	-0.33	94.0-
	5	0		0.8	· •	9.0	0.5	0.4	0.4	•

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-1.30	-1.22	-1.02	1	1	-1.30	-0.40	-1.00	-1.23	-0.74	-0.10	-0.72	-0.11	-0.27	-0.12	0.48	2	-0.28	N	-1.42	-0.38	19.0-	99.0-	-0.77	91	99.0-	-0.76	44.0-	05.0-	0.21
-1.45	-1.25	-0.87	-0.62	-0.82	-1.17	-0.37	96.0-	-1.24	69.0-	-0.13	-0.77	0.01	-0.28	-0.13	0.38	-0.07	-0.35	-0.22	-1.42	-0.38	49.0-	-0.65	-0.70	-0.82	-0.72	17.0-	04.0-	-0.44	0.15
-1.63	-1.22	-0.70	-0.55	-0.87	-1.10	44.0-	-0.92	-1.20	89.0-	-0.21	-0.76	0.03	-0.30	-0.17	0.28	-6.22	-0.45	-0.22	-1.41	-0.37	-0.61	-0.59	19.0-	-0.84	-0.73	-0.80	-0.34	94.0-	0.16
-1.72	-1.08	-0.54	-0.54	-0.95	-0.92	-0.55	-0.92	-1.21	99.0-	-0.37	-0.65	0.08	0.18	-0.19	0.08	-0.22	-0.32	-0.20	-1.02	-0.34	-0.54	-0.55	-0.63	-0.92	-0.82	-0.86	-0.38	44.0-	01.0
-1.79	-1.02	-0.43	-0.59	-1.02	-0.78	-0.72	76.0-	-1.19	-0.67	-0.46	-0.42	-0.04	0.13	-0.24	-0.07	-0.24	-0.20	-0.19	06.0-	-0.38	-0.44	09.0-	-0.62	-1.11	-0.89	-0.84	-0.50	-0.36	0.04
-1.72	06.0-	-0.42	99.0-	-1.06	-0.72	-0.82	-0.92	-1.12	-0.70	-0.76	-0.22	-0.12	0.05	-0.24	-0.32	-0.14	-0.12	-0.17	-0.80	-0.53	-0.34	-0.68	-0.66	-0.92	-1.02	-0.78	-0.54	-0.37	-0.02
-1.62	-0.87	-0.47	-0.72	-0.96	-0.66	-0.93	-0.86	-1.08	-0.77	-1.00	-0.02	-0.35	-0.04	-0.25	-0.30	-0.03	0.06	-0.15	-0.66	-0.74	-0.37	-0.66	-0.54	-0.94	-1.12	-0.84	-0.59	-0.24	-0.08
	-0.90	• 6		9	1.	-		0	6.	6.			•																-0.19
-1.06	-0.96	-0.86	-1.13	-0.90	-0.64	-1.22	-0.68	76.0-	-1.05	-0.90	0.03	-0.66	-0.12	-0.32	-0.31	0.42	0.10	-0.28	-0.40	-1.24	-0.46	-0.68	-0.59	-0.94	-1.24	-0.82	-0.70	-0.35	-0.29
-0.17	-1.11	-1.04	-1.22	-0.84	-0.67	-1.32	-0.52	-1.02	-1-19	-0.92	0.10	-0.57	-0.15	-0.28	-0.22	0.53	-0.02	-0.28	-0.30	-1.34	-0.43	-0.72	-0.64	-0.84	66.0-	-0.61	-0.74	-0.40	-0.34
800))									1000								,		1200									

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1400	0.18	3 0.06	-0.04	-0.14	-0.34	-0.44	-0.59	-0.69	-0.55	-0.35
	-0.23	1	-0.03	0.01	-0.01	-0.16	-0.32	-0.49	-0.54	-0-41
	-0.28	1	-0.25	-0.35	-0.37	-0.42	-0.44	-0.50	-0.44	94.0-
	-0.56	1	-0.75	-0.76	-0.85	-1.11	-1.14	-1.04	69.0-	-0.59
	-0.5	1	-0.37	-0.24	-0.32	-0.34	-0.54	-0.81	-0.91	-0.84
	-0.48	400,000 00	-0.49	-0.58	-0.57	-0.76	-0.72	-0.72	-0.64	-0-44
	-0.24	1	-0.44	-0.40	-0.54	-0.73	-0.82	-0.89	-1.08	-0.79
	-0.86	1	-0.70	-0.76	-0.85	+6.0-	86.0-	-1.06	-1.19	-1.12
	-1.4	1	-1.47	-1.56	-1.69	-1.74	-2.44	-2.44	-2.09	-2.19
	-1.7		-2.46	-2.36	-1.64	-1.99	-1.64	-1.46	-1.44	-1.72
1600	-2.93	1	-3.19	-3.09	-2.97	-2.72	-2.64	-2.46	-2.41	-2.29
	-2.2	1	-2.44	-2.49	-2.64	-2.69	-2.34	-2.37	-2.39	-2.42
	-2.5	1	-2.59	-2.69	-2.59	-2.79	-2.74	-2.64	-2.69	-2.74
	-2.7	1	-2.72	-2.94	-3.23	-3.36	-3.54	-3.63	-3.66	-3.74
	-3.76	1	-3.79	-3.49	-3.49	-3.39	-3.41	-3.42	-3.43	-3.39
	-3.66	'	-3.62	-3.54	-3.21	-3.27	-3.39	-3.39	-3.59	-3.69
	-3.9	١	-4.06	-3.94	-4.04	-3.74	-4.25	-3.69	-3.97	-4.17
	-4.0	1	-4.39	-4.59	4.56	-4.47	-4.27	-4.01	-4.06	-3.88
	-3.8	1	-3.79	-3.45	-3.59	-3.77	-3.84	-4.41	-4.24	-4.29
	-4.3		-4.34	-4.42	95.5-	-4.41	44.4-	-4.59	-4.57	-4.64
1800	-4.5	1	-5.09	-5.19	-5.09	-5.04	-5.02	-5.34	-5.42	-5.34
	-5.0	1	-4.03	-4.00	-3.69	-3.91	-3.89	-3.99	-3.97	-3.79
	-3.92	1	-3.87	-3.56	-3.59	-3.57	-3.79	-4.17	-4.33	-4.56
	41.4-	1	-5.21	-5.19	-5.09	69.4-	-4.29	-3.78	-3.48	-3.59
	-3.84	4 -4.13	-4.34	-4.39	62.4-	-4.41	-4.29	-3.84	-3.61	-3.67

		Ω	SMOOTHED KNOX - 1900 FT.	T - VONV	JA OOF	LWOFO	TWO FOOT SPACING	2	
Ó	81	4	9	80	10	12	14	16	18
0.00778	0.01000	0.00111	.0333	0.04111	-0.38111		-6.00067	0.23000	0.2211
0.04889	-0.2133	0	0.33222	.13	.04	-0.36222	-0.09778	0.19333	0.17889
.0133	-0.03	.1688	.3677	-0.55444	-0.12556	0.09778	9	0.30111	0.65778
.3322	-0.16	-0.18222	0	-0.51222	-0.05444	0.18333	0.44333	0.25667	.0.19567
. 1044	-0.275	.338	0	-0.17889	-0-23444	0.14333	-0.01000	-0.00778	C-13e89
	0.01	0-04444	-0.1200	0.17111	-0.00444	-0.11778	-0-13000	0.00556	-6.0333
-0.00667	-0.04		0.04667		-0.07444	-0.03111	-0.04778	-0.00444	0.3400
55	0.0244			-0.23667	.3233	-0-14333	-0.06222	0.12889	L-25556
0.15556	0.1522	.1877	-0-	.12			0	-0.00R89	-0.05444
-0.05222	0.05	0	0	-0.02111	•	0.07222	0.04111	0.00778	0.02111
200 -0.01556	0	0.08778		22		-0.07333	÷	0.58778	-0.15111
0.17222	-0.31	-0.34667	0.04000	0.01778	•	0.02333	0-	.0111	0-17000
-0.01333	-0.16	-0.11111	-0.0833	0.12111	0.07889	•	00050.0	.0922	0.09778
-0.08222	-0.19	-	-0.02222	-0.09333	-0.03333	77770-0-	0.1100	-088e	0-11000
17	-0.09	-0.04000	0.090.0	0.07111	3388	-0.02333	0.0733	.0566	-6.66667
.1500	0	-0.07000	-0.00	•	0.30444	0.08667		-0.00667	-0.04556
-0.13000	-0.13	-0.16556	0.12	-0.02333	-0.03000		-0.01111	0.00333	-0.01333
0.03889	0.04	-0.01333	•	.01	-0.01222	-0.05778	-0.08444	-0.07556	0.00883
0.31778	ĭ	0.0522		-0.04444	•	0.00444	-0.01657	0.00444	9250-0-
0.03778	ĭ	C	-0.04667	-0.0333	0.02849	-0.15444	-0.33222	-0.33000	-0-37589
400 0.58222	0.52222	0	0.1200	-0.11111	-0.11444	-0.11778	-0.3422	0.06111	-0.09667
-0.02778		0	0	0.14667	0.10667	-0.01222	7	1	-0.18555
-0.14667	ĭ	0	-0.02778	-0.08111	. 0.04667	0.12667)		0.09000
0.11444	0	0	-0.02556	-0.06444	-0.12000	-0.14111			0.64775
-0.08778	ĭ	0	-0.18444	-0.16222	0.15889	0.51222	0		-0.19222
-0.27444	1	1	-0.03667	-0.06444	-0.34333	0	0.25111	-0-	-0.00222
-0.05444			0.25111	0.01000	-0-05333	-0-	-0.02444	•	0.04222
94444	Ö	J	0	-0.13556	-0-14111	0	-0.03444	-6.01222	-0-0-3111
-0.01111	_	0	0	0.20444	.0.00889	1	-0.11222	7	-0.06667
-0.05111	-0-10000	9		0.10444		0-11111	0.55111		11:00
6000.34111	ĭ	-0.316	-0.17000	-0.03222	-0.00778	.0322	0.10556	•	0.15es9
0.06667		-0-	-0-15222	-0-12778	-0.13111		-0.16667	0.02869	11160-0-
0.14222	_	0	0.35556	0.11444	2		-0.48333	-0.29778	-6.32600
-0.45000	0.14	0.08889	0.26556	0.57111	0.55667		99	-0.04333	-0-2822
-0.27556	-0.26000	-0.29333	-0.27444	-0.17333	-0.03778	0.22111	0.32667	0.42889	0.46567
0-11556	0.04	-0.19000	-0-27667	-0-28889	•26	-0-17444	• 0666	0.02333	0.07333
0.24667	0.224	0.19556	•	2	•	•	-087	-0.16778	-0.1533
-0-14778	0.036	.208	.2455	0.24222	•16	.0155	-0.15333	.2677	-0.34000
-0.27444	-0.0366	0.05444	0.10333	0.19667	0.19556	• 2	0.11444	0.04883	-0.01444

CONT.	
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TABLE	

	,														10)4	-															
1306	U368	1150	.2005	.1805	.0155	.1168	.3633	-0222	.1153	50.	30.	04.	.23	.0811	.0011	.0322	.266	-0144	.0533	.0377	99000	.1768	.1108	0	-0.06061	0		0	0	C.09222	97760.0	-0.24000
0070.	.0855	.0122	8960	.2122	0	0.	0	0	-	070	1265	0855	1622	0033	0722	0577	2690	0322	0588	0233	•	•	.0788	•	•	-0744	.0566	7	.2377	.0688	655	-0.33899
.0555	0.13000	-0.01111	.0788	.1633	.0288	.0322	.0677	.1066	.2277	.1155	-0.1744	C	O	1	1	()	7	9	0	0.03444	0.10222	0.11778				•	•	5	-0.21667	.035	-0.24111	-0.26667
.158	.136	.056	.04	- 10	00.	.04	.09	.06	.19	.21	•	•	•	•	•	•	•	•	•	-0744	.0533	-0044	.0533	.0477	.0955	•0666	.0388	-0.21556	-0.07222	-0.01889	-0.30667	-0.00111
.2288	.1186	1011	.1100	.0088	-0844	.0500	* 0944		0344	1511	2155	0233	0588	2077	0400	0200	0544	2033	0611	0288	0355	11911	.0533	.0188	0.00889	.0144	0.03111	-0.09778	0.06111	-0.03222	-0.26889	77701 0
7	•	-0.13000	.100		-0.08778		0.11444	-0.12444	0.07444	0.12222	0.12222	-0.00556	-0.23111	-0.16333	0.01111	0.04444	0.03000	0.16600	0.13444			0	.040	•	•	•	.03	.05	.191	.002	.012	1844
	0.09111	.0433	.1122	-0044	00060-0-	0.010.0	•	•	•	0	.0133	.0000	.15556	.11444	.14889	.07111	.03889	.06556	.07556	.02222	00333	7997	-0.12667	-0.04778	-0.00778	•	.0333	.0588	0.19000	.014	.0522	2288
0.14111	-0.02333	-0.06111	0.02111	75760-0-	-0.03111	0.04333	0.01667	-0.21333	0.16000	-0.14778	0	0	1300	0.02889		.0155		0.03222	-0.02000	0.05333	-0.03889	-0.04222	-0.28667	-0-13667	0	0.05778	-0.01000	0.06556	0.15667	0.11889	-0.00444	0 08556
-0.02333	-0.26556	-0.05556	95560.0	-0.14333	-0.02444	0,06111	-0.05333	-0-10889	0.18111	-0-15444		.04556	-0.09222	0.21333	0.13556	-0.01667	0.07778	-0.19889	-0.05889	-0.03000	0.04667	-0.07667	-0.23444	-0.05111	-0.02333	79900-0	-0.04778	0.08222	0.18778	0.11444	0.03444	11170 0
-0.13111 -	-0.36111	-0.04689	0.08444	-0.22111-	0.08889	-0.01889	-0.13333 -	-0-10667	0.19556	0.00333	6	-0.05222	0.01111	0.28778	0.01778	-0.00222	0.10000	-0.22778 -		-0.07667		-0.03000-	0.01889 -	0.15667	-0.02333 -	-0.04111	Ċ			0.08333	0.04000	00000
	-0.02333 0.14111 0.23000 0.24111 0.22889 0.158e7 0.05556 -0.07000 -0.15e6	-0.02333 0.14111 0.23000 0.24111 0.22889 0.15867 0.05556 -0.07000 -0.1506 -0.26556 -0.02333 0.09111 0.10567 0.11889 0.13667 0.13000 0.08556 0.0388	-0.02333 0.14111 0.23000 0.24111 0.22889 0.15869 0.05556 -0.07000 -0.1566 -0.26556 -0.02333 0.09111 0.10667 0.11889 0.13667 0.13600 0.08556 0.0388 -0.05556 -0.06111 -0.04333 -0.13000 -0.10111 -0.05667 -0.01111 0.01222 0.0477	-0.02333 0.14111 0.23000 0.24111 0.22889 0.15869 0.05556 -0.07000 -0.1566 -0.02556 -0.07000 -0.1566 -0.025556 -0.02333 0.09111 0.10567 0.11889 0.13667 0.13600 0.08556 0.0388 -0.05556 -0.06111 -0.04333 -0.13000 -0.10111 -0.05667 -0.01111 0.01222 0.0477 0.09556 0.02111 0.11222 0.10000 0.11000 0.04556 -0.07889 -0.09889 -0.20355	-0.02333 0.14111 0.23000 0.24111 0.22889 0.15859 0.05556 -0.07000 -0.1556 0.02556 -0.07000 -0.1556 0.0358 -0.26556 0.0358 0.0358 -0.06111 0.10667 0.11889 0.13667 0.13600 0.08556 0.0358 -0.0358 -0.06111 -0.04333 -0.13000 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1500	0.14556	0.24000		0.03889	0.02667	-0.16778	-0.14667		-0.09222	0.08889
	0.26444	0.10556		0.14444	0.07556	-0.05333	-0.09222	-0.12000	-0.27667	0.03778
	-0.01889	0.03444		0.10222			0.02556	0.02111	-0.03000	0.11889
	-0.12778	0.02667		0.13389			-0.51111	-0.39444	0.05556	0.03000
		-0.36111	-0.39000	-0.36000	0.27667	•	0.32000	0.56556	0.67778	.0.55889
1600		-0.54222	-0.55111	-0.33778	-0.14111		0.03333	C.12222.	0.10000	0.16667
	0.15778	0.08333	-0.00000	-0.05444	-0.19333	-0.22889	0.14000	0.13556	0.12667	0.10222
		-0.10889	0.01222	-0.06000	0.07000	-0.10778	-0.05111	9.06000	0.01333	0.00222
	0.06111	0.17000	0.24000	0.12444	-0.06333	-0.08111	-0.14222	-0.10222		-0.08889
		-0.24111	-0.16556	0.10778	0.07333	0.13222	0.08667	0.05556		0.11556
		-0.22000	-0.15333	-0.07778	0.27444	0.21778	0.12333	0.16222	0.02000	0.00111
	-0.14333	-0.15444	-0.14889	-0.01778	-0.08657	0.24111	-0.26333	0.32556		-0.04333
	0.19778	-0.07778	-0.08333	-0.27889	-0.26111	-0.18667	-0.03667	0.09111		0.00889
	-0.07889	0.53556	-0.07333	0.30556	0.20556	0.07222	0.13444	-0.35778		-0.04667
	-0.07000	-0.10667	0.04667	0.00556	-0.00333	0.07444	0.05556	-0.04444		0.06111
1800	0.18667	0.05333	-0.19889	-0.21333	-0.02667	0.10667	0.15333	-0.23333	-0.39778	-0.43689
	-0.27889			0.30778	0.46556	0.10778	0.06444	-0.08444		0.08667
		-0.18444	-0.08667	0.24556	0.27556	0.36667		-0.00000-0-	- 1	-0.02889
	-0.04000	-0.42000	-0.39667	-0.43778	-0.45778	-0.18556	0.06111	0.45111	0.65667	0.46889
	0.23000		-0.20000	-0.21000	-0.60778	-0.24667	-0.20556	0.11444	0.18111	-0-11444

TABLE 2B
AUTOCORROLATION KNOX

ķ		1-1900 ft.		1-950 ft.		950-1900 ft.
0		0.0298		0.0313		0.0282
i		0.0168		0.0176		0.0159
2		0.0062		0.0063	Trans.	0.0062
3		0.0036	_	0.0027	_	0.0046
4	_	0.0106	_	0.0108	_	0.0103
5	_	0.0124		0.0143	_	0.0105
6	_	0.0126	_	0.0142	_	0.0111
1		0.0103		0.0122		0.0085
क्ष	_	0.0061	_	0.0074	_	0.0048
9	-	0.0009		0.0022		0.0005
10		0.0028		0.0040		0.0018
11		0.0063		0.0084		0.0045
12		0.0062		0.0081		0.0044
13		0.0065		0.0079		0.0053
14		0.0061		0.0082		0.0040
15		0.0032		0.0059		0.0005
		0.0000		0.9014		0.0014
17		0.0033		0.0027		0.0040
18	_	0.0042		0.0027	_	0.0029_
19		0.0654		0.0091		0.0017
20	_	0.0047	_	0.0091		0.0007
21		0.0035	-			0.0007
22	_	0.9033		0.0076		0.0018
23		0.0010	=	0.0037		0.0020
			•	0.0002		0.0016
24		0.0032		0.0050		0.0012
25		0.0051		0.0092	_	0.0009_
26		0.0047.		0.0103		0.0012
27		0.0034		0.0078	_	0.0017
28		0.0012		0.0037		0.0021
29	_	0.0001		0.0013	Ξ	0.0020
30		0.0023		0.0032		0.0020
31	-	0.0043	-	0.0071	_	0.0321
32		0.0045.		0.0092		0.0021
33	_	0.0033	-	0.0085		0.0721
34	-	0.0011		0.0055		0.0039
3.5		0.0003	-	0.0018		0.0030
		0.0021		0.0024		
37		0.0037		0.0067		0.0014
38		0,0034		0.0067		0.0005
39		0.0036		0.0073		0.0009
40		0.0021		0.0071	H -	0.0029
41		0.0004		0.9039	4.7	0.0135
42	-	0.0018	-	0.0011	-	0.0033
43		0.0026		0.0049		0.0015
44	-	0.0029	-	. 0.0076		0.0006
45	-	0,0027	-	0.0081		0.0018
46	-	0.0019	-	0.0071	•	0.0029
47	·	0.0013		0.0050		0.0725
48	-	0.0007	-	0.0017		0.0007
49		0.0005		0.0017		0.0000
50		0.0021		0.9072	_	0.0019

TABLE 2C

POWER	SPECTRAL	DENSITY
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KNOX

SMOOTHED DATA

λ	-1-1900 ft.	1-750 ft.	750-1900 ft.
-0.00000	0.00042	0.00671	0.00370
0.00500	0.00720	0.01029	0.00320
0.01000	0.01298	0.00370	0.01912
0.01500	0.04368	0.04129	0.04414
0.02000	0.07294	0.05083	0.08900
_0.02500	0.15443	0.10686	0.20543
0.03000	0.30829	0.25217	0.36722
0.03500	0.70377	0.91519	0.47923
0.04000	0.95972	1.23784	0.70308
0.04500	0.62758	0.55708	0.70781
0.05000	0.37234	0.31496	0.42663
_0.05500	0.32086	0.34175	0.29879
0.06000	0.25758	0.27263	0.22290
0.06500	0.20449	0.19727	0.20981
0.07000	0.17224	0.13977	0.20400
_0.07500	0.14001	0.09833	0.18329
0.08000	0.11015	0.09414	0.12397
0.08500	0.C9245	0.09207	0.09176
0.09000	0.08590	0.08895	0.08302
0.09500	0.09092	0.11834	0.06334
0.10000	0.10691	0.12/19	0.08727
0.10500	0.09625	0.06292	0.10923
0.11000	0.05753	0.03763	0.07654
_0.11500	0.04477	0.03153	0.05721
0.12000	0.04182	0.03970	0.0+418
0.12500	0.02707	0.03119	0.02227
0.13000	0.03201	0.04644	0.01699
0.13500	0.04835	0.07903	0.01712
0.14000	0.04166	0.07103	0.01261
_0.14500	0.03002	0.04927	0.01036
0.15000	0.03273	0.05482	0.01078
0.15500	0.02732	0.03717	0.01696
0.16000	0.02926	0.03036	0.02822
0.16500	0.02874	0.03079	0.02643
0.17000	0.03466	0.04867	0.02006
_0.17500	0.04625	0.96493	0.02786
0.18000	0.05175	0.66003	0.04419
0.18500	0.04985	0.04807	0.05188
0.19000	0.04212	.0.04421.	0.04014
0.19500	0.03042	0.03048	0.02969
0.20000	0.02984	9.03509	.0.02445
0.20500	0.03110	0.03580	0.02669
_0.21000	0.02781	_0.02533_	_0.02995_
0.21500	0.03494	0.01424	0.05535
0.22000	0.04699	0.02498	0.06947
0.22500	0.04260	0.03039	0.05490
0.23000	0.03358	0.03105	0.03609
0.23500	0.02569	0.02256	0.02850
0.24000	0.02368	0.02238 0.02678	0.01033
0.24500	0.02117	0.02772	0.01434
			0.01209
0.25000	0.01932	0.02633	0.01209

POWER SPECTRAL DENSITY KNOX λ 0-950 ft 9-475 ft 476-950 ft 0.000 ------------0.005 59.36506 41.53824 18.46144 0.010 4.83219 1.37743 7.11799 0.015 3.34020 3.15744 3.37538 0,020 1.86347 1,29860 2.27377 0.025 1.73363 1.19961 2.30615 0.030 1.81786 1.48694 2.16534 0.0352.50689 3.25999 1.70706 0.040 2.23221 2.87909 1.63529 0.045 1.04178 .92475 1.17496 0.050 . 47123 .39861 . 53994 0.055 . 32756 . 34889 .30503 0.060 . 22286 . 25319 . 19285 0.065 . 15668 . 15115 . 16076 0.070. 12158 .09866 .14400 0.075 .09436 .06627 . 12353 .08240 0.080 .06257 .07321 .06192 .06238 .06213 0.085 .06037 0.090 .06251 .05835 0.095 .06801 .08852 .04238 0.100 .08659 . 10302 .06068 .05585 0.105 .07655 .09695 0.110 .05629 .03681 .07498 0.115 .04823 .03397 .06164 0.120 .04921 .04672 .05199 0.125 .03425 . 03947 .02818 .08186 0.130 .04264 .02263 .10799 0.135.06607 .02448 0.140 .05684 .09692 .01720 0.145 .03987 .06544 .01376 .06938 0.150 .04142 .01364 0.155 .03244 .04413 .02014 0.160 .03116 .03231 .03352 0.165 .02702 .02944 .03155 0.170 .03306 .04643 .01913 .05815 .02495 0.175.04142 0.180.04403 .05108 .03760 0.185.04088 .03942 .04254 0.190.03351 .03224 .03383 0.195.02433 .02438 .02375 0.200 .02417 .02842 .01980 0.205 .02981 .02223 .02590 .02600 0.210.02411 .02199.05066 0.215.03198 . 01303 0.220.04568 .02428 .06754 .05688 0.225 .03147 .04413 .03979 .03423 0.230.03702 0.235.02994.02629 . 03322 0.240 .02884 .03506 . 02233

.03474

.03332

0.245

0.250

.02653

.02445

.01797

. 01530

		18	0.	8	8.94	2.2	2.5	3.6	4.5	.3	3.4	1.7	1.4	2.7	3.6	0.9	.2	.4		4.	9.	8	1.	.2	9.	4.	.2	9.	0		0.	0	0.	6.0	1.8	3.0	3.3	8	4.1	4.5	4.2	3.6
		16	8	8	8.61	2.1	2.4	3.6	4.5	.4	3.5	1.8	1.3	2.6	3.8	1.1	r.	S.	4.	4.	• 6		6.	•	.6	7.	.2	œ	8		0	0		0.7	1.7	3.0	3.2	8	4.2	4.4	4.2	3.8
		4.	. 7	• 2	8.22	2.0	2.2	3.6	4.4	.5	3.6	2.1	1.3	2.5	3.9	1.4	6.	• 6	4.	4.	• 6	1.		8	5	• 9	.2	6.	S.	٦.	0	6.	.2	9.0	1.6	2.7	3.2	8	4.2	4.4	4.2	3.9
	ING	12	•		7.83	1.8	2.0	3.5	4.3	• 6	3.6	2.3	1.3	2.3	3.7	1.7	7.	8	• 6	4.	• 6	1.	3		5	4.	•5	0	•2	•2	0	6.	£.	0.4	1.5	2.5	3.1	7.	4.2	4.4	4.3	0
	TWO FOOT SPACING	10		0	7.50	1.6	2.0	3.4	4.1	•	3.7	2.5	1.2	2.2	3.1	2.0	•	٥.		5	•		•	٥.	L.	9	•	-	9	5	9	5	٠,	0.2	1.4	2.3	3.0		4.1	,	4.4	_
က	TWO FC	œ	7	0	7.16	1.2	2.0	3.3	4.0	5	3.8	2.8	1:1	2.1	3.5	2.3	6.		6.	.4	5	. 7	.7	4.	6.0	• 4		.2	ω.	•2	0.	6.	9.	0.0	1.4	2.2	2.9	1.	4.1	4.3	4.5	4.1
TABLE		9	8	2.	6.80	9.0	2.0	3.2	3.9	. 7	3.9	3.0	1.2	2.0	3.3	2.6	3	3	• 9	4	.5	• 6	0	. 7	0.8	6.	•	7	8		0	• 9	8		1.3	2.2	2.9	9.	4.0	4.2	4.6	4.1
	500 Ft.	4	6	30	6.38	0.1	2.0	3.0	3.8	9.	4.0	3.0	1.3	1.8	3.2	2.9	1.6	4.	0	.2	5	• 6	2.	0	0.7	3	3	•2				6.	6.	3	1.1	2.0	2.8	9.	3.9	4.2	4.5	4.1
	YUMA - 150	8	7	9	5.76	7.	2.1	2.8	3.7	1.	4.1	3.2	1.5	1.7	3.1	3.2	0.1	• 6			5	• 6	4.	• 2	• 2	• 6	. 7	.2	3			6.	0	• 6		1.9	2.8	3	3.9	4.1	3	4.3
	1 X	C	7	4	5.36	2	0	9.	7	Ŋ	2.	3	9	9.	6.	4.	5	0	.2		5	9.	• 6	5	. 1	•	• 6	.2	2	0	. 1	0.	0	8	1.0	1.9	2.9	4.	3.9	4.1	4.5	
													0										,	3								0	,									
			0	20								i	200										,	400				•				600)									

-9.68 -9.93 -11.05

-9.38

-9.19

-9.18

-9.36 -9.79 -10.07

-9.40 -9.81 -10.05

-1.42 -0.84 -1.38 -1.25 -3.06 -8.56 -8.56 -8.56 -9.50 -9.61 -16.98 -10.61 -10.33 -10.46 -10.95 -10.56 -8.67 -8.67 -9.17 -9.18 -9.18 -9.18 -9.16 -1.45 -1.18 -2.91 -4.89 -7.68 -8.12 -9.17 -9.07 -9.88 -9.54 -10.02 -10.68 -10.03 -10.44 -10.46 -10.30 -10.47 -10.95 -10.68 -10.11 -9.05 -8.71 -9.15 -9.46 -1.69 12.00 10.00 -9.93 -10.66 -9.91 -10.23 -10.93 -10.68 -9.18 -8.69 -9.05 -9.42 -9.52 -i0.42 -10.46 -10.06 -2.11 -0.03 -0.23 -1.51 -2.40 -4.33 -7.18 -9.16 -9.16 -9.75 -10.11 -10.30 -10.39 -10.53 -8.69 -8.66 -9.02 -10.91 -10.53 -10.73 -10.27 -10.15 -2.58 -0.24 -1.154 -1.100 -7.00 -7.00 -7.00 -9.14 -9.67 -9.65 -9.40 -10.52 -10.41 -10.22 -10.37 -10.61 -11.05 -10.88 -10.88 -10.76 -10.76 -9.35 -8.71 -8.57 -8.42 -9.04 -9.04 -9.60 -9.57 -9.31 -10.37 -10.06 -10.31 -10.68 -11.20 -10.83 -10.70 -10.73 -10.12 -9.42 -8.68 -2.67 -0.35 0.07 -1.51 -1.91 -4.08 -6.68 -2.89 -0.45 -1.45 -1.07 -1.51 -3.93 -7.06 -9.26 -9.26 -9.26 -9.26 -10.27 -10.29 -10.83 -11.14 -10.77 -10.77 -10.77 -10.71 -10.42 -9.55 -8.67 -3.05 -0.65 -0.20 -1.20 -9.65 -8.82 -8.63 -9.26 -9.26 -9.57 -9.51 -9.61 -10.45 -10.88 -10.77 -10.72 -10.61 -10.47 2 -3.30 -0.92 -0.92 -0.18 -1.21 -1.23 -8.27 -8.27 -9.32 -9.32 -9.32 -9.32 -9.32 -10.27 -10.27 -10.77 -10.34 -10.50 -10.60 -10.95 -9.72 -8.85 -8.85 -3.16 -5.55 -8.90 -7.90 -8.98 -9.93 -9.93 -10.29 -10.29 -10.31 -10.31 -10.78 -9.90 -8.91 -8.66 -1.05 -1.31 -1.46 -1.18 0.18 1200 1000 800

	TWO FOOT SPACING	
TABLE 3A	SMOOTHED YUMA - 1500 Ft.	

	,									
	0	2	4	9	∞	10	12	14	16	18
C	0.0000	0.02556	-0.06778	0.04778	0.17444	0.03556	-0.04556	-0.60889	-0-50444	-0.19333
	-0.02778	0.10889	0.15556	0.43444	0.10667	0	0.03222	0.02222	-0.40889	-6.53222
	-0 36444	(1)		1477	0.09111	0.03333	-0.02222	-0.03778-	19900-0	. 00050-0-
	-0.17111	00090-0-		.0322		.3033	0.36778	0.23444	0.22667	0.20778
	-0.00333	0.02667	•	0060-	9	-15	-0-16556	-0.07111	-0.03111	-0.03778
	-0.05111		0.05889	0.13333	0.11222	0.09222	6.02889	0.05333	-0.01000	-0.04000
	-0.00667	-0.08000	•	2700	٠	-0.02222	0.06556	0.05667	0.09111	0.06333
	-0.06555		-0.01889	0.14889	-0.09000	0.07667	0.13222	0.06445	0.07778	0.00556
	-0-00222	- 1	-0.01778	-0.03667	-0.02000	-0.05111	-0-02333	0.06333	.0.05556	.0.02772
	0.02778		0.04667		0.11778	-0.03222	-0.02000	-0.01333	7	-0.08444
200	-0.05778	0-0-		7	-0-17444	-0.12111	-0.05889	-0-01444	-0.08333	-0-07111
	0.00000	0.0-	-0.02222	0.06333	-0.01111	-0.01444	-0.02333	-0.02444	9440	-0.02556
	0.01333	0.01889	-0.01444	-0.00889	0.08111	0.16333	0-17111	0.29222		7
	0.10667	0.10667	0.05222	0.03667	-0.02667	0.02111	0.03333	0.01778	0.04222	0.11111
	0.05333	0.09	•	0			.3	-0.05444	0.20444	-0-17222-
	0.06111	0-0-		-0.12667	-0.09667	-0.06889	-0.08111	-0.00667	-0.02111	-0.02556
	-0.02222	_	0	-0244	-0-07000			C-03222-	-0.03111-	-00010-0-
	0.03444	_	-0.15222	-0.18000	-0.01778	.0500	0.08222	0	0.03667	-0.01444
	-0.02889				97500 0		0.03333	0.04111	0.05333	-0-01067
	0.01778	-0.0	-0.03222	-0.01778	-0.02556	0.05000	-0.01333	-0.00111	0.07222	0.13889
400	0.25889	0.14	è			.0033	-0.05333	-0.02667	-0.03000	-0-01556
	0.08111	0.1		.01	-0.03222	.0611	2366	0.12556	0.09111	0.09556
	0.03778	0.08			-0.13222	-0.01556	-0.13556	-0.19333	7	-0.16444
	-0.20222	-0.26	-0.45556	-0.38778	-0-32444	-0.17333	0	-0.00333	.1277	0.08556
	0.05333	9	-0-02556	0.06889	0.06556	0.09111	0.05222	0.07778	•	-0-06444-
	0.04556		0.02667	0.01000	0.10111	0.06778	0.04889	0.04222	0.06222	0.06444
	0.06333	0.01	-0-11556	-0-42444	-0-42000	-0-23333	-0.12889		11511	-2055
	0.13000		0.03111	0.00889	0.05111	0.11222	0.03111	.0044	-0.01222	-0.02778
	-0-01444	-0.01000	0.02222	-0.00667	-0.01000-	-0011		•	•	
0	0.11444		0.07778	-0.78556	0.09889	.0777	•	0.03556	•	99460-0
900	00060 0	91-9	0.01444	0.08889	0.01889	-0.01333	000000	0.00333	0.05111	0-11333-
	0.07333		-0.00111	-0.01778	-0.02667	-0.05444	-0.03000	-0.04556	•	0
	-0.07333	-0.00667	0.03111	-0.05000	-0.01778	0.00778	-0.00222	-0-01444	•	0
	-0.00222		0.02889	-0.00111	0.12444	0.09556	-0.01667	-0.06889	-0.30444	-0.18778
	-0-04889	-	77760-0	0.08222	C-02889	0.01222	79910.0	000000-0-	0.03667	
	-0.01778	-0.04444	-0.04000	0.00222	-6.04778	-0.04222	-0.01000	0.01111	-0.01444	0.03444
	0.00778	-0.00556	0.07889	-0.03222	-0.06556	0.01000	-0.06600	-0-04222	-0.08111	0-05333
	0.07889	0.06556	-0.00111	0.00222	-0.01889	0.01222		88	0	C.
	-0-01556	-0.02111	-0.07889	-0-11556	-0-09556	-0*04000	-0.05889	0.08689	-00050-0-	0-03222
	0.05111	-0.11222	0.01000	0.01111	-0.09333	-0.08111	-0.07222	-0.05778	-0.03667	-0.07000

CONT.
У
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3A
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国
TABL
A
H

800	00000	9850	012	.0244	.0222	-		-0.10444	0.00667	0.03000
200	00010-0-	0000	200	1144	0511	0188	.206	2	0.02222	0.0122
	0.02222	0.00	125	1677	1311	1044	.081	03	-0.04222	-0.0633
	0.05778	0.073	3	15.23	1666	1377	084	07	-0.04889	-0.0277
	-0.11333	-0.0177	. 023	7761.	6630	0555	003	80	0.07778	0.0522
	-0.00556	0.0488	023	7711	0000	0000	045	63	-0.16556	-0.6911
	-0.C6778	0.2588	0.0	1197	2770	20.00	70.	7	00090 0	-6.9422
	0.03444	0.1744	.010	<i>دد/۱</i> .	1771	1177	721	, -	0.26222	-0.6533
	-0.01667	0.0100	010	. 0911	017	61133	- 0		-0 08111	0.50
	79899	-0.7388	192	4666	.3566	2133	000	יות ניכי	0.00111	0.70
	0.770	2000	910	1011	.042	.2477	.243	5	0.14111	0.7766
	0.147.0	00000	0 0		0144	0455	.046	5	-0.02111	-0.021
TOOO	0.1773	0.19550	77700 ::	-0 17222	0.04333	7+467-0	(1)	0	-0.02222	(.4.
	0.03778	2	*;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;;	•		7 11 11 11	C	80.01	-0.12000	A 0-
	0.21389	053	·60·0-	7	. 04.	00000	0 0	-0.03	3.13556	90.0-
	-0.11389	136	0.02556	0	9.06333	20.7	5	0000	-0.19222	-0-13vC
	-C 33333	203	0.022	w-4	7	1770	2 (0.00	1 20223	-0.03
	20000	0 1 0	0.170	\sim	1) 111	2	11.0-	66667-6-	
	0.0000	7 7 7		3		: 533	3	0.7.0	0.45000	
	6.15222	1,54	0.10	` 4	1	113	7	-0.01	0.46333	, • O -
	0.19556	404	-0.55-	Ç,	ا ب ر		-	6.23	0.30555	-0.1
	-0.3±111	300	11.0-	-	ζ,	7 1 1	· -	77	0.11773	0
	0.02778	13	-0.201	21,5	• T T • ' -	7,77			0.21803	0
1200	0.05444	40	J. Orb.	1877	-6.4755	3556	٠.		2550	d
	F 7 7 6 C		101	35	0.0000	1,65	Š		000000	
	0.0000	0	7.	5 5 3	- 3.0582	7555	ò	0.0	-0.0.000	0
	777/0.0	7		1 4 4	0.0636	522	Ö	30.0	0.07884	0
	999	90	7	7 1 1	000-01	3200	2	0.0	0.02444	0.11.0
	0.05333	0	3077.01	0.000	2770	11/1	õ	0.0-	-0.97778	0-
	-C.18111	-0.33	VC50.0	0.520	2007	7756	Ö	0.0-	0.03667	74.40
	-0.05111	0.02	1700.0-	0 - Lada		2770	0	0.0	0.00889	00.0
	0.03657	0.03	0.0144	0.0244	0.0211	140	C	0	0.02667	0.1255
	0.00773	0.01	0.0322	-C.CC5-	110.0-		0	0	-0.04778	-0.0244
	-0.00556	-0.03	0.0005	0.0433	0.10.0	7770	•	0	0.03556	-0.02
1400	0.01111	0.05	0.0277	-c.c022	-0.0300	1001	0		-0.10667	0.17
1	00000-0-	-0.10	-0.6266	0.0200	-0.0122	.020	٠.	0 0	-0.67111	-0.62
	-0.08222	C	0.22111	0.28889	0.21889	-0.01222	-0.16222		0-01000	9.008
	-0.03554	0.01	0.0188	-0.017	-0.0133	.020.	•	0	-0.10567	0.1
	-0.0122	-0.00	0.0522	0.1123	0.2400	.401	•	•		

TABLE 3B AUTOCORRELATION YUMA

	1	-1500 ft.		1-750 ft.		750-1500 ft.
k	+			0.0199		0.0245
0		0.0222		0.0104		0.0117
1		0.0110		0.0061		0.0005
2		0.0033		0.0008	_	0.0084
3	-	0.0038		0.0026	-	0.0086
4	_	0.0056		0.0026	_	0.0044
5	-	0.0040	_	0.0042	-	0.0012
6	-	0.0027		0.0023		0.0009
7	_	0.0007				0.0024
8		0.0004	-	0.0016		0.0025
-9		0.0013		0.0001		0.0019
10		0.0014		0.0009	-	0.0013
11	the second second second second second	0.0002	1	0.0017	_	0.0023
	_	0.0003		0.0017		0.0028
12		0.0010		0.0007	-	0.0009
13	191	0.0010	-	0.0011		0.0001
14		0.0014	-	0.0030		0.0011
15	_	0.0004	-	0.0021	σ , with the last time last and with other last last σ	0.0004
16		0.0006	-	0.0019		0.0004
17		0.0008	_	0,0012		0.0026
18		0.0014	-	6.0064	-	0.0038
19		0.0019	-	0.0001		0.0014
20		0.0005		0.0004	-	
21	-	0.0002	-	0.0007		0.0013
22		0.0015		0.0068		0.0039
23		0.0014		0.0015		0.0045
24		0.0001		0.0023		0.0021
25	-	0.0016		0.0024	Sec. 10.00	0.0009
26	-		_	0.6020	-	0.0025
27		0.0022		0.0018		0.0030
28		0.0023	_	0.0011	-	0.0019
29	-	0.0014		0,0002	_	0.0005
30		0.0001		0.0003		0.0007
31		0.0005		0.0003		0,0029
32		0.0013	_	0.0002		0.0032
33		0.0014	ppen	0.0001		0.0013
34		0.0006		0.0001	_	0.0008
35	-	0.0004		0.0003		0.0024
36	-	0.0010		0.0009	-	0.0021
37	_	0.0005				0.0026
38	=	0.0007		0.0012		0.000
		0.0004		0.0012	10174	0.0003
39		0.0005		0.0014	_	0.000
40		0.6000		0.0005		0.000
41		0.0001	-	0.0000	-	
42	_	0.0001		0.0002		0.001
43		0.0002		0.0012		0.001
44		0.0002		0.0012		0.002
45				0.0020		0.000
46	-	0.0005		0.0013		0.000
47	-	0.0004	l	0.0011	_	0.000
48	-	0.0009		0.0004	-	0,000
49	-	0.000)	0.0004		0.000

TABLE 3C

POWER SP	ECTRAL DENSITY		YUMA
λ	1-1500 ft.	1-750 ft.	750-150 ft.
-0.	C.C3108	0.03456	0.02306
0.00500	0.07247	0.10978	0.03372
0.01000	0.13516	0.20851	0.06286
0.01500	0.14861	0.20158	0.09640
0.02000	C.13675	0.14153	0.13135
0.02500	0.14893	0.19571	0.10132
0.03000	0.15961	0.22116	0.09656
	0.17715	0.22175	0.12855
0.04000	0.24567	0.27879	0.21640
0.04500	0.32252	0.33020	0.32251
0.05000	0.27873	0.31742	0.24115
0.05500	0.23495	0.18960	0.27628
0.06000	0.28751	0.10343	0.47389
	0.26737	0.03711	0.45017
0.07000	0.10216	0.03261	0.24130
	0.11908	0.08066	0.15674
0.08000	0.13259	0.09052	0.17507
	0.13532	0.07840	0.19329
0.09900	0.12021	0.05484	U. 18576
	0.03172	0.03633	0.12576
0.10000	0.07369	0.02887	0.11815
	0.07263	0.02206	0.12353
0.11300	0.00451	0.02919	(°C9947
0.11500		0.04391	0.06105
0.12000	0.04383	0.04288	0.04431
0.12500		0.03544	0.03113
0.13000	0.02457	0.03030	0.01783
	0.02363	0.01989	0.02712
	0.02436	0.01866	0.03149
0.14000 0.14500		0.03298	0.02077
		0.03970	0.02139
0.15000	0.03062	0.03570	0.01969
	0.02?93	0.02317	0.02126
0.16000	0.02263	0.02536	0.03232
	0.02875	0.02109	0.02520
0.17000	0.02305	0.02008	0.01443
0.17500	G.01729.	0.02507	
0.18000	0.02132	7.03344	0.61648 0.02931
0.18500	0.03152		
		2.03681	0.03860
0.17500	0.02909	0.02959	0.02791
. 0.20000	-0.03089	0.02759	0.03398
0.20500	0.03366	0.02945	0.03741
_ 0.21000	0.04219	0.04606	0.03887
0.21500	0.04774	0.05372	0.04212
0.22000	0.04089	0.04846	0.03335
0.22500	0.72683	0.03541	0.01791
0.23000	0.02347	0.02730	0.01889
0.23500	0.02575	0.02803	0.02322
0.24000	0.03009	0.025/1	0.03422
0.24500	0.03508	0.02217	-0.0483€
0.25000	0.03683	0.01836	0.04387

-115-TABLE 3D

POWER S	PECTRAL DENSIT	Y	YUMA
λ	0-750 ft	0-375 ft	376-750 ft
0.000	••••		THE PERSON
0.005	418.09392	633.34277	194. 53742
0.010	50.31736	77.62410	23. 40152
0.015	11.36420	15. 41482	7. 37170
0.020	3. 49368	3.61580	
0.025	1. 67188	2. 19704	3.35572
0.030	. 94115	1. 30409	1. 13741
0.035	. 63102		. 56937
0.040	. 57140	. 78989	. 45790
0.045	. 53538	. 64843	. 50332
0.050	, 35276	. 54813	. 53536
0.055	. 23986	. 40172	. 30621
0.060	. 24876	. 19356	. 28205
0.065	, 20486	. 08949	. 41002
0.070	. 11446	. 06674	. 34492
0.075		. 05831	. 17039
0.080	. 08025 . 08813	. 05436	. 10564
0.085		. 06016	. 11636
0.085	. 09131	. 05290	. 13043
	. 08448	. 03854	. 13056
0.095	. 06128	. 02717	. 09422
0.100	. 05968 •	. 02338	. 09570
0.105	. 06446	. 01958	. 10965
0.110	. 06312	. 02856	. 09732
0.115	. 05666	. 04731	. 06578
0.120	. 05158	. 05046	. 05221
0. 125	. 04263	. 04485	. 03939
0.130	. 03272	. 04036	. 02375
0.135	. 03229	. 02717	. 03705
0.140	. 03405	. 02546	. 04296
0.145	. 0356 8	. 04380	. 02758
0.150	. 03875	. 05024	. 02707
0.155	. 02730	. 03095	. 02338
0.160	. 02499	. 02558	. 02345
0.165	. 02946	. 02598	. 03311
0.170	. 02199	. 02012	. 02404
0.175	. 01548	. 01798	. 01292
0.180	. 01814	. 02133	. 01402
0.185	. 02585	. 02742	. 02403
0.190	. 03009	. 02956	. 03100
0.195	. 02327	. 02367	. 02232
0.200	. 02494	. 02234	. 02752
0.205	. 02803	. 02494	. 03116
0.210	. 03663	. 03999	. 03375
0.215	. 04388	. 04917	. 03855
0.220	. 03975	. 04711	
0. 225	. 02779		. 03242
0. 230		. 03668	. 01855
0. 235	. 02588	. 03010	. 02083
	. 03001	. 03267	. 02706
0.240	. 03665	. 03132	. 04169
0.245	. 04396	. 02778	. 06053
0 .250	. 03901	. 02323	. 05552

TABLE 4

BATTLEFIELD DAY - 6201 FT.

ONE FOOT SPACING

	0	1	2	3	4	5	6	7	8	9	10	11
	00003	00009		00022	00027	-		00048	00053	00059	00064	
	00074					_	_	00122				
			00180					00214				
	00252		00255					00294				
	00331		00344					00413				
		00478		00498			00534		00549			
			00605					00682				
	00736	00737	00728	00718				00687				
	00666	00660	00657	00648	00642	00638		00636				
	00617	00610	00606	20800				00580				
			00539					00512				
	00520	00520	00518	00512	00512	00510	00507	00506	00505	00540	00513	00527
	00533	00545	00560	00573	00588	00608	00634	00653	00666	00669	00660	00651
	00640	00627	00613	00595	00572	00557	00551	00548	00540	00535	00535	00533
	00526	00517	00502	00484	00468	00451	00432	00409	00390	00376	00367	00347
	00339	00332	00327	00324	00324	00324	00327	00324	00320	00317	00314	00308
	00307	00300	00295	00292	00287	00283	00279	00275	00273	00268	00266	00268
	00269	00275	00281	00285	00293	00296	00296	00296	00294	00299	00303	00302
000	00294	00290	00288	00288	00288	00286	00286	00290	00291	00295	00297	00297
298	00298	00300	00302	00305	00303	00300	00296	00297	00298	00297	00296	00297
	00301	00302	00302	00302	00302	00304	00307	00309	00309	00311	00315	00316
	00320	00323	00324	00324	00327	00330	00332	00335	00335	00334	00331	00330
	00328	00326	00326	00325	00335	00335	00335	00334	00336	00337	00339	00341
		00334	00332	00336	00340	00342	00343	00346	00340	00336	00340	00341
	00341	00340					00338	00338	00344	00346	00344	00343
			00353					00355				
		00398					00421		00428	00426	00422	00416
	00414			-		-	00385		00378	00374	00370	00368
336		00365					00376			00391		00403
		00406	_			00395		00381	00375			00363
		00363		00364		00374	00381	00388			00391	00389
	00392		00388					00377				00401
		00432						00479				00460
	00451							00445		00440		
	00431	00430		00431				00448			00477	
	00541	00505		00511				00527 00588	00528	00530 00597		00603
	ma							00646				00688
456	00694						00030			00709		
		00713		00721				00726				00743
		00752			00759			00756				
		00791						00849				
	2 - 115						_	00812	_	00816		
	_							00804				
								00851				
								00898				
								00885				
564								00828				
								00860				
	00873	00879	00881	00883	00885	00887	00887	00883	00872	00865	00860	00857
	00852	00846	00843	00842	00848	00851	00857	00869	00879	00887	00894	00897
	00898	00900	00898	00892	00887	00884	00878	00876	00878	00880	00883	00885
												00898
												00914
	00908	00908	00902	00900	00895	00910	00921	00922	00921	00926	00930	00935
	00938	00944	00948	00952	00952	00954	00956	00959	00960	00964	00970	00976
	00982	00992	01001	01008	01017	01025	01030	01031	01028	01024	01021	01017

	0	1	2	3	4	. 5	6	7	8	9	10	11
004		-	_					•	_	-		
684	01107	011011	01012	01020	01030	01040	01045	01058	01069	01079	01090	01105
								01089				
								01185				
								01202				
								01277				
	-							01320	-			
								01377				
								01446				
								01476				
804								01588				
								01591				
								01693				
	01714	01712	01706	01701	01694	01687	01682	01683	01684	01685	01685	01686
	01688	01690	01695	01700	01705	01715	01723	01730	01737	01740	01742	01746
	01751	01754	01755	01756	01757	01760	01760	01759	01761	01762	01765	01769
								01801				
	01821	01824	01832	01847	01868	01893	01911	01925	01939	01955	01963	01962
	01956	01946	01939	01931	01922	01911	01909	01925	01931	01936	01944	01947
912	01952	01958	01964	01971	01998	02006	02027	02036	02034	02032	02030	02026
	02018	02017	02018	02013	02014	02016	02017	02017	02016	02013	02008	02010
	02013	02019	02028	02035	02040	02044	02052	02063	02073	02083	02095	02103
	02110	02117	02126	02134	02141	02144	02146	02148	02149	02155	02160	02165
	02173	02186	02196	02207	02218	02227	02237	02241	02245	02249	02257	02265
								02340			-	
				_		_		02468	-		-	
								02609		-		
*		02631						02569				
								02602				
1032		02624						02722				
								02743	_			
		02821						02791				
								02735		_		
								02853 02856				
								02950				
								03043				
								03075				
1140	03111							03185				
1110								03251				
								03310				
								03416				
								03467				
						_	_	03452				
	03447	03451	03463	03475	03485	03493	03498	03500	03500	03496	03491	03487
								03495				
	03546	03551	03555	03554	03549	03540	03535	03535	03533	03532	03532	03530
	03529	03527	03529	03534	03538	03540	03543	03553	03561	03568	03581	03599
1260	03601	03606	03610	03609	03610	03611	03613	03615	03618	03626	03634	03633
	03639	03647	03652	03656	03664	03670	03674	03675	03675	03671	03667	03664
	03661	03662	03661	03659	03664	03664	03665	03670	03673	03679	03687	03693
								03715				
	03723	03730	03735	03738	03740	03746	03746	03748	03749	03749	03753	03755
	03762	03765	03769	03777	03787	03792	03794	03794	03798	03800	03802	03805
								03872				
								03934				
	03973	03984	03995	04004	04016	04029	04036	04039	04039	04043	04052	04054

-118-TABLE 4 CONT.

	0	1	2	3	4	5	6	7	8	9	10	11
1368	04059	04067	04066	04093	04112	04125	04134	04141	04145	04156	04167	04176
	04174	04174					04204			04226		
	04258	04271	04281				04301					
	04352		04371	04381		04394		04405		04425		
			04469				04512			04541	04551	04562
		.04586	04594		04609		04615			04635		
		04669					04716				-	-
							04852			04886		
4	04910		04920						04957			
	04993		05010				05037		05046	05052	_	
1488	05083		05110			05113		05107		05095		
			05101					05154		05151		
							05148		_	05173		
	Mar -	05181					05154					
		05101					05097					
	05060	05056					05063			05075		
	05095	05091	05089				05081					
	05055	05050	05041			05014		04992			04962	
	04942	04938	04938	04937	04944	04952	04967	04963	04959	04955	04953	04950
	04953	04952	04951	04954	04950	04946	04945	04942	04940	04935	04938	04940
1608	04926	04923	04918	04914	04919	04912	04909	04907	04916	04918	04914	04914
	04910	04906	04901	04905	04903	04908	04902	04897	04893	04888	04882	04876
	04871	04866	04866	04875	04883	04898	04908	04909	04911	04906	04901	04910
	04918	04914	04910	04906	04904	04901	04894	04893	04887	04883	04879	04878
			04872			04901	04916	04914	04914	04916	04914	04911
			04927	04933	04934	04935	04939	04949	04986	04987	04986	04980
		04997			04983			04970		04962		
			04960		04958				-	04969		04985
1716		04987			04994		04999			04993		
	04987						05009					
		05020		05010			05003	05007			05019	
		05044	05052				05047		05044		05057	
	05051						05061		05058		05051	
	05043		05031				05037				05028	
		05032			05049			05097		05110		
			05134				05129			05119		
	05125		05115				05098			05105		
	05253			05265				05284	05291		05303	
1836							05394			05428		
		05482					05532			05583		
			05648					05690	05701			05740
			05771				05826			05858	05865	
	-						05870			05854		
							05932					
							06007					
							06114					
							06189					
1944							06253					
	_						06354					
	06399	06404	06408	06414	06414	06416	06422	06427	06435	06442	06456	06482
	06501	06511	06515	06520	06535	06545	06550	06552	06553	06553	06552	06550
	06548	06553	06556	06570	06572	06577	06588	06603	06614	05620	06627	06634
	06638	06653	06663	06665	06682	06692	06694	06696	06701	06706	06708	06709
	06709	06710	06710	06710	06712	06715	06718	06770	06724	06731	06748	06757
	06765	06771	06784	06790	06792	05794	06/93	06790	06 785	06779	06775	06778
	06778	06781	06782	06783	06783	06786	06786	06797	06808	06806	06804	06802
	06795	06792	06786	06/81	06774	06779	06779	06773	06777	06774	06772	06769

	0	1	2	3	4	5	6	7	8	9	10	11
2064	06768	06766	06759	06754	06748	06746	06745	06744	06744	06743	06742	06742
2001	06739	06737	06734							06702		
	06694	06690	06682			06666				06644		
	06626	06622	06620			06630			06648	06650	06654	06665
	06683	06699	06712		06736		06761		06778	06785	06795	06807
	06818	06831		06856				06907	_	06940		
	06978	06989	06996						_			
	07205	07225	07237	07265	07286	07293	07306	07319	07328	07335	07351	07369
	07381	07393	07407	07415	07431	07450	07465	07471	07477	07483	07487	07487
	07486		07507	07516	07523	07529	07535	07558	07569	07585	07604	07613
2184	07618			07637				_	07667	07674	07676	
	07685		_			07727			07741	07742	07749	07756
	07758		07762						_	07823		
	07826			07840						07883		
	07913	07914				07950			_	07957		-
	07967			07977						07995		
	08002	08009	08025	08031	08032	08029	08027	08021	08014	80080	08002	07999
	08005		07998				08021	08034	08058	08060	08067	08075
	08081	_	08094				08134			08167		
	08204		08233					08314	08332	08349	08392	08421
2304			08503				08612	_		08664		
			08770							08910		
			09038				09154	09184	_	09241		09294
	09322		09368				09481	09498		09534		09573
			09638					09734		09776		
	09836	09856	09881	09910	09936	09952	09971	09994		10038		
	10099		10130		10167		10195	10210		10230		
	10259		10279				10344			10412	-	10466
		10508		10569			10564	10560		10545		10551
		10525		10497			10454	10440		10410	-	
2436			10358				10334			10310		
1 100,	10326	10320		10304				10258	10244		10216	
			10159				10113	10099		10058		
			09980		09949		09927			09895		
			09828					09819		09792		
			09745							09675	_	
		09661		09657				09624		09607		
		09583		09570					_		_	
		09488	09495	09502			09474			09440		
		09288		09296				09261				_
2556							09271			09240		
2000		09195 09135		09179			09156	09148		09141	09138	
						09140				09165		
	09165					09198				09218		
			09228							09242		
		09255		09257						09286		
	09301	09303		09310				_		09333		
	09341		-	09347					_	09340	_	
											-	
		09365		09368				_	-			
	09488	09499	09503	09505	09506	09515	0.8219	09522	09525	09527	09529	09531

-12 0-

	0	1	2	3	4	5	6	7	8	9	10	11
0676	09531	09532										
2676	09539						09536					
	09585		09537				09553					
							09627					
- 0	09631						09620					
	09655			09635					-		09646	
			09653 09637		09644		09646	-				
			09661				09670		09651		-	
	09688						09708					
	09730			09740							09759	
2796	09761		09769		09790						09828	
2130	-	09842			09855						09912	
	09924		09934		09956			09981			10003	
		10020	10024		10024						10048	
			10052		10069		10084			_	10107	
		10114	10121				10159					10217
		10219					10203					
				10230							10214	
	10226						10284				10292	
		10291	10289		10313		10316				10327	
2916	10345		10345		10344			10324	10314		10299	
	10320	10337	10344	10352	10361	10370	10378	10384	10390	10397	10396	10396
	10402	10403	10397	10395	10390	10389	10389	10385	10388	10387	10388	10388
	10385	10384	10384	10383	10381	10380	10380	10377	10375	10371	10369	10367
	10366	10366	10365	10363	10364	10362	10360	10359	10355	10353	10352	10351
	10348	10351	10349	10346	10342	10337	10333	10321	10316	10312	10308	10307
	10306	10308	10313	10320	10327	10537	10344	10350			10358	10358
	10358	-		10349			10349	10351		10363	10368	
	10382			10411			10451	10460	10467			10494
2026	10512			10565				10591		10573	10561	10554
3036		10539			10523		10521	10522	10522	10523	10526	
	10537		10553		10570			10567			10548	
	10540	10540	10541	_	10544		10549		10552	10552	10546	10540
	10535		10520		10508		10501	10495	10493		10503	
			10503		10502		10495	10491	10485		10469	
		10454	10453		10447		10451	10455	10459			10470
	10473		10473	10471		10465	10465	10464	10464	10464	10478	10483
	10485	10491	10495		10495		10498	10495	10492		10488	10485
		10487		10490		10507					10506	
	10484	10472	10460	10439		10417	10411	10413	10405		10245	
3156	10365		10342	10329	10314	10303		10275	10230		10245	
		10229		10230 10194	10231	10233					10154	
			10167		10172		10173				10203	
			10246	10254		10267	10271	10788		-	10287	
			10317			10335	10235			102327		10316
		10320	10317		10330	10350	10360	10331	10384	10395	10403	10412
	10418	10426	10429	10432	10439	10446	10456	10468	10477		10494	10508
	10515	10524	10538	10547		10574	10588	10595			10610	
	-	_		10625								10687
	10014	10020	1002. 3	10060	10000	10044	10001	*******	1 (11/0)		. 0.000	.000

	0	1	2	3	4	5	6	7	8	9	10	11	
3276	10684	10680	10679	10680	10685	10692	10702	10710	10717	10726	10736	10745	
	10751	10756	10762	10770	10776	10780	10783	10786	10790	10797	10803	10807	
	10812	10817	10819	10821	10824	10828	10832	10838	10847	10855	10864	10870	
	10881	10887	10397	10904	10912	10919	10927	10937	10951	10959	10970	10978	
	10988	10996	10999	10999	10993	10982	10967	10955	10946	10938	10927	10916	
	10919	10917	10933	10934	10935	10939	10938	10937	10938	10942	10950	10950	
	10945	10944	10947	10947	10948	10947	10945	10948	10948	10950	10952	10954	
	10957	10958	10958	10959	10960	10964	10967	10969	10977	10981	10981	10978	
	10977	10976	10974	10968	10960	10949	10944	10937	10932	10926	10917	10911	
	10908	10901	10885	10878	10882	10894	10900	10902	10904	10900	10896	10891	
	10887	10886	10886	10884	10880	10874	10887	10884	10881	10879	10878	10876	
3402	10870	10864	10863	10864	10964	10865	10868	10875	10884	10891	10997	10910	
	10927	10938	10949	10959	10968	10973	10977	10978	10977	10974	10966	10960	
	10953	10946	10940	10937	10934	10932	10928	10943	10944	10946	10948	10951	
	10954	10973	10987	11000	11002	11006	11008	11012	11016	11021	11027	11032	
	11032	11035	11037	11040	11043	11053	11067	11069	11071	11072	11084	11085	
	11085	11083	11084	11084	11083	11077	11074	11069	11064	11065	11063	11062	
	11060	11061	11064	11065	11069	11074	11086	11096	11104	11107	11109	11109	
	11106	11104	11099	11095	11092	11091	11091	11092	11098	11104	11122	11142	
	11149	11154	11157	11160	11160	11154	11148	11145	11145	11144	11144	11146	
	11149	11149	11150	11154	11152	11154	11151	11147	11147	11148	11150	11153	
3528	11148	11158	11157	11158	11158	11157	11153	11148	11146	11144	11141	11138	
	11142	11148	11153	11161	11164	11171	11176	11182	11191	11195	11193	11191	
	11188	11187				11179	11175	11177	11179	11182	11184	11186	
		11188	11188	11183	11170	11176	11176	11175	11172	11170	11169	11168	
		11167		11179			11179				11157		
	-	11142		11134				11123		11117		11107"	
		11104				11095		11085		11077		11065	
	11061	11057		11049			1026	11017		10993 10898		10973 10890	
	10972	10962	10953	10944	10865			10915		10820		10799	
3648	10791	10784	10779	-	10763		10756			10752		10738	
3040	10738	10738	107731	10770	10726		10719			10703		10697	
	10694	10693	10692		10687		10686		10683	10680	10678	10675	
	10674	10676	10676		10637			10672		10668		10664	
	10660	10657	10656		10655				10654		10653	10655	
	10660	10659			10675		10674		10683		10694	10702	
	10711		10723		10737		10738			_		10745	
	10747	10751	10757		10771	10778	10785	10793		10804		10821	
	10829	10839				10859	10865		-	10881	10890	10907	
	10912	10920			10942		10961	10969		10091	11001	11011	
3768	11020		11037			11062		11064		11060		11056	
0.00	11060	11065			11084	11096		11128		11152		11164	
	11165	11168				11185	11197	11207	11220	11229	11234	11239	
	11245	11244	11240	11237	11234	11230	11228	11229	11228	11229	11232	11243	
	11252	11258	11262	11266	11267	11267	11268	11272	11277	11283	11290	11296	
	11299	11306				11337		11350		11360		11372	
	11379	11388				11420		11425	11428	11433	11436	11439	
	11440	11442		11449	11454	11449	11464	11472	11481	11489	11495	11504	
	11513	11521	11526	11532	11537	11543	11551	11557	11563	11569	11578	11593	
	11607	11620	11631	11641	11649	11658	11662	11667	11672	11677	11678	11679	
	11683	11690	11697	11708	11719	11736	11747	11759	11774	11791	11805	11820	

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3900	11833	11846	11857	11865	11871	11877	11883	11895	11904	11912	11918	11929
								12024				
								12149				
	12186	12195	12205	12219	12232	12248	12260	12277	12288	12299	12311	12322
	12333	12343	12353	12364	12376	12386	12305	12404	12411	12420	12428	12435
	12442	12448	12456	12464	12471	12478	12486	12494	12503	12513	12517	12522
	12525	12534	12542	12549	12556	12564	12571	12578	12588	12597	12600	12624
	12648	12667	12679	12686	12685	12676	12661	12638	12612	1.2585	12560	12558
								12540				
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4020								12588				
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4152								12283				
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								12523				
4260								12560				
1200								12592				
								12615				
								12661				
								12600				
								12581				
	12580	12581	12580	12580	12579	12578	12577	12576	12574	12574	12573	12580
	12584	12583	12602	12599	12598	12597	12593	12592	12589	12587	12586	12584
	12579	12578	12577	12575	12573	12572	12570	12569	12569	12570	12568	1256/
	12566	12565	12563	12563	12563	12563	12563	12561	12561	1 2559	12559	12558
	12557	12554	12551	12548	12544	12541	12537	12532	12528	12532	12529	12524
	12521	12517	12514	12511	12509	12509	12508	12501	12498	12494	12492	12493
4392												12503
	12501	12499	12498	12496	12491	12487	12482	12475	12482	12483	12478	12471
	12460	12456	12452	12449	12447	12442	12436	12430	12427	12423	12418	12412
	12404	12395	12386	12381	12375	12369	12360	12350	12343	12338	12333	12323
	12318	12312	12309	12309	12307	12304	12302	12303	12309	12310	12313	12317
	12325	12331	12336	12342	12348	12353	12360	12367	12373	12382	12390	12400
	12410	12417	12425	12437	12435	12442	12452	12554	12462	12480	12993	12503
	12516	12532	12540	12552	12564	12576	12590	12602	12614	12630	12644	12656
	12669	12687	12704	12720	12734	12753	12760	12782	12799	12811	12818	12824
	12832	12841	12847	12852	12860	12867	12879	12894	12911	12925	12943	12965

TABLE 4 CONT.

	0	1	2	3	4	5	6	7	8	9	10	11
4512	12983	13010	13030	13051	13071	13087	13103	13126	13143	13160	13174	13197
				13247								
				13403								
	13533	13546	13551	13563	13585	13604	13623	13639	13652	14673	13690	13706
				13760								
	13874	13898	13913	13926	13943	13956	13963	13978	13991	13991	13998	14007
	14016	14023	14034	14035	14031	14049	14057	14071	14087	14095	14111	14122
	14135	14148	14167	14179	14192	14200	14210	14220	14229	14239	14244	14247
				14284								
	14392	14406	14416	14425	14433	14442	14453	14461	14468	14475	14481	14486
	14494	14501	14508	14515	14522	14533	14543	14557	14570	14581	14594	14608
4644	14621	14631	14644	14657	14663	14674	14689	14701	14712	14726	14736	14741
	14759	4766	14781	14790	14799	14810	14819	14833	14841	14852	14860	14866
	14873	14883	14897	14910	14925	14931	14937	14938	14947	14955	14960	14966
	14974	14982	14988	14996	15008	15018	15030	15041	15047	15052	15069	15071
	15082	15095	15105	15112	15116	15124	15135	15142	15149	15166	15174	15181
	15195	15220	15231	15248	15264	15284	15303	15319	15333	15351	15362	15373
	15385	15403	15412	15421	15432	15449	15461	15466	15482	15495	15503	15509
	15523	15532	15543	15551	15558	15572	15574	15582	15608	15618	15634	15645
	15659	15672	15681	15683	15692	15709	15733	15740	15755	15765	15776	15785
	15802	15823	15844	15860	15880	15900	15910	15923	15933	15943	15946	15946
	15949	15952	15956	15963	15971	15982	15989	15996	16007	16017	16020	16023
4776	16027	16030	16034	16038	16045	16051	16061	16071	16086	16099	16106	16117
	16127	16134	16143	16154	16163	16171	16179	16188	16195	16202	16214	16220
	16223	16228	16235	16238	16237	16242	16248	16252	16253	16254	16254	16252
	16251	16252	16253	16256	16256	16259	16266	16278	16297	16314	16328	16344
	16367	16395	16413	16421	16424	16424	16423	16422	16420	16416	16423	16435
	16448	16459	16472	16479	16488	16503	16513	16524	16531	16542	16551	16553
	16562	16568	16584	16697	16606	16616	16623	16625	16627	16628	16632	16638
	16647	16661	16670	16690	16687	16691	16697	16704	16710	16711	16710	16713
	16719	16727	16734	16740	16749	16757	16763	16770	16776	16783	16787	16793
	16797	16800	16809	16810	16828	16833	16844	16854	16866	16874	16883	16891
4896	16896	16905	16911	16917	16924	16932	16939	16945	16954	16965	16969	16979
	16989	16999	17008	17016	17026	17034	17043	17057	17067	17077	17085	17095
	17104	17112	17119	17126	17133	17140	17149	17157	17165	17174	17182	17197
	17209	17218	17234	17242	17250	17254	17260	17265	17269	17272	17277	17283
	17303	17311	17324	17335	17343	17355	17361	17368	17376	17383	17393	17403
	17407	17410	17417	17422	17428	17434	17440	17447	17450	17456	17463	17468
	17473	17472	17454	17488	17491	17491	17499	17513	17521	17533	17536	17541
	17549	17560	17575	17594	17602	17612	17625	17624	17628	17638	17642	17645
	17652	17657	17665	17666	17670	17685	17692	17700	17706	17717	17724	17731
5016	17738	17749	17756	17762	17775	17786	17794	17799	17808	17814	17821	17827
5016	17839	17844	17850	17855								
			17921			17957						
	Act of the last of			18040								agent with the section
				18166								
		10777	10763	10770	18383	18393	18407	18420	18420	18430	10452	18466
	18478	18491	18509	18522	18538	18556	18571	18583	18599	18617	18633	18645

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.5100		18891						18995				
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	-		18169					18165				10 mg
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3040								17815				
								17773		the second section is		
								17716				
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	****							17644			THE TAX .	Sec. 2. 5. 1
								17660				
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	de anteres de la							17739			THE R. P. LEWIS CO., LANSING, MICH.	
								17750				
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												17972
5724	17980	17992	17008	18005	18010	18011	1800F	18001	17988	17970	17956	17947
1795	17933	17916	17011	17916	17930	17940	17943	17947	17966	17983	17992	17997
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TABLE 4 CONT.

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	17992	17981	17972	17961	17955	17957	17950	17944	17944	17948	17951	17952
	17953	17956	17959	17960	17961	17962	17966	17971	17973	17978	17989	17993
	17989	17987	17986	17983	17980	17981	17980	17980	17980	17985	17993	18006
	18018	18027	18036	18043	18051	18064	18072	18080	18085	18083	18083	18081
	18085	18081	18078	18076	18077	18082	18088	18098	18113	18128	18145	18154
	18169	18173	18173	18172	18171	18170	18170	18170	18170	18171	18173	18175
	18179	18186	18191	18197	18204	18212	18218	18226	18239	18259	18266	18274
	18279	18281	18282	18283	18284	18284	18283	18296	18293	18291	18291	18289
	18288	18286	18285	18285	18284	18284	18284	18280	18279	18280	18283	18284
5856	18285	18289	18292	18293	18295	18299	18306	18312	18318	18325	18333	18337
	1.8341	18344	18346	18345	18344	18344	18343	18340	18338	18339	18339	18340
	18345	18348	18349	18351	18353	18352	18353	18354	18355	18359	18361	18358
	18353	18351			18351		18351	18352				
	18384	18401	18403	18406	18406	18404	18404	18407	18412	18419	18429	18433
			18450					18445				
	18430				18428			18429				
	18475		18501				-	18522	-			
	18512		-	_				18526				
								18519				
5988								18526			and the second	and the
0000	18559	18570						18646				
	18647	18638	18631					18620				
	18657		18660					18677				
	18708	18710	18709				18684	18675			18670	18675
	18689	18707			18768	18791		18816			18819	18813
	18799	18788			18751	18741	18738	18738			18764	
	18784	18791	18796		18809			18821	18831		18851	18859
		18873			18890			18894				18921
		18934		18941		18924					18932	
	18925	18925	18925					18940			pro	
6120	18961	18965		18975				18983			18966	
	18955			18954		18943		18942			and the same of the	18950
	18964	18983			19010			19020				
	18990							19006				
	19044							19083				
	19124				19140			19148				An
		19201							19241	19245	19245	19250
	19257	19463	19264	19269	19278	19285	19293	19300	19308			

	ONE FOO
TABLE 4A	SMOOTHED BATTLEFIELD DAY -6201 Feet

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8 0.002554 0.0 .05556 000013333 000133333 000133333 000133333 0001333 00013333 0001333 0001333 0001333 00013 000133 000133 000133 000133 000133 000133 000133 000133 00013 000133 000133 000133 00013 000133 000133 00013 000133 000133 00013 0.03567 0.02111 0.05222 0.01889 0.02657 0.02650 0.02657 0.02657 0.03889 0.02444 0.02778 0.02689 0.02689 0.02689 0.05222 0.05222 0.05222 0.05222 0.01778 0.01778 0.01222 0.0222 0.0333 0.0323 0.0323 0.0323 0.0323 0.0323 0.0323 0.0323 0.0323 0.0323 02657 001110 0211 -0.02555 -0.02222 0.01444 -0.05889 0.03333 -0.000111 -0.03333 -0.03222 -0.03333 -0.0022444 -0.0024444 -0.0024444 -0.0024444 -0.003333 -0.00222 -0.003333 -0.00222 -0.003333 -0.00222 -0.003333 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 -0.00222 0.03889 -0.05778 0.05000 -0.08444 0.00222 -0.001111 -0.001718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01718 -0.01756 -0 -0.03000 0.05000 -0.03222 0.04222 -0.00444 -0.001111 -0.00222 -0.00222 -0.00222 -0.002111 0.002111 0.00222 -0.003556 -0.003778 -0.003778 -0.003778 -0.003778 -0.003778 -0.003778 -0.003778 -0.00222 -0.003778 -0.01333 0.02889 -0.03111 0.00000 -3.00778 -3.03000 -3.01667 -3.03333 -0.06000 -0.06556 -0.04889 -0.04778 -0.00333 -0.03556 -0.12000 -0.10556 -0.10556 -0.01222 -0.01222 -0.01222 -0.01333 -0.00111 -0.00333 -0.00333 -0.00333 -0.00333 -0.00333 -0.00333 -0.00333 -0.00333 -0.00333 -3.01444 3.03000 -3.02222 5.62000 .00778 0

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1	0255	.003	144	.030	.0200	000	.0156	.0255	.3133	.322
	0.09111	990-0	.0077	0077	.030	.0455	0.32303	.33	50	.210
	.0933	165	90	.0100	.0366	.0555	.0588	.0722	.0377	·116
	.0077	.02	1116.	.0344	5	4	.0433	5	.3133	0-033
	.0066	.020	1150*	.0333	. 3266	.0288	. 3100	.0544	.0722	.057
1	•	0.00	.3333	.0366	.3244	.013	11110.	.0383	533	276
	•	0.10	.0844	.0733	009	.0322	.3622	.0355	.0455	.043
1		0+0-0-	.0300	.0211	0	0.04444	.0400	244	. 2677	0.3
		-0.0411	1160.	.0811	.1211	.0588	* 500.	.0422	.035	1+0-
		0.0233	544	-0466	.0633	.0411	.0144	65	-0.02111	-336
	•	0.078	.1088	.1244	311		.1677	.1122	-2052	-0-388B
1	C	-0-1633	.1155	.0844	.1166	. 1055	.3788	.036	-2.0044	.091
	•	0.1433	11	.1211	-0933	-0366	. 1133	.331	-3.1044	.132
1	Ċ	-0.0766	.0733	-0322	344	.0411	. 1211	.332	-2.2433	900
	•	-0.0088	.1116.	-0044	.0088	.0122	.3177	87746.0-	-2.3055	.027
		0.025	.0388	.0677	.3888	.0755	. 2000	56	-0.0277	.337
•	o	-0.05	.0766	.0966	.0611	00	.0633	.033	3.3344	.256
		0.0655	0.05556	.0700	CO9	.0200	.3098	18	-0.0600	.355
1	o	0.0-	.0100	.0133	.014	.0055	.0055	.018	2.057	0.0588
		-0.0177	.0533	.0466	.0500	.0366	.0055	11	0.0055	.338
		0.0100	055	.0300	35	.0611	1746.	.337	-3.3233	.03
•	ď	-0.0838	.1466	.0166	.3688	455	.3130	14	-3,3155	.001
•	o	0.0188	.0255	.0333	.0055	.0033	.0122	.013	-0.035	.041
1	o	-0.0311		.0077	0.00889	0.01111	2433	0.07657	3.0946	0-3866
	o.	0.0177	.0022	.0266	.0655	.0988	1122	.373	-3.3311	• 224
1	Ċ	-0.0222	.0088	-0100	1164	.1466	.1330	-073	3.3333	0
1	C C	0	0088	0 -	190	0544	-0.04222	-0.02444	2	1900
•	5 0	1055	0000	0400	0071-	0321	0222	200	7656	0,00
		-0.0		0355	38	355	0.31778	215	940	23
1		-0.0244	.0377	.0655	.0733	.0788	0010	-0322	. 3238	.020
		-0.0011	-0044	0277	.0355	.0388	.0430	0288	033	328
		0.0	940	.0011	.0244	17770.	.3600	.3777	.3333	.31
		900.0	.0144	0040	544	.0300	.3077	.0133	.0333	+CC-
1		-0.02		-0.04333	.033	166	0.02000	11	4	9990-0
		0.0	-0322	.0077	*0022	0.	.3522	.0455	.0522	.355
1	-	-0.154	119	.024	110.	.354	0	457	49	2.3211
		0.06445	0.03556	-0.00111	-0.06667	-0.07657	855	-0.03444	3.04445	100

	တ			1		0.0177		-3-1033	3.1577							1			1		'	1		1		'	•		•			0-000
	œ	-3-39444	0.05445	-0-0500	3.02222	00000-0	-5.57111	-0-03666	0.19000	3.30222	3.16111	-3.32222				-0.00889	-1,13333	-0.01000	-3.37222	0.10557	-3.39557	3.31778	0.08778	-2.24222	3.16555	1			-3.32557	-2.30111	0.00555	-0.00889
	2			1	0.00000	15900-0	-0.01111	-0.03000	0.15111	-0.03444	0.03111	-0.04333	-0.04333	0.32555	-0.03444	3.32555	-0.33333	0.01222	-0.02111	0.13834	-0-03333	0.01657	0.000.0	-0.02000	0.13557	-0.09555	0.0333	0.09000	0.01222	-0.01839	-0.01839	0.00657
	9	-0.07333	0-38445	-0.37567	-0.04222	0.01556	0.01333	-0.05555	3.38773	-0.18555	0.04000	-0.30779	-0.37222	0.33567	-0		-0-	0.31555	1			-0.30222	0			-0.12222	-0.00111	0.38779	0.30222	0-20002	-0.06303	0.02445
A CONT.	ഗ	-0.07333	0.11000	-0.05667	-0.03555	0.02778	0.02667	-0.03111	0.05889	-0.15778	-0.13333	-0.03444	-0.07657	0.03778	-0.01111	0.04111	0.03222	-0.02222	0.01889	0.23222	-0.05667	0.01222	-0.09000	0.03333	0.04889	-0.13333	-0.04667	0.07889	-0.02222	-0.02667	-0.03555	0.04222
TABLE 4A CONT	4	-0.06889	0.09333	-0-01889	-0.03555	0.02333	-0.00444	0.00778	0.07333	-0.08000	-0.10111	4444C-0-	19910.0-	19500.0-	0.01222	0.33556	0-31445	-0.02667	0.02556	0.16111	-0.35222	0.02445	-0.09111	0.04111	111110-0-	-0.03111	-0.08222	0.31667	-0.02333	-0.02222	-0-01667	0.34667
	က	-0-02000	0.04333	0.01889	-0.04333	0.00333	-0.02333	0.03667	0.04778	-0.02444	-0.06000	0.01111	-0.01333	-0.02333	0.04000	0.01445	0.01000	-0.02444	0.01667	0.07556	-0.04222	0.05222	-0.07778	0.09222	-0.05222	0.06445	-0.05222	-0.03444	-0.00778	0.00000	0.00111	0.03333
	2	-0-00889	A7700 0	00000	0.010	00000	-0.01567	0.05556	-0.05333	0.01333	-0.03667	0.03556	0.32222	-0.05000	0.06000	-0.30555	-0.02000	-0.31667	0.01445	-0.01333	-0.050.0-	0.07000	-0.06000	0.12222	-0.10333	0.09333	-0.06444	-0.06111	0.02556	-0.05222	0.02667	00040-0
	П.	0 07880	0.0000	97770	0.01446	0.02645	00000	0-03333	-0.12444	0.03222	-0.03555	0.06889	0.03222	-0.04555	0.04111	-0.01444	-0.03333	-0.00111	0.02333	-0.09222	-0.01111	-0.01778	-0.09333	0.11333	-0.12778	0.08667	-0.04667	-0.08555	0.05445	-0.05555	0.05000	0.00667
	0		0.12333	9701-0-	0.08000	*******	44400	000000	-2-14222	0.10111	0.03333	00011.0	0.02333	-3.00778	0.02000	-2.05778	-0.03778	0.01556	0.02667	-3.08444	3.05222	-2.06000	-2.04889	0.09111	-3.12333	0.11333	-0.01555	-0.07555	00090.0	-0.04000	0.02556	5.00333

667 0.00778 0.02556 0.02778 -0.02000 -0.05111 -0.05555 -0.03839 -0.05778 0.000899 -0.00667 0.03222 0.032778 0.025111 0.00252 0.002020 0.03211 0.00222 0.00222 0.00222 0.002111 0.00222 0.00222 0.00222 0.002111 0.002111 0.00222 0.00267 0.00252 0.00667 0.00233 0.00222 0.002111 0.002111 0.00745 0.00252 0.00667 0.00252 0.00067 0.00222 0.00267 0.002111 0.002459 0.00667 0.00111 0.00349 0.00267 0.00111 0.00349 0.00267 0.00111 0.00349 0.00222 0.00267 0.00111 0.00349 0.00222 0.00267 0.00111 0.00349 0.00222 0.00267 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.00349 0.00222 0.00111 0.001555 0.00249 0.00333 0.00244 0.00267 0.00249 0.00333 0.00244 0.00267 0.00249 0.00333 0.00244 0.00267 0.00249 0.00333 0.00244 0.00252 0.00249 0.00333 0.00244 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00111 0.00349 0.00252 0.00244 0.00274 0.00252 0.00252 0.00244 0.00252 0.00252 0.00252 0.00244 0.00252 0.00252 0.00252 0.00244 0.00252						
667 0.00378 0.02556 0.02778 -0.02000 -0.05111 -0.05555 -0.03889 -0.03577 0.00378 0.00889 -0.00667 -0.01111 0.00667 0.03222 0.02778 0.05111 0.02520 0.00389 -0.02551 0.00889 -0.00889 -0.00811 -0.02022 -0.02650 0.03164 -0.02020 0.02223 0.03899 0.00889 0.008989 0.008989 0.008989 0.008989 0.008989 0.008989 0.00899 0.00898	o .	0700 0222 0044 0044	00000000000000000000000000000000000000	0344 0533 0322 0555 0233 0366	0.0144 0.0355 0.0755 0.0244 0.0033	00.00000000000000000000000000000000000
667 0.00778 0.02556 0.02778 -0.02000 -0.01111 -0.05555 -0.03889 0.00089 -0.00667 0.001111 0.000222 0.02222 0.03111 0.000222 0.00333 0.00689 0.00645 0.00657 0.03222 0.03111 0.02222 0.03111 0.00222 0.00555 0.03111 0.02222 0.03111 0.02222 0.03111 0.03455 0.03555 0.03557 0.0323 0.00689 0.00667 0.03111 0.03222 0.03111 0.03452 0.03111 0.03222 0.03111 0.03452 0.03111 0.03222 0.03111 0.03452 0.03111 0.03222 0.03111 0.03452 0.03111 0.03222 0.03111 0.03145 0.03452 0.03111 0.03222 0.03111 0.03452 0.03111 0.03222 0.03111 0.03452 0.03111 0.03145 0.03144 0.03145 0.03144 0.03145 0.03144 0.03144 0.03145 0.03111 0.03145 0.03111 0.03145 0.03144 0.0	80	.0577	0.0455 0.0344 0.0088 0.0533 0.0900		0.0189 0.0211 0.033 0.0433 0.022	0.06550 0.06550 0.01680 0.01887 0.01880 0.00811 0.015660 0.016600
647 0.00778 0.02556 0.02778 -0.02000 -0.05111 -0.0555 647 0.00389 -0.00667 -0.01111 0.00667 0.03252 0.03274 1010 0.0333 0.00889 0.00445 -0.00555 0.03657 0.03657 1010 0.0333 0.00889 0.00445 0.03411 0.01525 0.03657 102667 0.00567 0.01111 0.03467 0.03413 0.00555 0.03657 10333 -0.03778 -0.05000 -0.08111 -0.07000 -0.05555 -0.00744 1045 0.14556 0.07445 0.02889 0.03413 0.03445 0.01748 10566 -0.0522 -0.08889 0.00788 0.01778 0.03444 0.03444 10566 -0.05555 -0.00889 0.0383 0.04644 0.03444 10566 -0.05555 -0.00778 0.00233 0.04647 0.03444 0.03444 105655 -0.00778 0.00778 0.00778 0.00755 0.03657 10565 -0.00778 0.00778 0.00778 0.00757 0.00757 0.0222 0.03657 10566 -0.00555 0.00778 0.00555 0.00778 0.01667 0.02557 0.03657 10566 -0.00555 0.00778 0.00778 0.00778 0.01647 0.03567 0.02567 0.02567 0.00778 0.00778 0.01644 0.03657 0.00778 0.00778 0.01778 0.03667 0.00778		.0388 .0511 .0203	.0277 .0377 .0377 .0346 .0846	0.0922 0.0922 0.09444 0.0377 0.05111 0.0144	0.0355 0.0333 0.0333 0.0377	00.02888 00.01888 00.04888 00.04888 00.03688 00.04888
667 0.00778 0.02556 0.02778 -0.02000 -0.05111 0.00687 0.0322 0.00889 -0.00667 -0.01111 0.00667 0.0322 0.00889 0.00889 0.00445 0.00522 -0.02000 1111 0.00333 0.00889 0.008811 -0.07000 -0.0555 0.0355 0.0354 0.00867 0.03445 0.0355 0.0345 0.0355 0.0355 0.0355 0.0345 0.0345 0.0355 0.0355 0.0345 0.0345 0.0355 0.0355 0.0345 0.0355 0.0355 0.0345 0.0355 0.0345 0.0355 0.0355 0.0355 0.0355 0.0355 0.0355 0.0252 0.0233 0.0555 0.0255 0.0252 0.0233 0.0555 0.0555 0.0252 0.00211 0.00555 0.0355 0.00555 0.00278 0.00233 0.0555 0.0355 0.00555 0.00578 0.00555 0.00555 0.00555 0.00578 0.00578 0.00555 0.0355 0.00578 0.00555 0.0355 0.0357 0.0355 0.035	9	.0555 .0277 .0144	.0077 .0056 .0122 .0588 .0388	. 2822 . 2822 . 2338 . 2322 . 2256 . 2200		0.0144 0.0277 0.0270 0.02111 0.0224 0.0223 0.0228 0.0218
667 0.00778 0.02556 0.02778 -0.0020 667 0.00889 -0.00667 -0.01111 0.005 669 0.00889 -0.001111 -0.005 111 0.00333 0.00889 -0.01111 -0.005 111 0.003478 -0.05000 -0.08111 -0.006 445 0.02667 0.01111 0.03445 0.014 666 0.05522 -0.08889 -0.034 0.014 667 0.00667 0.01111 0.03467 0.014 666 0.05522 -0.08889 -0.03465 0.007 556 0.02667 0.01111 0.01555 0.025 578 0.02667 0.01333 -0.026 0.025 578 0.02667 0.01333 -0.026 0.025 578 0.02667 0.02000 0.00222 -0.001 500 0.02556 0.02667 0.01346 -0.02667 0.011 500 0.02567 0.02667 0.026	ū	.0511 .0322 .0200	.0555 .0122 .0344 .0111 .0177	.0422 .0355 .0444 .0211 .0366 .0165		0344 0222 0222 0220 0322 0322 0323 0323
10.02556 0.02778 0.02556 0.02777 0.00778 0.00255 0.002111 0.00332 0.00889 0.00667 0.002111 0.00332 0.00889 0.00889 0.00667 0.002667 0.00567 0.00567 0.00567 0.00567 0.00567 0.00567 0.00567 0.00567 0.00567 0.00111 0.003444 0.00522 0.00578 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00778 0.00555 0.00555 0.00778 0.00555 0	4	.0066	.0700 .0411 .0133 .0122 .0388	0.0466 0.0250 0.0666 0.0033 0.0011		710. 000. 0000. 0000. 0000. 0000. 0000. 0000. 0000.
1	بر	0277	0.0811 0.0466 0.0344 0.0288 0.0988	0.0033 0.0033 0.00366 0.0022 0.0022	.0366 .0366 .0022 .0500 .0233	0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.0
5657 1111	7	.0255 .0066 .0311	.0500 .0577. .01111. .01111. .011111.	0555 0555 0133 0555 0200	.0055 .0366 .0355 .0355	0788 0733 0733 0733 0733 0733 0733 0733
000 m 4 0 4 0 M 0 M 1 M 1 M 1 M 1 M 1 M 1 M 1 M 1 M	т,	0000	000-00	000000000000000000000000000000000000000	000000	00000000000
0.0000000000000000000000000000000000000	Oi.	0.0166 0.0166 0.0100	0.0344 0.0344 0.0166 0.0244 0.0766	04778 04778 07111 03222 06000	0.0344 0.0277 0.0088 0.0100 0.1100	0.0322 0.1866 0.0277 0.0533 0.0533 0.0555 0.0488 0.0511 0.0511

TABLE 4A CONT.

o	0.0711	2745.0	- J. J. 25	٠,			'			'					•		10	7	0.0	C	1	0	-0-0044	0-1344	0.3277	-0-0011	56	22	0-0588	-2.2511	0-0922	7 7 2 7 7	
ω	3.39555	;	111111-0-	٠,	2.01557					•					•	•					•		1	-0	ဂ်	0-0700	-2.32330	2222	2005		3.36111	7 76111	11100-0
7	0.11111	•	0.33555	-0.02030	٠,			•							•	•		•	•		•		•	-0.07222	0.04839	0	-0-	0.021	0	215	0-10334	000	•
9		; ;	0.0	ן ו	0	C	0.04889													0.00445				-0-		-0-1	-0-	ċ	-0-	0.0	0.31567	2 4	
மு	0.11111	-0.08222	•	-0.03666	-0.04444	0.02778	0.04657	-0.13333	55550-0-	-0.02222	-0.05656	0.07667	-0.08555	0.05222	-0.01444	0.01445	81160.0-	0.01334	-0.04889	-0.03222	-0.03555	0.07111	-0.04666	-0.01555	0.01778	-0.09111	-0.03889	0.05222	-0.04689	0311	-0.07778	1000	-0.07662
4	0.06667	5	5	0.01445	-0.32111	0.32889	0.02667	-0.09778	-0.32889	ï	1	0	7	C	1	0	7	7	1	0	-0.02111	J	-0.07333	.063		•	G	0.05111	.05		-0.12444		•
ю.	0.00667	-0.08000	0.02778	0.11667	-0.00222			-0.06111	-0.01000	-0.01333	-0.03111	-0.00111	-0.12333	0.00556	-0.01889	0.09445	-0-11656	0.03889	0.00556	0.03445	0.00111	0.04111	-0-	0.08000	9	0-	-0-	0-0	9	9	-045	0.0000	
7	-0.08778	-0.04778	0.02000	0.10667	-0.00555	0.01000	0.02667	-0.02333	0.03889	55400	03000	00090	81190	98586	01333	15778	94995	05111	02556	03334	04111	01555	-0.34303	0.09778	-0.02889	0.02111	-0.02222	0.07667	0	2164	-0.33778	2210	
٦	-0.12555	-0.02222	0.07222	0.05445	-0.05111	0.03667	87760.0-	1	O	7	1	7	1	0	O	U	U	O	0	7	0	7	7	0.05333	9	9	-036	0	0.06445	.0311	00 0	-0-01889	
0	-3,16111	3.03222	0.07567	-3.01778) C :	0.09334	0.01111	0.00334	00051.0	0.01889	-0.04111	0.00334	-0.00333	0.06111	-0.01222	-3.10222	0.00556	0.08333	-0.01666	-0.03666	0.04000	-0.07444		2.00445		•	•		- 19	.048	076	2700-	•

ø	.3333	.1120.	.3322	.3555	.3165	.0511	-0366	.3>22	.338B	.0522	.3388	.0033	- 3344	· 2544	.0233	.3244	.334¢	.0811	1700-	.3322	- J365	-3155	.0388	.4544	.0400	.3388	.0710	.3365	. 3244	.0000	.0533	.3255	0.05889	11110-	-3144
∞	.0511	.3877	.0322	.3333	. 2522	-0155	.3022	. 3833	- 3452	.0533	.3577	.3455	.3344	.3955	.0433	.3388	603	.3122	. 3255	. 3733	.1155	.3055	-3455	. 3277	.3333	- 2055	.0133	-3452	.0233	.0100	.0833	. 3333	0.06223	.1233	033
1	.1233	\$41C.	-0344	.0377	.0383	.0522	.3365	.3103	.3311	.3211	.0877	.0133	. 3233	.1265	+95C-	.0565	3355	1100.	.0300	.3465	.0200	.0155	.0144	.0255	-4325	.0155	1150	.0388	.000	1140.	.1133	. 3565		244	-0.00333
9.	399C	.1222	- 3222	.0055	.1256	.0700	.3011	1110.	.3288	.0055	- 3722	.0333	.0133	.3888	. 3833	.3355	.0211	.3056	.3400	.100.	1190.	.3388	-3044	. 3177	.4300	.0500	.3122	.3122	-3344	* 3444	. 3233	. 3177	CI	550C-	.3088
- t3	.0544	.1188	9	.0544	1100	.1333	-0366	.0544	.0533	-0266	.0122	-0722	.0377	.0722	.0788	.0144	0.333	- 0244	.1333	.0377	-0422	.0465	-0422	.0255	.4288	.0155	.0366	.0377	9900.	.0144	.0511	.0211	2	.0165	-0277
4*	.0811	.3522	.0333	.0433	.3330	.1522	.3455	.3577	1101.	.3288	.0600	.1211	.0133	.3430	-0366	.0688	.3630	.3322	.1377	.3611	1140-	. 3255	.3277	.0088	.4511	1100.	.0100	.3277	.0233	-0400	.3777	.0377	0	.0055	.0388
က	.0377	-0066	.0433	.0111	550	.1700	.0088	.0333	.0477	.0077	-0477	.0455	.0166	.0422	-0344	.0600	.0122	.0322	.3833	.0755	.0077	5500	-0044	.0055	.5388	.0111	.1022	.0211	.0277	.0388	.0155	.0477	-0.05222	.1266	833
21	.3211	.3266	.0755	.0155	.3422	.1188	.0244	.3344	.1033	3044	.0466	.0400	.3055	.0144	.0222	.0833	2033	.0133	- 3844	.0711	.3366	.0022	.0066	.0033	.4266	.0155	.0811	.0255	.0422	.0055	.1055	.1222	.00	.0888	0.10667
7	.0222	.0344	.0544	.0177	.0544	.0511	.0355	.0122	1770.	.0122	-0344	.0433	.0156	.0333	-1022	.0411	.0238	.0122	.0722	.0822	44C0-	.0522	9900-	.0133	.4333	.0011	.0700	17701.	.0344	.0188	.1644	.0711	0.00556	.0211	.0311
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	01	.0088	3.0344	3.0888	0.0233	0.0356	3.0088	3.0133	9801.0	5.0633	0.0177	3.483C	0011.0	3.0722	.0277	3.0822	.0356	0040.0	.0177	3.0177	.0588	3.0765	.0277	0.1566	0.0500	.0033	7.00 C	7 000.0	0.0377	.0311	3.0255	.0455	.0333	0.0944	.0266	0.0555	.0344	0.0111	3.0266	0.0788	.0711

	တ	3.3	0000.0	9910.0	3.3368	5.3377	0.3822	0.0188	2 2164	7.46	2000	20000	1.00.0	0.031	0.0111	0.331	44C-C-	-0.032	737	4000	070.0	-3.686	-0.038	210.0	0.026	0.013	2.02	0.036	-0-051	0.057	0.033	-0.035	0.146	2.074	0.040	3.336	0.004	0.000	111		0.0122	0040-0	-0.07333	0.0388	-0.0133		
	ω	.0.0333	3.3122	0.0211	1100.0	-7.00.C-	-7.0627	7200	7.76.6		3.331	900.0	3.063	0.031	0.050	0.027	-3.324	2100	71000	110-6	-0.008	3.355	-2.075	2.017	0.043	-0.015	0.060	-0.011	-0.04333	3.581	0.063	-2.314	3.09 E	3.36	0.03	-2.07	2000	00.0		30.0	7710.0	3,3355	3 -2.07000	3,344¢	-2.0333		
	2	-0-02838	-0.33222	0.0122	-3.31555	0.05112	0 05223	0.0000	0.03112	-0.33555	3.3265	0.037	0.084	0.013	-0.337	0.075	2 4 0	70.00	+10.0-	-0.025	-0.036	0.070	770.0-	700.0	0.035	0.016	0.051	750-0-	-0.05111	0.135	0.058	0.041	-0.033	-0.035	9 + C - C	901.01	240	0.0	-0.00	0.00	.0355	0.0333	7 -0.0533	0.0411	7466		
	9	00110	0.0156	0.0144	7 7533	0000	10000	.0.1655	0.3956	-0.00.6	5.0022	-2.0388	0.132	-7.7377	7256	20.00		-0.021	-0.314	-2.027	9.028	0.382	74C.C-	2,015	0.021	0.038	-7.364	180 0-	-0.04111	132	7000	200	0.05	411	0 2 4 8		277.0-	0.040	-0.30	0.00	0.0033	-0 2166	2577	0.0311	24.70	-0.0407	
CONT.	ιĊ	6660	0.0222	2000	0.00.0	0.020	0.020	0.0300	0.1511	0.0155	-0.0100	0.040	0 021	0.023	20.00	0.000	-0.022	-0.010	0.025	-0.022	0.045	0.053	0.00	0.00	-0.035	2000	40.00	730.0-	-0.02000	0.00	001.0	-0.031	0.0+0	071.0-	-0.0.	0.10	-0.15	0.03	-0.02	0.00	11000	110000	-0.0000	-0-0-0-	0.0000	-0.0577	
TABLE 4A CONT	4,		0.0100	0.027	0.020.0	0.100	0.0111	-0.1044	0.167	0.017	-0 012	210	3000	0.000	160.0-	+0c · 0-	-0.038	0.005	0.034	900.0	7.50	7000	22000	0.00	0.00	0.00	0.00	81C C	-0.34565	0.020	0.010	-0.083	0.018	-0.06	-0.02	-0.04	-0.07	0.03	-0.03	0000	0 0133	-0.0133	0.0	-0.0422	0.0435	-0.0522	
T	ମ		0.01656	0.00222	0.06555	0.01777	0.0077	0.0722	A R C C	7410	00000	0.01009	21100.0	0.03655	0.026	900.0	0.010	00000	0.045	000	220.0	0000	0.002	0.057	0.036	-0.061	000.0	0.032	0.00334	0.032	-0-041	-0.056	0.016	-0.045	-0.01	-0.078	-0.004	0.01	0.010	00.0-		-0-01/1	0.0-	-0.0366	0.0288	-0.0244	
	¢.		1110-0	0.0344	0.0233	0.0511	0.322	7770		0.01	3.025	0.000	0.024	0.010	500.0-	820-0-	0.002	010	0.00	0000	-0.016	-0-036	-0.028	0.081	0.026	-0.038	-0.053	0.023	0.014	0.022	-0.031	-0.02	0.00	-0.04	0.00	-0.05	0.07	0.01		9 0	00.0-	-0.01	0	0.03	-0-01	-000	
	~	ı	001000	0.0333	0.0311	0.0688	0.033	77000	0.000	0.1046	0.0244	0.035	0.033	900.0	000100	0.043	2000	300	0.000	0.00	0.012	CZ0°0-	0.051	.0.065	0.048	-0.031	-0.003	0.021	0	0.005	-0.093	-0.016	-0.002	-0.02	0.04	-0.01	0.07		000	0000	0.0222	-0.0133	0.01334	0.0288	-0.0355	0.0100	
	c	,	1110	1222	1997	6000		0007	2661	8555	5550	2112	5333	1778	0889	2444	2222	6777	7771	3888	11111	68800	111150	111160	19910	11150	00555	12556	01567	1334	02000	05445	02667	11111	08667	02645	04445	02556	00000	05778	01556	01000	.00778	03886	.05886	334	

o	0.05667		0.0344	2000	9999	ဂ်ဂ်ဂ်ဂ်ဂ်		0.05556 0.01222 0.00000 -0.0333 -0.15333
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7	0.03112	0.04223	0000	0.02223 0.02556 -0.07555 0.03839	-0.18222 -0.01444 -0.32939 0.34889	0.0388	000000000000000000000000000000000000000	-0.01655 -0.02777 -0.17656 -0.01111 -0.15111 0.07778
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4	-0.04333 -0.03333 -0.02555	0.10223	0.05567 -0.02555 -0.00555 -0.02999	0.01334 -0.05444 0.03556 0.034445 -0.05777	1 1 1		0.08667 -0.01222 0.00001 -0.02444 0.01334 -0.08333 -0.03555	-0.03222 0.06556 0.03555 0.01334 0.02223 -0.28777
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o	010	-3.05000 0.01112 -3.05999	9959	0.01667 -0.01889 0.17112 -0.06111	-3.04111 -3.00222 -3.09222 -3.01222	.0133 .0600 .0544 .0266 .0211		0.01889 0.05445 0.00445 0.06556 -0.03222 -0.16111

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	6	-0.33222	0	-7-73111	27.00		133	2455	-0.05333	0333	213	9900	2210	250	2000	2000	7740	5.21657	3322	5500	0	77.00	2555	0.03334	3144	316¢	-0.05222	9950-	.3111	· DI	665C-	-0.03655		- 35	2	0
	&	0.	1		00000	3.0244	1110	9446-	-2.3622	2.0144	3.3144	-2.0144	2.222	-0.0555	2.0034	20.00	7.7166	2.6	-3.114¢	-2.2055	0.0100	3.3255	3.3033	3.3566	7.0183	0.0100	0.000	-2.3265	-7.1177	-0-0038	-3.33777	-0.00555	1100.	233	3	-3.32333
	1	0	1120-	1170-	441000	1155	9400	. 3377	-0.04444	0	21100.0	C	0	9 0	3 0	3 6		0.23334	7	0	G	G		0.08445		C	-0.33838	-0.02222	C		.0283	5	.00	171C-	1140.	0.03000
	9	0.33223	4600	2003	00000	, .			-0.30333	0	3	C	0.04555			•	• .	0-19556	C	0		G	7		, .	0.30778	-	-3344	12	.3155	. 2055	-0.00666	.3011	0.30111	.3033	-0.30999
CONT.	က	07	0.0133	7710	5610-	7710	0322	.024	0.03001	-0.03111	-0.01888	0.02778	•	•	21110-0	•	•	0-14667	-0.08111	•		•	•	0.06556	•	0.03778	•	•	0.00667	•	0.01778	666000-0-	-0.03888	-0.02666	-0.07444	-0.09110
TABLE 4A CONT	4	0.09223	1190-0	-0.02888		21140-0	0211	C	0.02556	-0.01444	0.00889	.0155	110.	-0088	24	0020-	.0033	0.05445	.1200	0.10667	.001	•	.0033	058	6600-	.012	.0055	2210	9	-0377	.0488	.3188	.0066		-0.04666	-0.04888
F	က	•	0	င် င	10.	3 6		.0665	0.02223	0.00334	0.04223	•	•	•	0.04112	•	•	-0-06666		•	•	•	•	0.05334	- 0000	.0088	-0066	0300	.0411	.0277	.0655	0.00001	0211	.0155	-0.04777	-0.02555
	7	•	-0588	9440	1150-	0.02778	0111	.0344	0.02556	-0466	.0622	. 3722	0-00334	.0055	.0011	.0211	-0044	- 00	. 3222	.0377	.0311	.0722	.0033	0.01334	4000	.0044	-0088	0155	.2233	1100-	.0788	.0044	.0311	C	.0211	C
	1	90	-0322	.0188	-1255	-0135	5	0	0	0	0	0.	0	9	0	0	0.0	-0-00222	0.0	0	0	0	0.		3	. 0	.013	.013	.062	.035	.0533	.0166	.0344	-012	.0100	.0377
	0	2.051	-0388	.0155	0.0300	.0311	0088	33	93	.07	.03	20.0	00.	00.	.02	0.c	00.	-0 04777	2.17	.06	0	.0	.01	.03	5		0	00.	.05	00.	.05	.03	0	0	.02	.03

6	0.000000000000000000000000000000000000
∞,	-0.021111 -0.02535 -0.02535 -0.025334 -0.01000 -0.02334 -0.010333 -0.02555 -0.01223
1	0.003034 0.003021 0.003021 0.003055 0.003051 0.003057 0.003030 0.0030 0.0030 0.003030 0.00300
9	0.000000000000000000000000000000000000
5	0.05778 -0.02222 -0.03339 -0.05778 -0.051112 -0.051112 -0.0555 -0.05555
4	0.0222 0.02339 0.026333 0.026333 0.026333 0.02645 0.02645 0.02666 0.02666 0.02666 0.00677 0.01222 0.01334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334 0.02334
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2	0.00555 0.00667 0.005667 0.005666
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0	0.04444 0.04223 0.024223 0.02445 0.02334 0.02334 0.01667 0.01667 0.01667 0.01667 0.01333 0.00444 0.001112 0.011112 0.02667 0.02667 0.03333 0.06455 0.06455 0.06455 0.06455 0.06657 0.06657 0.06657 0.06657 0.06657 0.06657 0.06657 0.06657

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6	.5	-0.0	0.0	-0.0	0.0	ċ	-0-	0-	ċ	-0-	-0-	-0-0	-0-	2.0	-2.0	0.0	0.0	-0-0	ċ	0.0	0-0-	-0-0	•	0	1	ċ	1	o o	-0.0	2.0	0.0	-0.0	-0-0	0.0	0.0	-0.0	0.0	0.0	ŀ	-0-
8	300	.1377	.0433	.1411	.0188	. 3211	.0339	.3277	.0311	.3555	.3077	0.00555	. 2088	. 3233	11110-	.3122	-3355	.3422	. 3555	-3288	. 3255	5	-3465	-0377	.1255	.1189	.0311	.1511	.0699	.0533	- 3089	.0122	.0011	++0C-	-3055	. 3233	.3189		.0011	-3488
2	.0503	-0.33838	-0183	-0422	.0244	.3277	.3365	.0277	.0155	.3233	.0355	0.04112	.0311	.3146	.0433	-0-02555	-0322	3544	.0333	.0411	.3165	0.01779	0322	-3422	11111	.1733	.0383	-1463	6640-	. 3255	-2155	-0365	.0133	.314¢	.0211	.0111	.0244	5	1700.	0455
9	13	3	7	3	G	7	7		G	7	3	0.32334	7	7	7	7		7	7	7		0.31112	0.00001	.0211	1176.	11977	.3088	-0344	.0011	. 3033	9900.	. 3222	.0033	.3288	.3156	5500	.3544		. 3256	489
ĸ	.0311	.0255	.0188	.0444	-0.02777	.0311	.0266	9900-	-0477	-0344	-0299	0.0000	.0288	-0188	-0322	88	.1155	-0544	.0144	-0-04444	.0166	-0433	19900.0	6600	.0211	89	.0655	0.03445	-0744	-0.03221	-0133	4440-	.0244	.0155	.0144	.0166	.0544	CO	.0533	1100.
4	.053	.023	.005	-027	140-	.001	-312	-022	+1C-	. J17	.036	-0.01332	.040	.081	.318	.020	-082	.106	0.31779	.019	.041	012	N	0.	-0244	0	.1233	-0133	-0344	•	9900	1110.	.0255	1110.	-028B	.0088	.0666	1110	.0211	6600-
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2	-0.00555	21	-0277	1110-	-0522	-0289	4400	.0033	7700	2155	.0933	03	-0122	.0299	.0255	.0311	-0233	.0233	-076	-0622	-0344	-3466	0.01445	0		0	0.		0	C	7			0	0	0	0	-0.37888	0.	0
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ਖ਼ਾ	-0.03888 0.05445 0.01223 -0.01777	000000	-0.01110 -0.03777 0.03777 -0.03333 0.06112 -0.00777	0.00890 0.06890 0.09332 0.05223 0.03779 0.00001	0.04667 0.02779 0.02779 0.01412 0.00779 0.10557 0.01557 0.0112 0.02890 0.07112
ო	-0.04777 0.07112 -0.00777 -0.00444	0.01110 0.01657 0.00445 -0.02332 -0.04333	-0.01888 -0.02999 0.02658 -0.04444 0.04112	0.02556 0.11223 -0.13666 0.04223 0.00112 0.01445	0.08001 0.01658 -0.01555 - 0.01555 -0.01255 -0.05111 - 0.06446 -0.02221 -0.02449 -
62	0.01890 -0.02110 0.00556 0.00334	-0.05665 -0.056644 -0.02334 -0.04888 -0.04333	0.01223 -0.00777 -0.01555 -0.04110 0.04555 0.02890	0.09445 0.10223 -0.08211 0.07112 0.00777 -0.05888	0.12445 -0.00999 -0.07666 0.16556 -0.02110 -0.06333 0.03334 0.03668 0.03668
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0	-0.23666 -0.07777 0.00223 -0.00999	-2.00999 -2.05444 0.00001 -2.05666 5.01667	-0.00555 -0.01221 -0.00444 -0.08110 0.09556 -0.10777	0.17334 -0.01555 0.02001 0.0334 0.02334 0.02112 0.03779	-0.14777 -0.05566 -0.10110 0.0023 -0.02666 -0.02666 -0.04690 0.04690 0.04690 0.04690 0.04690 0.04690 0.04690 0.04690 0.04690

6	2311	1100-	.332	-0222	0.003	.3266	3.3788	.0344	.0288	.3355	.0511	.1322	.3189	.0711	446C.	.0455	.083	9900-	.3388	.041	.3344	771C.	. J 46	.3177	.3111	. 3222	.0365	-0833	0-0-0	. 332	3.322	-3544	· 3249	. 331	.2133	.145	.1283	.154	.043	.1100-	.008	.0533
	600	3.3244	3.3022	. 3233	.0544	.3333	.0533	.3155	.3555	.1439	.1244	.0333	. 3255	. 3244	.3811	.0133	· 3844	. 3255	.3377	1140.	. 3255	.3133	1146.	.3155	.0211	. 3522	.3722	. 3977	. 3333	. 3855	. 3144	.0599	. 2477	. 2433	.2055	1311	.1077	.2154	.1433	.3022	. 3277	.0311
7	.057	-0122	.3255	. 3555	.0411	.1477	.3333	\$51C*	.0811	.3255	.1565	.0555	.0477	.0583	.0533	-3355	-0522	*51C*	233	. 3255	.0133	. 3377	.0483	. 3255	.3165	1160.	. 3577	.1044	.0203	. 3355	. 3255	. 3383	1101.	. 3522	.1577	.1793	.3755	.1344	-2322	. 3255	.0155	.0144
9	-0.0399	0.3089	774C.0-	0.3677	4440 -0-	-0.3544	0.0611	0.3089	0.000	-0.0511	-6.3822	0.1555	-0.1011	7760.0-	0.3844	-0.3144	-0.1133	0.3339	-0-318	0.3289	5400.0	-0.3355	-0.0522	0.0189	0.3022	-0.0838	0.0533	+590-0-	0.0411	3.3055	3.3277	-0.3255	0.1055	-3.3755	0.1333	-0.1399	0.3522	-0.0777	3.1722	-0.3644	-0.3277	-0.0288
	015	0433	.0555	.0577	.038	.033	.032	· 0344	. 6244	.074	,4CO.	.0388	.073	.084	.0911	.0111	.0566	.0455		.078	9000	.037	-032	.026	.015	.032	.021	.016	.048	.043	.0388	.009	-133	.075	1.0.	660.	.057	.016	.117	.073	.0311	.013
4	.0366	.0211	9900.	.0188	.0211	.0022	.0911	.0055	.0455	.0188	. 3777	.0388	0.0499	.0144	.0077	0.0044	0.0222	.0300	TTTTC.0-	.0700	.0255	944C-0	.0233	0.0188	.0189	.0166	.0110	.0411	.3422	.0744	.0233	-3156	.1155	.3688	.0211	.0311	.0066	.0288	550C-	.0355	.0133	.0056
m	.0177	.0165	.0222	.0244	.0000	.0655	.1344	.0055	.0155	.0044	.0555	.0577	.0244	.0566	.0533	.0055	.0033	.0177	-0.10777	.0733	.0477	.0122	.0055	.0089	.0255	.0077	-0166	.0311	-0500	.0355	.0311	.0022	.0511	.0288	.0622	.0666	.0022	.0000	1101.	.0011	.0000	.0189
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TABLE 4B

AUTOCORRELATION BATTLEFIELD DAY

k	1-	6201 ft.	10-	1551 ft.	1551	-3101 ft
0		0.0031		0.0212		0.0277
1		0.0024	-	0.0007	_	0.0018
2		0.0014	_	0.0021	_	0.0030
3		0.0005		0.0032	_	0.0032
3	1000	0.0001	_	0.0037	_	0.0049
5	-			0.0006		0.0009
6	30.	0.0005		0.0002		0.0003
7	100	0.0009		0.0001	_	0.0002
8		0.0012		0.0004	_	0.0004
9	172	0.0013		0.0002	_	
10	-	0.0013		0.0003		0.0007
	0.00	0.0011	M = = =	0.0004		0.0007
11	-	0.0009		0.0004		0.0006
12	-	0.0005	-			0.0001
13	-	0.0002	-	0.0002	-	0.0010
14		0.0002		0.0000	-	0.0007
15		0.0005		0.0003	-	0.0004
16		0.0006		0.0003	-	0.0001
17		0.0007		0.0005		0.0048
18		0.0006		0.0000	-	0.0003
19		0.0005	-	0.0000	. -	0.0005
20		0.0003	-	0.0002	-	0.0007
21		0.0001	·	0.0001	-	0.0007
22	_	0.0000		0.0001		0.0004
23	_	0.0002	-	0.0002		0.0003
24	_	0.0004	_	0.0004	-	0.0001
25	_	0.0005	· -	0.0000	-	0.0002
26	-	0.0004		0.0000	_	0.0002
27	_	0.0004		0.0002	_	0.0001
28		0.0002		0.0002	-	0.0001
29	-		_	0.0001		0.0000
30	-	0.0001	_	0.0001	_	0.0001
31	-	0.0000		0.0000	_	0.0002
32		0.0001		0.0002		0.0000
33		0.0002		0.0005		0.0002
34		0.0003		0.0005		0.0002
35		0.0004	_	0.0000		0.0001
36		0.0004	_	0.0003	_	0.0000
		0.0003	_	0.0001		0.0001
37 - 38 -		0.0001	-			
39	-	0.0000		0.0002		0.0000
40	-	0.0001		0.0003		0.0000
41	-	0.0002		0.0001	-	0.0000
42	_	0.0003	-	0.0002		0.0000
43	_	0.0003		0.0002		0.0000
44	_	0.0003		0.0002		0.0000
45	_	0.0002	_	0.0002	-	0.0000
	_	0.0001		0.0001	_	0.0000
46	_	0.0000		0.0001		0.0000
47		0.0000		0.0002		0.0000
48		0.0001		0.0001	_	0.0000
49		0.0001	_	0.0002		0.0001
50		0.0001	_	0.0005		0.0001

	<u>k</u>	310	1-4651 ft.	4651-	6201 ft.
0					
ī			0.0816		0.0049
1 2			0.0112		0.0023
3			0.0120		0.0008
4		- 1 To 1 M	0.0101	AGE TO SECURE	0.0001
5		-	0.0150		0.0007
5			0.0039		0.0005
7			0.0026		0.0007
8			0.0014		0.0010
9			0.0010	-	0.0011
10		_	0.0002		0.0011
11			0.0002	-	0.0010
12		_	0.0004	-	0.0008
13		2	0.0010	-	0.0004
14		- 3	0.0034	-	0.0001
15		<u> </u>	0.0031		0.0001
16		Ī 1	0.0025		0.0003
17			0.0018		0.0004
18			0.0178	4621	0.0003
19			0.0024	'	0.0004
20		· · · = 3,,	0.0025	. 6 2 6 1	0.0002
21					0.0003
22		- u : K			0.0002
23			0.0015		0.0003
24			0.0012		0.0003
25			0.0004		0.0001
26		-	0.0002	-	0.0000
27		-	0.0003	-	0.0001
28			0.0000	-	0.0003
29		•	0.0002	=	0.0001
30		_	0.0004	• -	0.0001
31		-	0.0003	-	0.0001
32		_	0.0007	-	0.0001
33		· -	0.0000	-	0.0002
34			0.0003	-	0.0003
35		_	0.0003	-	0.0002
36		v <u>-</u>	0.0000	-	0.0001
37			0.0001	_	0.0001
38		7. (21)	0.0001 0.0003		0.0001
39		_	0.0003		0.0002
40		_	0.0002		0.0002
41		_	0.0002		0.0002
42			0.0006		0.0001
43			0.0006		0.0000
44			0.0002	=	0.0000
45		<u>_</u>	0.0002		0.0001
46		_	0.0002		0.0002
47		_	0.0002		0.0002
48		<u>-</u>	0.0002		0.0001
49		_	0.0003		0.0001
50		•	0.0001		0.0001
		•	U• 000 t		0.0000

TABLE 4C

POWER SPECTRAL DENSITY

BATTLEFIELD DAY

λ	1 6201 64	20 222	
Λ	1-6201 ft.	10-1551 ft.	1551-3101 ft.
-0.	0.00168	0.00091	
0.01000	0.00229	0.00258	0.00298
0.02000	0.00577	0.00675	0.00301
0.03000	0.01280	0.01474	0.00683
0.04000	0.02512	0.02694	0.01380
0.05000	0.04145	0.04695	0.01762
0.06000	0.04944	0.05605	0.02527
0.07000	0.04344	0.03540	0.03493
0.08000	0.03536	0.01792	0.03616
0.09000	0.03611	0.01177	0.03427
0.10000	0.04979	0.00838	0.03671
0.11000	0.07341	0.00832	0.04440
0.12000	0.08966	0.00805	0.05582
0.13000	0.08512	0.00783	0.06461
0.14000	0.07005	0.00775	0.06755
0.15000	0.06500	0.00582	0.06719
0.16000	0.07923	0.00456	0.07004
0.17000	0.10050	0.00518	0.06452
0.18000	0.10250	0.00509	0.06391
0.19000	0.07870	0.00359	0.06829
0.20000	0.05382	0.00284	0.06284
0.21000	0.04873	0.00244	0.05513
0.22000	0.05753	0.00177	0.05126
0.23000	0.06602	0.00107	0.04638
0.24000	0.06240	0.00080	0.04360
0.25000	0.04714	0.00074	0.03823
0.26000	0.03495	0.00089	0.03280
0.27000	0.03483 .	0.00113	0.03166
0.28000	0.04390	0.00086	0.03148
0.29000	0.05434	0.00054	0.03150
0.30000	0.05738	0.00064	0.03269
0.31000	0.05063	0.00075	0.03414
0.32000	0.04386	0.00075	0.03673
0.33000	0.04926	0.00088	0.03946
0.34000	0.06720	0.00088	0.04090
0.35000	0.08321	0.00079	0.04246
0-36000	0.08240	9.00088	0.04709
			0.05018
0.37000	0.06740	0.00100	0.04949
0.38000	0.05534	0.00107	0.04902
0.39000	0.05984	0.00100	0.05012
0.40000	0.07740	0.00100	0.05197
0.41000	0.08781	0.00088	0.05212
0.42000	0.07777	0.00064	0.04995
0.43000	0.05662	0.00061	0.04706
0.44000	0.04233	0.00063	0.04430
0.45000	0.04396	0.00057	0.04153
0.46000	0.05436	0.00058	0.03921
0.47000	0.05931	0.00050	0.03717
0.48000	0.05170	0.00046	0.03446
0.49000	0.03803	0.00051	0.03204
0.50000	0.03093	0.00050	0.03053

λ	31	01-4651	ft.		4651-6201 ft.
-0.		0.00141			
0.01000		0.00174			0.00121
0.02000		0.00440			0.00152
0.03000		0.01044			0.00472
0.04000		0.02629			0.01184
0.05000		0.04910			0.029/3
0.06000		0.06513			0.04447
0.07000		0.06949			0.04186
0.08000		0.06423			0.03319
0.09000		0.07831		1	0.02500
0.10000		0.13129			0.01763
0.11000		0.21706			0.01474
0.12000		0.27288			0.01247
0.13000		0.24698			0.01862
0.14000		0.18810		1.1	0.01691
0.15000		0.16925		:	0.01425
0.16000		0.23372			0.01423
0.17000		0.32036			0.01326
0.18000		0.32592			0.01320
0.19000		0.23869			0.00965
0.20000		0.14728		1	0.00948
0.21000		0.13192	1	1	0.00387
0.22000		0.17423		1	0.00764
0.23000		0.21311			0.00652
0.24000		0.20546			0.00538
0.25000		0.15049			0.00433
0.26000		0.10292	!		0.00381
0.27000		0.10256	,		0.00370
0.28000		0.13951			0.00355
0.29000		0.18088	1		0.00334
0.30000		0.19149)		0.00341
0.31000		0.16048	i		0.00439
0.32000		0.12967			0.00507
0.33000		0.14996			0.00483
0.34000		0.22070			0.00477
0.35000		0.28059			0.00482
0.36000		0.27510			0.00382
0.37000	,	0.21533			0.00364
0.30000		0.16497			0.0055
0.39000		0.18257			0.00433
0.40000		0.25152			0.00528
0-41000		0.29454			0.00328
0.42000		0.25698			0.00428
0.43000		0.17453			0.00399
0.44000		0.12038			0.00341
0.45000		0.13044			0.00289
0.46000		0.17513			0.00256
0.47000		0.19765			0.00214
0.48000		0.16954			0.00238
0.49000		0.11634			0.00289
0.50000		0.08904			0.00309

-146-TABLE 4D

BATTLEFIELD DAY

POWER	SPECTRAL	DENSITY
LOWER	DI LC I WILL	17174417111

1011	Dit of Letter	ili biitteil i		1,	
λ	1-6201 ft	10-1551 ft	1551-3101 ft	3101-4651 ft	4651-6201 ft
. 00					
. 01	13.21146	14,88454	17.36529	10.03841	8,76918
. 02	2.14806	2.51289	2.54267	1,63803	1,75716
03	. 97882	1.12717	1.05529	.79835	. 90540
04	. 64177	, 68826	. 45016	. 67166	.75954
05	. 46532	. 52706	. 28368	.55120	. 49922
06	. 29168	. 33067	. 20607		
07	.15317	.12482	.12750	. 38424	. 28412
08	. 08224			. 24502	.11703
09	.05994	.04168	.07971	.14939	. 05815
		.01954	.06094	.12999	. 02927
10	.06301	. 01053	.05619	. 16616	.01865
11	.07494	,00849	. 05699	.22160	.02173
12	. 07758	,00697	. 05590	.23610	.01182
13	. 06522	. 00600	. 05176	.18924	. 01427
14	. 04945	.00547	. 04743	.13278	.01194
15	.04381	. 00392	. 04721	.12678	.01147
16	. 05266	00303	. 04289	. 15535	, 00933
17	. 06782	,00350	. 04313	. 21619	.00895
18	. 07204	.00358	. 04800	. 22907	. 00802
19	. 05887	. 00269	. 04701	. 17855	.00722
20	. 04359	.00230	. 04466	.11930	.00768
				11770	
21	. 04325	. 00217	. 04550	.11710	, 00787
22	. 05629	.00173	. 04538	. 17048	.00748
23	. 07114	, 00135	. 04698	. 22963	. 00703
24	. 07344	, 00094	. 04499	. 24181	.00633
25	. 05966	.00094	. 04151	.19046	. 00548
26	. 04656	.00149	, 04217	. 13710	.00508
27	. 04760	.00154	. 04802	. 14015	, 00506
28	.05990	, 00117	. 04298	.19036	. 00484
29	.07217	. 00072	. 04342	. 24024	. 00444
30	. 07262	, 00081	. 04321	. 24235	. 00432
31	. 06012	,00089	. 04362	. 19057	.00521
32	. 04844	, 00083	. 043.48	. 14321	.00560
33	. 05048	, 00090	. 04191	. 15366	. 00495
34	.06412	. 00084	.04051	. 21057	.00455
35	.07453	.00071	. 04218	.25133	.00432
36	.07012	.00075	. 04270	.23409	.00325
37	05528	,00082	.04059	. 17661	.00299
38	. 04421	,00086	. 03938	. 13252	.00365
39	.04787	,00080	.04010	.14606	00422
40	.06269	,00081	. 04210	.20373	. 00428
41	.07314	.00073	. 04341	.24534	.00355
42	.06753	.00056	, 04337	.22315	.00333
43	.05183	. 00056	. 04337	. 15976	.00365
44	, 04115	, 00061	.04308	. 11704	.00332
45	.04555	.00059	.04307	. 13515	. 00332
46	.05994	. 00064	. 04324	. 19312	.00283
47	.06914	. 00058	. 04333	. 23040	. 00249
48	.06299 .04766	. 80056	.04198	$\frac{20655}{14581}$. 00290
49		. 00064	. 04016	. 14581	. 00362
. 50	. 03915	. 00063	.03864	. 14269	.00391

NOTE:

After making the calculations, several small errors were found in the raw data of Aberdeen, Knox and Yuma. Since their effects on the results would be minor, the calculations were not done over. The raw data table in each case is corrected, but the smoothed data tables are as computed.

LAND LOCOMOTION LABORATORY

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Report No. 8391

Date: November 1963

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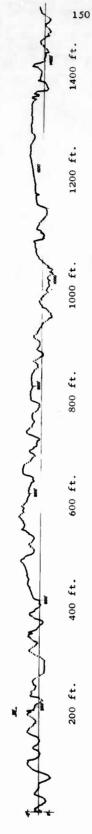


Figure 1 - PROFILE - ABERDEEN

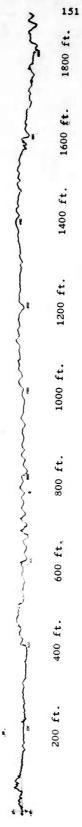


Figure 2 - PROFILE - FT. KNOX

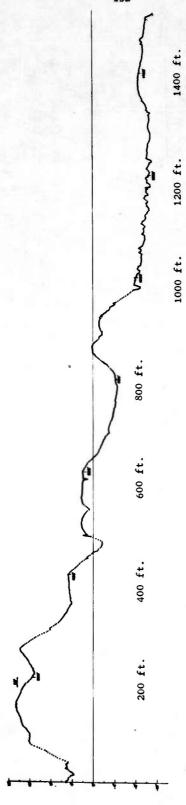


Figure 3 - PROFILE - YUMA

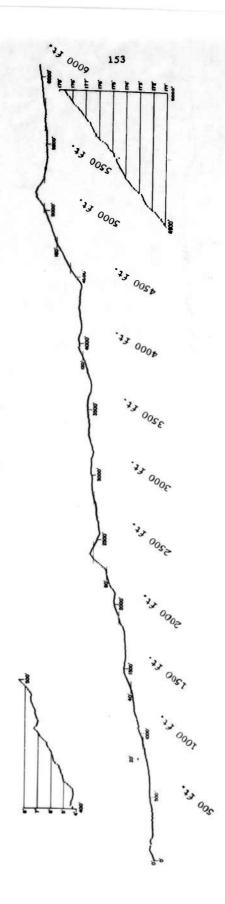
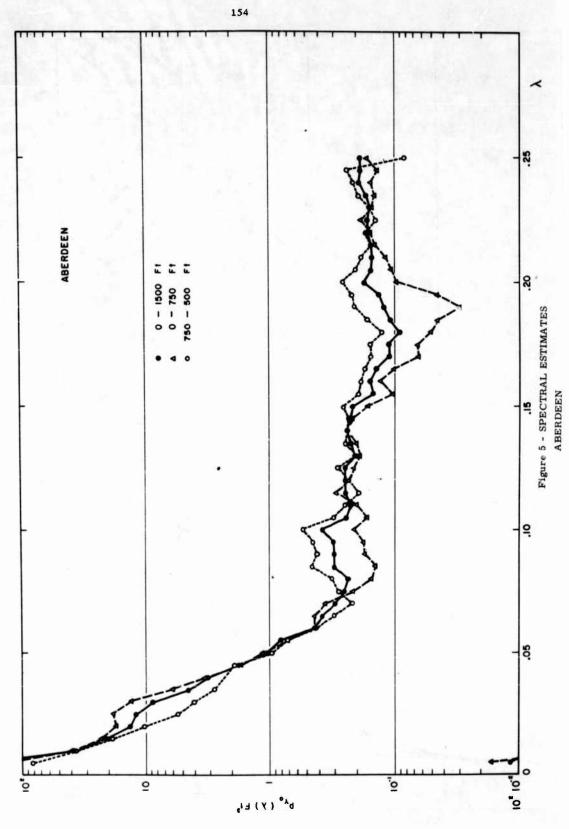
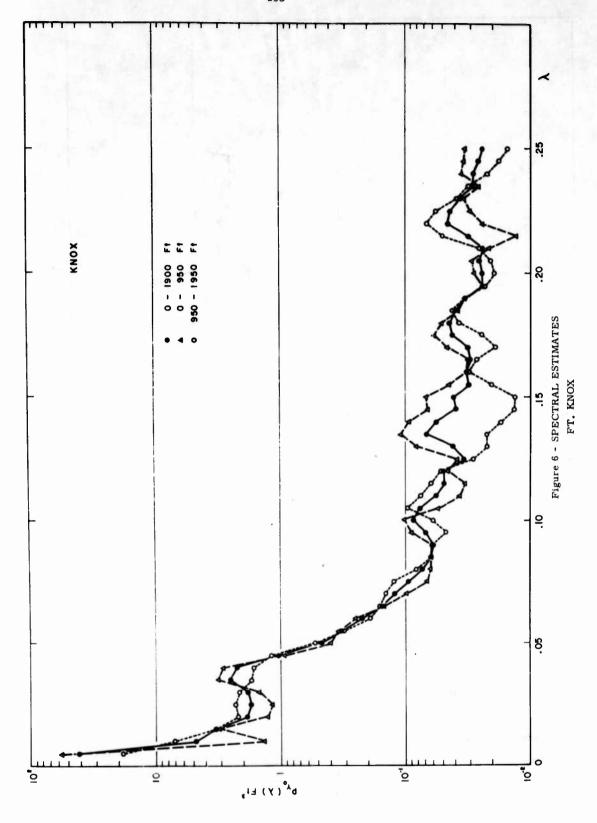
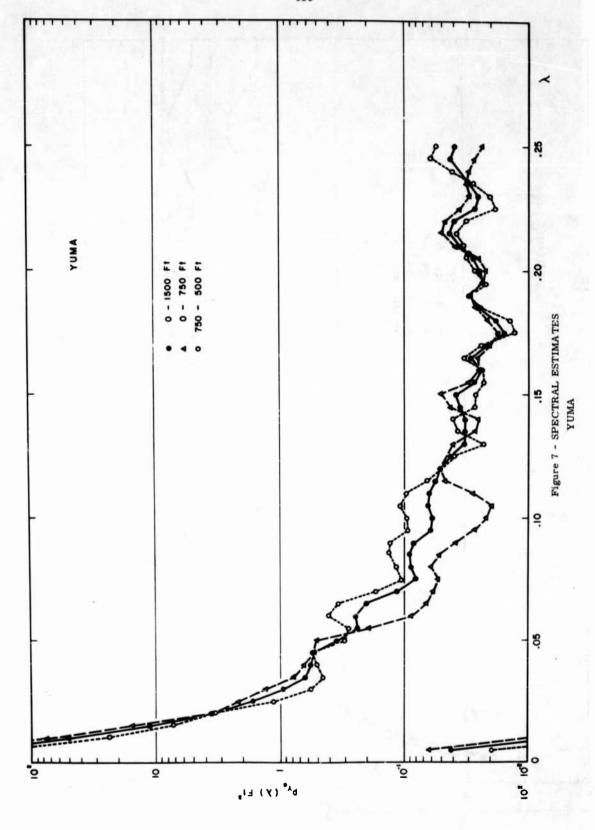
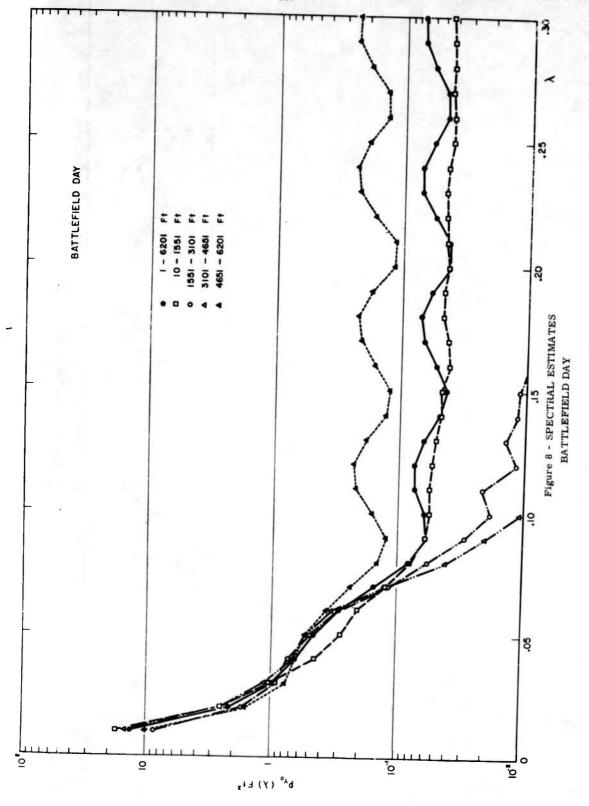


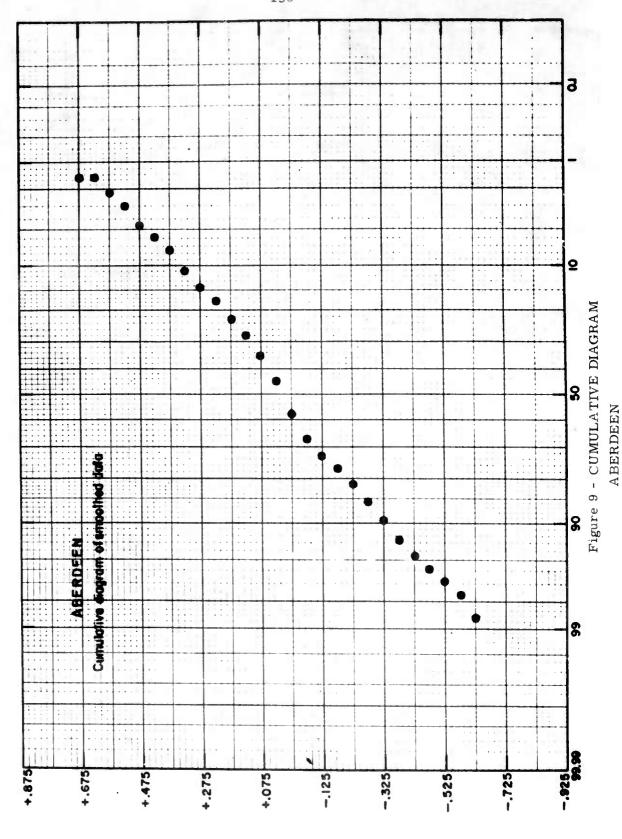
Figure 4 - PROFILE - BATTLEFIELD DAY

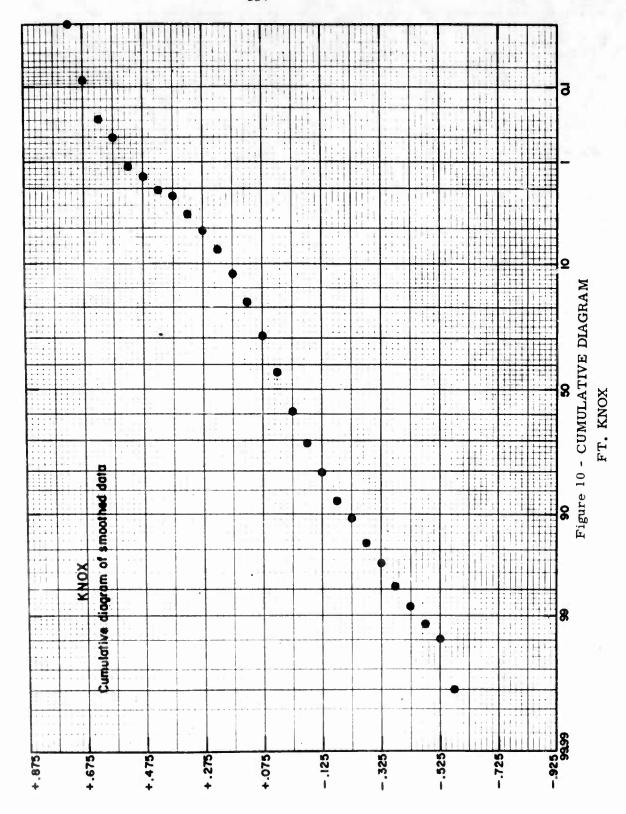


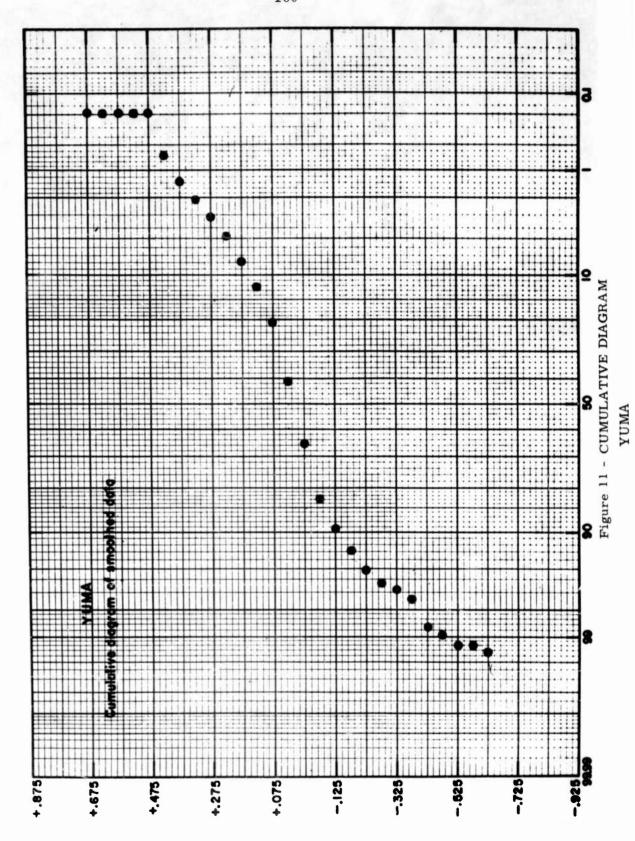












UNCLASS IFIED STATISTICAL STUDIES OF STABLE GROUND ROUGHNESS		UNCLASSIFIED STATISTICAL STUDIES OF: STABLE GROUND ROUGHNESS
AD Components Research and Development Laboratories, U. S. Army Tank-Automotive Center, Warren, Wichigan Army Tank-Automotive Center, Warren, Wichigan STATISTICAL STUDIES OF STABLE GROUND ROUGHNESS by Frank Kozin - Louis J. Cote - John L. Bogdanoff Report No. 8391 (Final), November 1963, 149 pp., Il figures Unclassified Report A method of estimating statistical characteristics of stable ground roughness for use in off-road land locomotion is developed. Computer programs for obtaining these estimates from survey data have been prepared. Several estimated power spectral densities obtained for survey data taken at military installations are	Pertinent features of these estimated power spectral dessities are interpretable in terms of visual features of plotted ground profiles. The estimated power spectral densities may be approximated by similar and simple form	Accession No. Components Research and Development Laboratories, U. S. Army Tank-Autonotive Center, Warren, Michigan STATISTICAL STUDIES OF STABLE GROUND ROUGHNESS by Frank Kozin - Louis J. Cote - John L. Bogdanoff Report No. 8391 (Final), November 1963, 149 pp., Il figures Unclassified Report A method of estimating statistical characteristics of stable ground roughness for use in off-road land locomotion is developed. Computer programs for obtaining these estimates from survey data have been prepared. Several estimated power spectral densities obtained from survey data taken at military installations are presented. Pertinent features of these estimated power spectral densities are interpretable in terms of visual features of plotted ground profiles. The estimated power spectral densities may be approximated by similar and simple form
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